

Research Paper

Cite this article: Ma L and Schluter PJ (2025). Weight misperception among Chinese children and adolescents: evidence from the repeated China Health and Nutrition Survey. *Public Health Nutrition* 28: e62, 1–8. doi: [10.1017/S1368980025000321](https://doi.org/10.1017/S1368980025000321)

Received: 19 April 2024
Revised: 28 October 2024
Accepted: 3 February 2025

Keywords:

Weight status; Weight misperception; Children and adolescents; Epidemiology; China

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Weight misperception among Chinese children and adolescents: evidence from the repeated China Health and Nutrition Survey

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Abstract

Objective: Weight misperception has been reported as a common problem in high-income countries, but there is a paucity of high-quality empirical evidence in low- and middle-income countries, especially among children and adolescents. This study estimates the prevalence of weight misperception and investigates changes over time among children and adolescents in China, as well as identifies factors that may affect this weight misperception. **Design:** The China Health and Nutrition Survey, which is a repeated, representative cross-sectional study employing multistage random cluster processes. **Setting:** A Chinese national survey across fifteen provinces and municipal cities. **Participants:** Data from children and adolescents aged 6–16 years from six consecutive waves between 2000 and 2015 were included. **Results:** The final sample totalled 7110 children and adolescents. The overall prevalence of weight misperception was largely stable between 2000 and 2015 (range: 34.1–37.3 %). Sex and age groups were associated with weight misperception, with boys and younger participants more likely to misperceive their weight status. In addition, dieting and being physically active or inactive were associated with increased rates of weight misperception. **Conclusions:** Weight misperception is common among youth in China and is unequally shared with several subpopulations at increased risk. Researchers and health promoters are called to recognise weight misperception when addressing overweight and obesity countermeasures, and more tailored public health initiatives are warranted to more effectively reach those with weight misperceptions.

Overweight and obesity in childhood and adolescence are complex issues and are known to increase the likelihood of adverse social, economic and health consequences over the life course⁽¹⁾. While these negative sequelae are well known, rates of overweight and obesity in children and adolescents continue to worsen across the globe. Between 1975 and 2016, the worldwide prevalence of obesity increased over eightfold among children and adolescents (i.e. from 0.7 to 5.6 % in girls and 0.9 to 7.8 % in boys)⁽²⁾. Although overweight and obesity rates among children and adolescents have plateaued in some high-income countries, they are now increasing in many low- and middle-income countries^(2,3). Consequently, an estimated 340 million children and adolescents were overweight or obese in 2016⁽⁴⁾. Overweight and obesity are largely preventable, but this, in part, relies on individual responsibility. Fundamental to this responsibility is having an accurate weight self-perception.

Like overweight and obesity itself, weight misperception (conceptualised as a discrepancy between measured and perceived weight status) is associated with a wide range of health problems, including eating disorders^(5–7), and psychological symptoms^(6,8). Among individuals within a normal weight range, those who perceive themselves as overweight are more likely to have unhealthy weight loss behaviours⁽⁹⁾. In contrast, overweight individuals who underestimate their weight status may have lower intention for healthy eating⁽¹⁰⁾, are less likely to control their weight⁽¹¹⁾ and are associated with increased risk of some obesity-related problems such as CVD⁽¹²⁾.

Similar to overweight and obesity, weight misperception has been initially reported as a problem in high-income countries^(9,10,13,14). More recently, there is emerging evidence that this is also an issue within low- and middle-income countries^(15–17), including China, where mean BMI and obesity rates have increased steadily since the early 1980s⁽¹⁸⁾. Although a number of large nationwide prevention programmes targeting overweight and obesity have been implemented in China over recent years, their efficacy is likely to have been hampered by this weight misperception. Overweight and obese individuals who underestimate their weight status, thereby normalise their body size, are unlikely to consider their excess weight as a health problem, resulting in low intentions for weight loss.

Concurrently, China has witnessed dramatic shifts in cultural beliefs and body image ideals following the rapid economic growth in recent decades^(19,20). These shifts have exerted

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additional influence on citizens' perception of their body image. Not surprisingly, a high proportion of Chinese children and adolescents now report body image dissatisfaction⁽²¹⁾. Further complicating matters, there is some evidence that weight misperception in China was found to be prevalent among children and adolescents^(17,22). In comparison with adults, children and adolescents are more vulnerable since they are experiencing a critical period in the life course, which encompasses elements of biological growth and major social role transitions^(23,24). As such, it is important to reliably understand the magnitude of weight misperception among youth in China and whether it is significantly worsening over time.

Extant literature focusing on weight misperception within low- and middle-income countries is rare and frequently limited to convenience sampling strategies, single cross-sectional designs and a lack of follow-up over a longer timeframe. Furthermore, understanding the factors that affect weight misperception is necessary for health professionals in developing tailored public health programmes aimed at addressing weight misperception and co-occurring outcomes. Prior findings from high-income countries and some middle-income countries indicated that weight misperception in adults is susceptible to a range of factors, which might vary by countries due to different participant characteristics^(10,25,26). Whether the influence of these factors on youth in China is open to conjecture due to its unique cultural and socio-economic characteristics, and thus, they warrant further investigation.

Using data collected from the repeatedly implemented China Health and Nutrition Survey (CHNS), this study aimed to estimate the prevalence of weight misperception among children and adolescents in China and investigate changes in this prevalence over time. The study also aimed to investigate and report on the demographic, family and sociocultural factors associated with weight misperception.

Methods

Study design

The CHNS is a repeated, nationally representative, cross-sectional study. It is an internationally collaborative project between the University of North Carolina at Chapel Hill and the Chinese Centers for Disease Control and Prevention and was designed to examine how the social and economic transformation of Chinese society affects a wide array of nutrition and health-related outcomes⁽²⁷⁾. Initiated in 1989, the CHNS has been conducted eleven times, with the latest survey undertaken in 2015. However, information on body image has been collected since 2000. Therefore, six consecutive CHNS measurement waves were employed for the present study; the first measurement wave started in 2000, when both measured and perceived weight status were elicited, through to 2015.

Participants

Detailed information on the CHNS study population, sample and quality control procedures appears in <https://www.cpc.unc.edu/projects/china/about/design/sample>. In brief, the study population is drawn from multiple districts (i.e. provinces and autonomous cities) of China, with variation in a wide range of socio-economic factors and other related health, nutritional and demographic measures. A multistage, random cluster process was used to draw the sample surveyed in each of the districts. Counties in the

districts were stratified by income (low, middle and high), and a weighted sampling scheme was used to randomly select four counties in each district. Villages and townships within the counties and urban/suburban neighbourhoods within the cities were selected randomly. For the present study, children and adolescents aged between 6 years and 16 years deriving from the CHNS measurement waves of 2000, 2004, 2006, 2009, 2011 and 2015 were eligible for our analyses.

Primary measure

Weight and height were measured by trained interviewers while the participants wore light clothing with shoes removed. Weight was measured to the nearest 0.1 kg using a calibrated scale. Height was measured using a fixed stadiometer and recorded to the nearest 0.1 cm. BMI was calculated by dividing weight in kilograms by the square of height in metres (kg/m^2). In accordance with the WHO Child Growth Standards^(28,29), BMI-for-age Z-scores were selected as an indicator of measured weight status. Under this criterion, measured weight status was collapsed into three categories using the SD scores (SDs): underweight (< -2 SDs); normal weight (-2 SDs to 1 SDs); and overweight (> 1 SD)⁽²⁹⁾. Children with BMI z-scores < -5 or > 5 were flagged as being biologically implausible and were excluded. Perceived weight status was measured by 'Do you think you are now underweight, normal or overweight?' Weight misperception was defined as discordance between the measured and perceived weight status classifications.

Demographic, family, lifestyle and media-related variables

In addition to the primary variables, information about demographic, family, lifestyle and media-related variables was collected from the CHNS questionnaire. The choice of these variables that might produce an influence on weight misperception was based on the relevant literature^(24,30). Five indicators were selected as demographic variables, including sex, age group, locality, ethnicity and region of residence. Paternal and maternal presence were assessed as family variables. We considered physical activity, dieting, fast food consumption, snacking while watching TV and watching TV while eating meals as lifestyle variables. In addition, access to internet and bedroom TV were evaluated as media-related variables. A detailed description of the names, measures, response options and codings of these variables is presented within the online Supplementary Materials Table S1.

Statistical analysis

Reporting of analyses was informed by the STROBE guidelines (www.strobe-statement.org). Frequencies, together with weighted percentages on participants' demographic characteristics, were described by measurement waves. Weight misperception for the total observations was calculated, and trends in the prevalence were described.

The pattern of weight misperception rates over time was investigated using degree-2 fractional polynomial regression models from the set of powers (-2 ; -1 ; -0.5 ; 0 ; 0.5 ; 1 ; 2 ; 3)⁽³¹⁾. The best time polynomial specification was then used in all subsequent regression analyses. In addition to the overall rates, we stratified the sample by age group and sex to explore whether weight misperception was equally distributed between these subgroups.

The potential factors of weight misperception were explored through complete case multivariable analyses. As conventionally employed logistic regression models produce biased and inflated estimates when the outcome of interest is not rare, a modified Poisson regression approach (with log-link function and robust variance estimator) was used to estimate prevalence ratios (PR) directly⁽³²⁾. A base model and adjusted models were employed. For the base model, only demographic variables (i.e. sex, age group, ethnicity, region of residence and locality) were analysed. A fully adjusted model was subsequently conducted. Rather than employing bivariable analyses to screen risk factors, all candidate variables were included in the adjusted model regardless of their statistical significance⁽³³⁾. The area under the curve of the receiver operating characteristic (AUC) was used to assess this adjusted model's predictive accuracy. In accordance with Hosmer and Lemeshow's recommendations, an AUC of 0.5 indicates no discrimination, 0.7–0.8 is regarded as acceptable, 0.8–0.9 is regarded as excellent and more than 0.9 is regarded as outstanding⁽³⁴⁾. Finally, in an effort to account for the missing data, sensitivity analyses were conducted on all regression models using chained equations multiple imputation ($M = 50$) methods⁽³⁵⁾. All demographic and study variables were included within the multiple imputations. PR and associated 95 % CI were reported, and Wald's type III χ^2 statistic was used to determine variable significance within these regression models. All the analyses were performed with Stata SE version 18.0 (StataCorp), with a two-sided $P \leq 0.05$ considered significant.

Results

Participants

The final sample was composed of 1149 (16.1 %) participants from measurement wave 1 (2000), 1355 (19.1 %) from measurement wave 2 (2004), 1072 (15.1 %) from measurement wave 3 (2006), 1016 (14.3 %) from measurement wave 4 (2009), 1420 (20.0 %) from measurement wave 5 (2011) and 1098 (15.4 %) from measurement wave 6 (2015) and when combined totalled 7110 children and adolescents. Participants' mean age was 11.0 years (range: 6–16 years), and 3393 (47.7 %) were female. Table 1 describes the participants' demographic characteristics. Overall, 55.8 % (3966) were children aged 6–11 years, 14.7 % (1042) were identified as ethnic minority, 19.5 % (1390) were from North, 17.3 % (1232) from West, 30.0 % (2132) from East and 33.1 % (2356) were from Central China, and 63.3 % (4455) were living in rural areas.

Weight misperception: distribution of measured and perceived weight status

Valid measured and self-report perception data were available from all 7110 participants. Based on anthropometric measurements, 1011 (14.2 %) children and adolescents were classified as underweight, 5054 (71.1 %) as normal and 1045 (14.7 %) as overweight. For perceived weight status, 1445 (20.3 %), 4885 (68.7 %) and 780 (11.0 %) reported being underweight, normal and overweight, respectively. Table 2 presents the matched distribution of these measured and perceived weight status classifications. Overall, 2530 (35.6 %) participants misperceived their weight status. The majority of underweight (57.2 %) and overweight (62.3 %) participants misperceived their measured weight classification. In contrast, most individuals within the normal range (74.3 %) perceived themselves accurately.

Prevalence of weight misperception

The overall prevalence of weight misperception was 35.6 %, ranging from 34.1 % in 2000 to 37.3 % in 2011. Patterns in weight misperception over measurement waves were investigated using degree-2 fractional modified Poisson regression models. There was no evidence for first- or second-order relationships, implying that the prevalence of weight misperception was largely constant over the study period. Figure 1 depicts the estimated prevalence of weight misperception and associated 95 % CI by year, together with the overall constant-only estimated mean prevalence and associated 95 % CI.

The potentially differential influence of age group and sex on weight misperception was next investigated. In modified Poisson regression analyses, a significant difference in misperception was found between sexes ($P = 0.02$) and age groups ($P < 0.001$), but no evidence for an age group \times sex interaction existed ($P = 0.52$). In these analyses, the estimated prevalence ratio (PR) for male misperception was 1.08 (95 % CI: 1.01, 1.15) compared with females and 1.24 (95 % CI: 1.16, 1.32) for children compared with adolescents. The estimated prevalence of weight misperception was 40.3 % (95 % CI: 38.4 %, 42.2 %) for male children, 37.2 % (95 % CI: 35.3 %, 39.3 %) for female children, 32.7 % (95 % CI: 30.8 %, 34.7 %) for male adolescents and 30.2 % (95 % CI: 28.4 %, 32.2 %) for female adolescents.

Potential factors affecting weight misperception

In the 2000 measurement wave, two variables of interest (access to internet and bedroom TV) were not collected, and thus, this secondary analysis was limited to the 2004–2015 measurement waves ($n = 5961$). The distribution of weight misperception for demographic and potentially confounding variables for this subsample appears in Table 3.

For the base model ($n = 5939$; 99.6 %), the effects of demographic variables (i.e. sex, age group, ethnicity, region of residence and locality) on weight misperception (Table 3) were initially investigated. In this model, age group was significant ($P < 0.001$), with younger participants more likely to misperceive their weight status than adolescents. Sex was also significant ($P = 0.02$), with boys having a higher PR of weight misperception compared with girls. In contrast, ethnicity, locality and region of residence were all non-significant.

Following the base model, a complete case multivariable model was undertaken ($n = 3292$; 55.2 %) investigating the effects of family, lifestyle and media-related variables on weight misperception; see Table 3. Among the considered potential factors affecting weight misperception, the median pairwise correlation was 0.01, with the highest between paternal and maternal presence ($r = 0.56$). In this model, age group remained significant ($P = 0.005$), as did physical activity ($P < 0.001$), but all other considered variables were non-significant. As before, weight misperception among children had a higher PR than among adolescents. Interestingly, both too little and too much self-reported physical activity were also associated with higher PR estimates of weight misperception. The estimated AUC for this model was 0.58 (95 % CI: 0.56, 0.60), which is not considered predictive acceptable, suggesting that other important unmeasured confounders exist⁽³⁴⁾. However, there was no evidence that the model assumptions were violated (deviance goodness-of-fit $P > 0.99$).

Given that only 55.2 % of the subsample was utilised in the complete case multivariable analysis, a sensitivity analysis was undertaken using multiple imputed data ($M = 50$). Table 3 and Fig. 2 present estimated PR and associated 95 % CI of the factors

Table 1 Participants' demographic characteristics by measurement wave

	2000		2004		2006		2011		2014		2015		Total	
	<i>n</i> 1149		<i>n</i> 1355		<i>n</i> 1072		<i>n</i> 1016		<i>n</i> 1420		<i>n</i> 1098		<i>n</i> 7110	
	<i>n</i>	%	<i>n</i>	%	<i>n</i>	%	<i>n</i>	%	<i>n</i>	%	<i>n</i>	%	<i>n</i>	%
Age (6–11 years)	541	47.1	624	46.1	597	55.7	560	55.1	858	60.4	786	71.6	3966	55.8
Sex (girls)	583	50.7	643	47.5	507	47.3	454	44.7	686	48.3	520	47.4	3393	47.7
Ethnicity (minority)*	153	13.4	199	14.7	181	16.9	172	17.0	180	12.7	157	14.4	1042	14.7
Region of residence														
North	241	21.0	316	23.3	227	21.2	154	15.2	267	18.8	185	16.8	1390	19.5
West	147	12.8	227	16.8	169	15.8	186	18.3	304	21.4	199	18.1	1232	17.3
East	494	43.0	44	3.2	317	29.6	304	29.9	285	20.1	288	26.2	2132	30.0
Central	267	23.2	368	27.2	359	33.5	372	36.6	564	39.7	426	38.8	2356	33.1
Locality (rural) [†]	741	68.4	882	65.4	703	65.6	676	66.5	779	54.9	674	61.4	4455	62.7

Note: *missing values for seven (2000), two (2009), four (2011) and nine (2015) participants; [†]missing data for sixty-six (2000), six (2004) and one (2015) participants.

Table 2 Participants' distribution of measured and perceived weight status (*n* 7110)

Measured weight status	Perceived weight status			Total	Overall discrepancy (misperception)
	Underweight	Normal	Overweight		
	<i>n</i> %	<i>n</i> %	<i>n</i> %	<i>n</i> %	<i>n</i> %
Underweight	433 (42.8)	546 (54.0)	323 (3.2)	1011 (14.2)	578 (57.2)
Normal	947 (18.7)	3753 (74.3)	354 (7.0)	5054 (71.1)	1301 (25.7)
Overweight	65 (6.2)	586 (6.2)	394 (37.7)	1045 (14.7)	651 (62.3)
Total	1445 (20.3)	4885 (68.7)	780 (11.0)	7110 (100)	2530 (35.6)

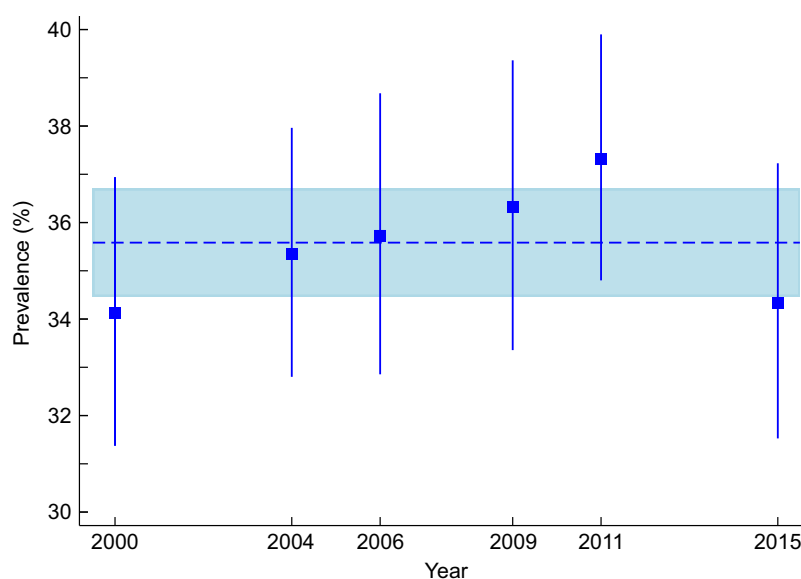


Fig. 1 Estimated prevalence of weight misperception (squares) and associated 95% CI (solid lines) by year, together with the overall estimated mean prevalence (dashed line) and associated 95% CI (shaded area).

associated with weight misperception derived from both complete case (*n* 3292) and multiple imputed (*n* 5961) analyses. In the multiple imputed multivariable modified Poisson regression analysis, age group remained significant ($P < 0.001$), as did sex ($P = 0.009$), physical activity ($P < 0.001$) and whether participants were on a diet ($P = 0.002$). The direction of the effect sizes for age group, sex and physical activity was the same as described earlier.

Being on a diet was, however, a new addition to this suite of significant variables – with those reporting dieting having an estimated adjusted PR of weight misperception 1.20 (95% CI: 1.07, 1.35) higher than their counterparts not dieting. When compared with the complete case PR estimates, the multiple imputed estimates generally exhibited modest shrinkage towards the null, with age group and dieting being notable exceptions; see Fig. 2.

Table 3 Distribution of weight misperception for demographic and potentially confounding variables, together with estimated proportional odds (PR) and associated 95 % CI for a base model including demographic variables (*n* 5939), a complete case multivariable models (*n* 3292) and multivariate multiple imputed model (*n* 5961) using modified Poisson regression analyses

	<i>n</i>	Misperception		Base model		Multivariable (complete case)		Multivariable (multiple imputed)	
		<i>n</i>	%	PR	95 % CI	PR	95 % CI	PR	95 % CI
Sex									
Male	3151	1174	37.3	1.08	1.01, 1.16	1.08	0.99, 1.18	1.10	1.02, 1.17
Female	2810	964	34.3	1	reference	1	reference	1	reference
Age group									
Child (6–11 years)	3425	1342	39.2	1.25	1.16, 1.34	1.16	1.05, 1.29	1.24	1.15, 1.34
Adolescent (12–16 years)	2536	796	31.4	1	reference	1	reference	1	reference
Ethnicity									
Han Chinese	5057	1827	36.1	1	reference	1	reference	1	reference
Minority	889	307	34.5	0.95	0.86, 1.05	0.90	0.78, 1.04	0.95	0.86, 1.06
Region of residence									
North	1149	407	35.4	1	reference	1	reference	1	reference
West	1085	381	35.1	0.97	0.87, 1.09	1.02	0.87, 1.20	1.00	0.89, 1.12
East	1638	601	36.7	1.01	0.91, 1.12	1.10	0.95, 1.27	1.03	0.93, 1.14
Central	2089	749	35.9	0.99	0.90, 1.09	1.03	0.89, 1.18	1.00	0.90, 1.11
Locality									
Urban	2240	788	35.2	1	reference	1	reference	1	reference
Rural	3714	1346	36.2	1.02	0.95, 1.09	1.00	0.90, 1.11	1.00	0.93, 1.09
Paternal presence									
Yes	4841	1701	35.1			1	reference	1	reference
No	1118	436	39.0			1.04	0.90, 1.20	1.09	0.99, 1.20
Maternal presence									
Yes	5124	1818	35.5			1	reference	1	reference
No	836	320	38.3			1.01	0.86, 1.18	1.01	0.91, 1.14
Physical activity									
Too little	1674	665	39.7			1.23	1.11, 1.35	1.21	1.12, 1.31
About right	3526	1162	33.0			1	reference	1	reference
Too much	207	104	50.2			1.56	1.30, 1.88	1.48	1.28, 1.71
Dieting									
No	5412	1908	35.3			1	reference	1	reference
Yes	452	196	43.4			1.11	0.95, 1.30	1.20	1.07, 1.35
Fast food consumption									
< 1–2 times/week	3188	1155	36.2			1	reference	1	reference
1–2 times/week	876	310	35.4			0.98	0.86, 1.11	0.98	0.88, 1.09
≥ 3 times/week	666	230	34.5			0.92	0.79, 1.06	0.96	0.84, 1.09
Snacking while watching TV									
Seldom	2883	1025	35.6			1	reference	1	reference
Sometimes	1117	407	36.4			1.00	0.89, 1.12	1.00	0.91, 1.10
Often	425	154	36.2			0.93	0.79, 1.10	1.00	0.87, 1.15
Eat meals while watching TV									
Seldom	2319	807	34.8			1	reference	1	reference

(Continued)

Table 3 (Continued)

	<i>n</i>	Misperception		Base model		Multivariable (complete case)		Multivariable (multiple imputed)	
		<i>n</i>	%	PR	95 % CI	PR	95 % CI	PR	95 % CI
Sometimes	1129	421	37.3			1.08	0.96, 1.21	1.07	0.97, 1.17
Often	982	360	36.7			1.02	0.91, 1.15	1.03	0.93, 1.14
Access to internet									
No	3966	1464	36.9			1	reference	1	reference
Yes	1881	633	33.7			0.94	0.85, 1.05	0.96	0.88, 1.04
Bedroom TV									
No	4798	1709	35.6			1	reference	1	reference
Yes	946	347	36.7			0.98	0.86, 1.10	1.01	0.92, 1.11

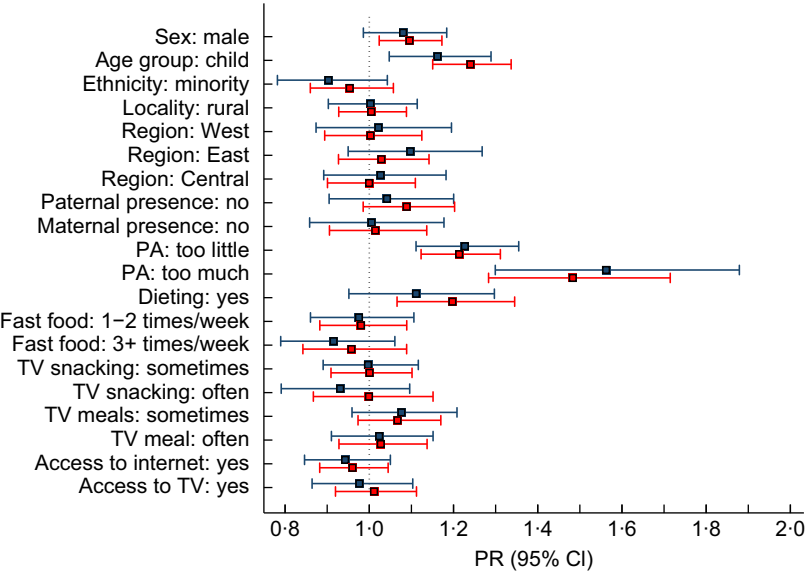


Fig. 2 Estimated prevalence ratio (PR) and associated 95 % CI of factors associated with weight misperception derived from complete case (blue; *n* 3292) and multiple imputed (red; *n* 5961) analyses from 2004 to 2015.

Discussion

Weight misperception among youth in China was observed to be common and largely constant between 2000 and 2015. Our prevalence estimates are somewhat smaller relative to those (range: 34.5–56.6 %) found in other studies targeting children and adolescents in China^(7,17,22), likely due to convenience sampling bias within these studies. When comparing between countries, our prevalence estimates are comparable to reports from some other low- and middle-income countries⁽¹⁶⁾ and high-income countries⁽³⁶⁾. Overall, these findings reinforce the view that weight misperception is a common public health challenge for youth in China. Since weight misperception can limit the effectiveness of public health initiatives aimed at reducing excess weight in this population, implementing prevention programmes recognising this discordance needs to be developed or tailored accordingly.

Both age group and sex were associated with weight misperception among children and adolescents. The prevalence decreased as age groups increased, a finding consistent with the literature⁽¹⁴⁾. One plausible interpretation for this finding is the higher cognitive ability for older than younger age groups. In comparison with girls, boys in this study seem more likely to

misperceive their weight status. It is possible that the pursuit of muscularity in boys may lead them to underestimate their weight status despite a high BMI^(37,38). It is also possible that girls have greater awareness of their weight status since they may be relatively more concerned about their body image. Owing to cultural norms and societal pressures from mass and social media, girls tend to demonstrate greater anti-fat attitudes and more positive views of thinness compared with boys⁽³⁹⁾. These sex differences are supported by two additional studies conducted in Chinese children and adolescents^(17,22). However, studies in other similarly aged populations in different countries have also observed the reverse, with more females than males misperceiving their weight^(8,40).

Another priority of this study was to explore potential factors that affect weight misperception among children and adolescents. In addition to the frequently evaluated demographic factors, a range of family, lifestyle and media-related factors were considered in our analyses, resulting in some significant findings. These findings were largely reliable even though a set of different models were employed, underlining that the factors identified in our analyses might be significant predictors of weight misperception among children and adolescents. Such information is important since it would assist health professionals in identifying subgroups

in which weight misperception is most prevalent, which could guide future programme development and implementation efforts targeting this population.

A finding of interest was that healthy eating and recommended physical activity levels reduced the risk of weight misperception among children and adolescents. Our results suggest that children and adolescents who are on a diet to lose/gain weight are more likely to misperceive their weight status. In support, several other studies have demonstrated the relationship between disordered eating behaviours and weight misperception among children and adolescents^(41–43). In addition, compared with the appropriate physical activity levels, both a higher and a lower level of physical activity were associated with a greater risk of weight misperception. It is possible that physically active individuals are more likely to underestimate their weight status and perhaps have a higher ratio of muscle to fat. In contrast, physically inactive individuals, especially those who have a lower ratio of muscle to fat, may overestimate their weight status. In line with this, several studies have confirmed that weight misperception is more prevalent among those physically active individuals^(10,44), and some other studies found weight misperception to be associated with physical inactivity^(26,41,45,46). Overall, these findings support that living a healthy nutritional and physical lifestyle may assist in reducing weight misperception among children and adolescents.

In addition to the aforementioned factors, it is noteworthy that several factors were not significantly associated with weight misperception. Although it might be interpreted that these factors do not affect weight misperception, some of these null findings warrant further investigation. For instance, since parents are critical in the socialisation of young children, their weight-related attitudes ought to play an important role in the development of weight bias in their children. This has been supported by a range of studies, including Rich and colleagues, who found that parental body dissatisfaction was associated with attributing negative traits to overweight and positive traits to thinness⁽⁴⁷⁾. Spiel and colleagues found that the father was crucial in the transmission of weight bias among children, especially boys⁽⁴⁸⁾. Other researchers showed that parents can positively or negatively influence their children's weight attitudes through the modelling of weight and dieting behaviours as well as their reinforcement through comments^(49,50). In contrast, neither paternal presence nor maternal presence was associated with children's weight misperception in our analyses; those living with a father/mother showed similar risk to others not living in this arrangement. Our results reveal that although parents exert important influence on their children, solely living with a father/mother does not significantly attenuate weight misperception. Future studies are needed to evaluate the effect of other parental factors on weight misperception.

Strengths and limitations

While this study encompasses important strengths, including the robust survey design, the large representative sample (i.e. multistage, random cluster sampling was applied, resulting in a wide spread of participants across China) and the repeated follow-up measures, which together have contributed to producing reliable and robust estimates, several limitations should be noted. Despite measured BMI z-scores, data collected for perceived weight status and other variables were primarily self-reported, which may lead to recall bias and response bias. Another limitation is that the repeated cross-sectional design of this study provided only correlational instead of causal relations. This study is also

limited to exploring the factors that may affect weight misperception, since other non-specified factors were not analysed. In addition, while the majority of geographical regions were consistent across the study period, three megacities joined this cohort since 2011, and three more provinces joined since 2015, contributing to potential sampling bias. Furthermore, information on body image was only collected from 2000, limiting our ability to identify the trend over a longer timeframe. Meanwhile, much has happened since the latest CHNS (held over 2015), including the outbreak of COVID-19, and so more recent patterns may be different.

Conclusions

Using a repeated cross-sectional design utilising data from six representative CHNS surveys, we assessed the prevalence and correlates of weight misperception among Chinese children and adolescents. The findings reveal that weight misperception is common and largely constant among youth in China. Weight misperception is also unequally distributed and more prevalent within a range of subpopulations (e.g. boys, young children, those on a diet and those being physically active or inactive). Researchers and health promoters are called to recognise weight misperception when addressing overweight and obesity countermeasures, and more tailored public health initiatives are warranted to more effectively reach those at risk.

Supplementary material. For supplementary material accompanying this paper, visit <https://doi.org/10.1017/S1368980025000321>

Acknowledgements. Not applicable.

Financial support. Not applicable.

Competing Interests. There are no conflicts of interest.

Authorship. L.M. conceived the paper, cleaned the database, performed the data analysis and drafted the initial manuscript. P.J.S. reviewed the study protocol, re-performed and finalised the data analysis and reviewed and revised the manuscript. Both authors approved the final version of the paper.

Ethics of human subject participation. This study is a secondary analysis of de-identified participants collected as part of the CHNS. These data are publicly available at www.cpc.unc.edu/projects/china/data/datasets/index.html. Identification of and dissemination to study participants is not possible or applicable given the nature of collection, public use and non-identifiable CHNS. This study was conducted according to the guidelines laid down in the Declaration of Helsinki.

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