

Article

Polygenic Risk Scores for Psychiatric Disorders Reveal Novel Clues About the Genetics of Disordered Gambling

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Abstract

Disordered gambling (DG) is a rare but serious condition that results in considerable financial and interpersonal harms. Twin studies indicate that DG is heritable but are silent with respect to specific genes or pathways involved. Existing genomewide association studies (GWAS) of DG have been substantially underpowered. Larger GWAS of other psychiatric disorders now permit calculation of polygenic risk scores (PRSs) that reflect the aggregated effects of common genetic variants contributing risk for the target condition. The current study investigated whether gambling and DG are associated with PRSs for four psychiatric conditions found to be comorbid with DG in epidemiologic surveys: major depressive disorder (MDD), attention-deficit hyperactivity disorder (ADHD), bipolar disorder (BD) and schizophrenia (SCZ). Genotype data and survey responses were analyzed from the Wave IV assessment (conducted in 2008) of the National Longitudinal Study of Adolescent to Adult Health, a representative sample of adolescents recruited in 1994–1995 and followed into adulthood. Among participants classified as having European ancestry based on genetic analysis ($N = 5215$), 78.4% reported ever having gambled, and 1.3% reported lifetime DG. Polygenic risk for BD was associated with decreased odds of lifetime gambling, $OR = 0.93$ [0.87, 0.99], $p = .045$, pseudo- $R^2(\%) = .12$. The SCZ PRS was associated with increased odds of DG, $OR = 1.54$ [1.07, 2.21], $p = .02$, pseudo- $R^2(\%) = .85$. Polygenic risk scores for MDD and ADHD were not related to either gambling outcome. Investigating features common to both SCZ and DG might generate valuable clues about the genetically influenced liabilities to DG.

Keywords: gambling; schizophrenia; polygenic risk score; comorbidity

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For many, gambling is a form of entertainment, but a subset of individuals progresses to develop a pattern of disordered gambling (DG) characterized by difficulty controlling gambling behavior, chasing losses, and serious financial and interpersonal hardships (Hodgins et al., 2011; Li et al., 2017). Surveys of nationally representative US samples estimate that approximately 2.3% of adults experience at least one lifetime symptom of gambling disorder (Kessler et al., 2008) and that the lifetime prevalence of diagnosable DG among adults is between 0.4% to 0.6% (Kessler et al., 2008; Petry et al., 2005).

Twin studies have established that DG is heritable, with additive genetic influences accounting for between 40% and 80% of the phenotypic variance in samples of adults and 5% and 20% of the variance in adolescents and young adults (Slutske, 2019). Classical twin studies provide an estimate of the aggregated effect of genes, but are silent with respect to the specific genes or biological pathways that may be involved in the disorder.

Starting from theory and empirical clues about possible neural mechanisms implicated in dysregulated gambling and related traits, researchers have tested associations between DG and a variety of functional candidate genes (Nautiyal et al., 2017; Slutske, 2019). These efforts have not yet identified strong and consistent associations

between selected variants and DG. Accumulating experience with large-scale, hypothesis-free genomewide association studies (GWAS) casts doubt on the utility of the candidate gene approach. The top GWAS hits for psychiatric traits frequently arise in unexpected genomic regions and implicate biological pathways unanticipated by prevailing etiologic theory, suggesting our intuitions about candidate genes may not be adequate for steering inquiry into the genetic determinants of disorders (Sullivan et al., 2018).

To date, there have been 2 GWAS of DG, including a total of only 2742 participants, with no genomewide significant single-nucleotide polymorphisms (SNPs) or genes detected in either study (Lang et al., 2016; Lind et al., 2012). Available evidence from GWAS involving other psychiatric conditions indicate that these disorders most likely arise from the combined influence of a large number of common variants, each of which has a small effect (Sullivan et al., 2018). Practically, this requires very large sample sizes to identify associated variants at a genomewide level of statistical significance. GWAS may yet make important contributions to elucidating the genetic underpinnings of DG, but this will require concerted efforts to amass much larger samples.

Epidemiologic studies indicate that DG is comorbid with numerous other psychiatric conditions (Edens & Rosenheck, 2012; Kessler et al., 2008; Petry et al., 2005). Investigating the extent to which these comorbidity patterns are attributable to overlapping genetic influences could provide valuable clues concerning the biological bases of DG. Multivariate twin studies have produced evidence for genetic correlations between DG and substance use disorders

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(Slutske et al., 2000, 2013; Xian et al., 2014), antisocial behaviors (Slutske et al., 2001) and obsessive-compulsive disorder features (Scherrer et al., 2015). Existing studies of DG and major depression have yielded mixed findings (Blanco et al., 2012; Potenza et al., 2005). Fitting joint twin models is tractable when studying overlap between DG and common mental disorders. However, because DG is itself fairly rare, sparse data often prevent meaningful twin analyses of genetic overlap between DG and less prevalent conditions. This may effectively limit what can be learned about genetic causes of comorbidity from biometric models using twin samples.

Although gambling research lags behind, a number of large-scale collaborative efforts have now been organized to conduct powerful GWAS investigations of several psychiatric disorders (Sullivan et al., 2018). This has generated new tools — polygenic risk scores (PRSs) — that can be leveraged to investigate the genetic bases of psychiatric disorders (Bogdan et al., 2018; Dudbridge, 2016; Maier et al., 2018). A PRS is constructed by applying regression weights derived from a target GWAS to genotype data collected from an independent sample, yielding a single summary score for each individual estimating their level of risk for the target disorder based on the aggregated effects of their common genetic variants. PRSs provide a method for examining how the phenotypic expression of an understudied disorder, such as DG, relates to common variant risk for other psychiatric disorders, including comparatively rare conditions that have been examined in large case-control GWASs.

The current study examines associations between DG and PRSs for four psychiatric conditions — attention-deficit hyperactivity disorder (ADHD), major depressive disorder (MDD), bipolar disorder (BD) and schizophrenia (SCZ) — in a nationally representative sample of young adults from the USA. Each of these disorders overlaps with DG at the phenotypic level in epidemiological samples (Clark et al., 2013; Edens & Rosenheck, 2012; Kessler et al., 2008; McIntyre et al., 2007; Park et al., 2011, 2005).

Genetic risks for MDD and ADHD are of interest because prominent theoretical accounts posit that mood disturbance and impulsivity are important risk factors (Blaszczynski & Nower, 2002; Sharpe, 2002). BD is the condition that shows the strongest association with DG in epidemiologic analyses of comorbidity (Kessler et al., 2008; Petry et al., 2005). SCZ has substantial genetic overlap with BD (Lichtenstein et al., 2009; Purcell et al., 2009). Polygenic risk scores for BD and SCZ are each related to a variety of substance use phenotypes (Carey et al., 2016; Hartz et al., 2017; Reginsson et al., 2017), suggesting they may influence multiple forms of addictive behaviors — possibly including DG.

Materials and Methods

Participants

Participants were drawn from the National Longitudinal Study of Adolescent to Adult Health (Add Health; Harris et al., 2009). The Add Health cohort is a nationally representative US sample that has been followed into young adulthood with four in-home interviews. Variables used were obtained in Wave IV ($N = 15,701$; response rate = 80%), which was conducted in 2008 when the participants were 24–34 years of age.

Procedure

Saliva samples were collected from consenting participants (96%) during the Wave IV assessment. Consent for long-term archiving was obtained from approximately 12,200 (80%) of those participants, making them eligible for genomewide genotyping. Approximately

80% of the sample genotyping was performed with the Illumina Omni1-Quad BeadChip and 20% was performed using the Illumina Omni2.5-Quad BeadChip. After quality control procedures, genotyped data were available for 9974 individuals (7917 from the Omni1 chip and 2057 from the Omni2 chip) on 609,130 SNPs common across the 2 platforms (Braudt & Harris, 2018; Highland et al., 2018).

To account for population stratification, the genotyped sample was limited to the 9129 individuals who could be assigned to one of four genetic ancestry groups based upon principal components analysis: European, African, Hispanic and East Asian. Genetic ancestry was strongly correlated ($r = .89$) with self-identified race/ethnicity. The self-identified race/ethnicity of the 9129 individuals was 5754 (63%) non-Hispanic White, 1940 (21%) non-Hispanic Black, 961 (11%) Hispanic, 449 (5%) Asian and 23 (<1%) Native American (Braudt & Harris, 2018).

Measures

Gambling and disordered gambling. The gambling assessment from the Add Health Wave IV interview included two questions: ‘Have you ever bought lottery tickets, played video games or slot machines for money, bet on horses or sporting events, or taken part in any other kinds of gambling for money?’ and (if yes to the previous question) ‘Has your gambling ever caused serious financial problems or problems in your relationships with any of your family members or friends?’ A dichotomous *gambling* phenotype was based on endorsing the first question, and a dichotomous *disordered gambling* phenotype was based on endorsing the second.

Polygenic risk scores. Details regarding the construction of PRSs can be found in an Add Health technical report (Braudt & Harris, 2018). Briefly, for a given PRS, each SNP was weighted by the regression coefficient for the corresponding SNP estimated in the discovery GWAS. These weighted effects were summed across all available variants (i.e., p -value threshold = 1.0) to yield a single quantitative estimate of risk for the target condition attributable to common variants for each participant in the independent target sample. These calculations were performed using the PRSice wrapper for R within the PLINK package (Chang et al., 2015). The PRSs were then standardized within each ancestry group, yielding within-group PRSs with a mean of 0 and a standard deviation of 1.

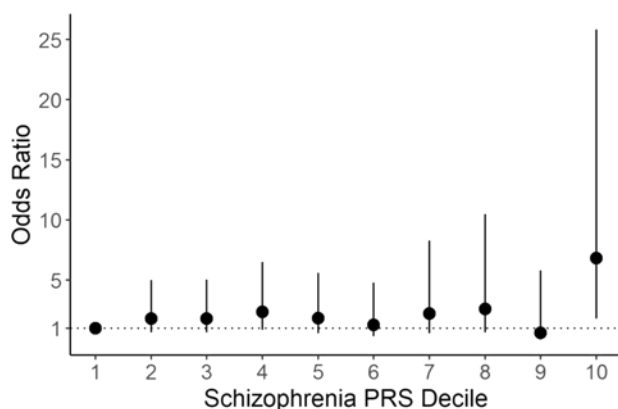
The MDD PRS was based upon summary statistics from a genomewide association meta-analysis involving 135,458 cases and 344,901 controls (Wray et al., 2018). Owing to data-sharing restrictions, the PRSs for MDD were based on summary statistics that excluded 75,607 cases and 231,747 controls from the 23andMe cohort. The ADHD score was calculated using statistics from a study of 20,183 cases and 35,191 controls (Demontis et al., 2019). The BD PRS was based upon results from a study of 7481 cases and 9250 controls (Sklar et al., 2011). The SCZ score used summary statistics from a GWAS investigation of 36,989 cases and 113,075 controls (Ripke et al., 2014). No Add Health data were included in any of these discovery GWASs.

Data Analysis

Data from close relatives were eliminated by retaining data from only a single member of each family. Because the discovery samples used to derive the PRSs were predominantly of European ancestry, analyses were restricted to individuals classified as having European

Table 1. Results from logistic regression models predicting gambling and disordered gambling

Polygenic risk score	Any gambling			Disordered gambling		
	OR [95% CI]	<i>p</i>	Pseudo- <i>R</i> ² (%)	OR [95% CI]	<i>p</i>	Pseudo- <i>R</i> ² (%)
Major depression	0.963 [0.882, 1.051]	.400	.0009	1.078 [0.786, 1.476]	.642	.0331
ADHD	0.994 [0.930, 1.063]	.868	.0009	1.007 [0.842, 1.361]	.578	.0160
Bipolar disorder	0.934 [0.874, 0.999]	.045	.1180	1.241 [0.973, 1.584]	.082	.4671
Schizophrenia	0.942 [0.853, 1.040]	.233	.0003	1.538 [1.072, 2.207]	.020	.8481

**Fig. 1.** Odds ratios and 95% confidence intervals for disordered gambling by deciles of schizophrenia polygenic risk score, with the lowest decile serving as the reference category.

ancestry according to the principal components analyses of genotyped SNPs. Together, these constraints yielded an analytic sample of 5215 individuals.

Logistic regressions were conducted predicting gambling and DG. Analyses of DG were limited to individuals who endorsed lifetime gambling. Covariates were age, sex and the first 10 ancestry principal components. For each model, the amount of variance in the phenotype accounted for by each PRS was estimated by computing the difference in Nagelkerke's pseudo-*R*² from a model that included both the covariate set and a particular PRS relative to a baseline model that included the covariates alone. This difference was multiplied by 100 to express it as a percentage of variance.

Results

Lifetime experience with gambling was reported by more than three-quarters of the analytic sample (78.4%, *n* = 4091). Table 1 summarizes results from four logistic regression analyses using individual PRSs to predict gambling endorsement. Surprisingly, higher polygenic risk for BD was associated with a significant reduction in the odds of gambling, *OR* = 0.93, 95% CI [0.87, 0.99], *p* = .045, accounting for 0.12% of the variance. No other PRS was associated with gambling behavior.

DG was comparatively rare (1.3%, *n* = 70). Table 1 also summarizes results from models predicting endorsement of DG. Polygenic risk for SCZ was associated with significantly increased risk of DG, *OR* = 1.54, 95% CI [1.07, 2.21], *p* = .020, accounting for 0.85% of variance. Figure 1 illustrates this effect by plotting the *OR* for DG as a function of SCZ PRS deciles, with the lowest decile as the reference category. This plot reveals that the effect was driven primarily by an increased risk of DG among those with the highest level of genetic risk for SCZ. Although not statistically significant, there

was suggestive evidence of a possible association between BD risk and DG, *OR* = 1.24, 95% CI [0.97, 1.58], *p* = .082, accounting for 0.47% of the variance. MDD and ADHD were not associated with DG.

In light of the similar effects for SCZ and BPD and the known genetic overlap between these conditions, we conducted an additional analysis predicting DG including both polygenic scores as predictors. The PRSs for SCZ and BPD were significantly correlated in the current sample, *r* = .28, *p* < .001. In the simultaneous logistic regression, the pair of PRSs accounted for 0.98% of the variance. Neither PRS was nominally significant, but SCZ was associated with a comparatively stronger effect. The SCZ PRS (*OR* = 1.44, 95% CI [0.97, 2.12], *p* = .07) accounted for an additional .51% of variance relative to the model including BD alone. The BD score (*OR* = 1.13, 95% CI [0.87, 1.48], *p* = .37) accounted for an additional 0.13% of variance in DG relative to a model including SCZ alone.

Discussion

The primary finding was that polygenic risk for SCZ was associated with DG in a nationally representative sample of young adults. This effect was small, accounting for <1% of the variance in DG. However, this is comparable to the average effect size for the SCZ PRS in predicting other psychiatric disorders and traits and is actually stronger than its prediction of some SCZ-related features and outcomes (Bogdan et al., 2018). This cross-disorder association suggests that common genetic factors have pleiotropic effects on DG and SCZ. Theoretically, this might indicate that one disorder is an intermediate phenotype that provides a crucial link in the causal chain, setting conditions that may facilitate the acquisition of the second disorder. Under this scenario, SCZ could facilitate dysregulated gambling, or DG might increase risk for onset of SCZ. Although one could generate hypotheses about how such effects could occur, it seems unlikely that casual transactions involving the fully diagnosable clinical conditions explain the overlap. Instead, it seems more plausible that subtle, subclinical manifestations of genetic risk for SCZ increase risk for DG.

The current findings suggest that investigating features common to both SCZ and DG might generate valuable clues about the genetically influenced liabilities to DG. Below, we speculate about some research domains that may merit closer scrutiny based on available evidence. Overlapping features that have been empirically associated with the SCZ PRS may represent especially interesting targets in light of the present findings.

Reduced cortical thickness has been observed in individuals with DG relative to healthy controls, a finding thought to be consistent with a diminished top-down control of impulsive behavior (Grant et al., 2015). Reduced cortical thickness is also characteristic of individuals with BD and SCZ (Knöchel et al., 2016; Rimol et al., 2010), and the extent of this cortical thinning is correlated with the SCZ PRS (Neilson et al., 2017). Interestingly, investigation of the genomic

region most strongly associated with SCZ in the discovery GWAS (hence, producing the most heavily weighted effects in the SCZ PRS) implicates variations in the immune-related complement system that might result in excessive synaptic pruning and cortical thinning during adolescence and early adulthood (Sekar et al., 2016). This pattern of evidence suggests that additional work examining a potential role of cortical thinning in DG may be warranted.

Aberrant dopamine functioning has long been hypothesized to play a major role in the pathophysiology of SCZ (Howes et al., 2016). Several lines of evidence also point to the involvement of dopaminergic mechanisms in DG (Nautiyal et al., 2017; Zack & Poulos, 2009). Compared to healthy controls, persons with DG show enhanced striatal dopamine release in response to an amphetamine challenge, and the magnitude of this effect is correlated with the severity of gambling problems (Boileau et al., 2014). Pharmacologic manipulations of dopamine modulate motivation to gamble and the rewarding effects of gambling in problem gamblers (Zack & Poulos, 2004, 2007), and treatment with dopamine agonists is associated with an increased risk of DG onset in patients with Parkinson's disease (Dagher & Robbins, 2009; Dodd et al., 2005). Some studies have found associations between dopamine-related candidate genes and gambling behaviors (e.g., Comings et al., 1996; Gray & MacKillop, 2014; Lobo et al., 2015). A SNP associated with the *DRD2* gene was among the 108 genomewide significant loci in the SCZ GWAS (Ripke et al., 2014). However, there is limited evidence that the SCZ PRS is related to SCZ features thought to be dopamine-related, such as positive symptoms or antipsychotic dosage (Hettige et al., 2016; Jones et al., 2016; Sengupta et al., 2017; Stepniak et al., 2014). Furthermore, common genetic variants implicated in dopamine functioning do not collectively discriminate SCZ cases from controls (Edwards et al., 2016). Thus, the polygenic liability shared by DG and SCZ might not prove to be strongly related to dopaminergic mechanisms.

Relative to healthy controls, individuals with DG (Kovács et al., 2017) and SCZ (Betz et al., 2019; Woodrow et al., 2018) exhibit impaired decision-making abilities on the Iowa Gambling Task (IGT; Bechara et al., 1994). The IGT is a complex task that taps a multifaceted construct and can reflect effects originating in a number of neural systems related to bottom-up impulsive processes and top-down reflective impulse control (Bechara, 2005; Buelow & Suhr, 2009). Moreover, impaired performance on this task is seen in a wide variety of psychiatric and neurological patient populations (Mukherjee & Kable, 2014). Thus, it is unclear whether the deficits observed in persons with DG and SCZ arise from a shared mechanism. There is some evidence that the common variant risk for SCZ is associated with risk-taking (Linnér et al., 2019; Maxwell et al., 2019), and that much of the association between the SCZ PRS and SCZ diagnosis is mediated by cognitive deficits (Toulopoulou et al., 2019). However, no studies have specifically examined whether SCZ polygenic risk is related to performance on neuropsychological gambling tasks. Future research examining how the SCZ PRS is related to parameters of IGT behavioral responding and measures of the neural underpinnings of task performance could help to determine whether and how common genetic risk for SCZ might contribute to the impaired decision-making, thereby potentially increasing risk for gambling problems.

Schizotypy is a multidimensional personality construct that is hypothesized to reflect subtle effects of genetic risk for SCZ (Lenzenweger, 2006; Meehl, 1990; Raine, 2006). Magical ideation, defined as belief in forms of causation that are invalid by

conventional standards, is a trait associated with schizotypy (Eckblad & Chapman, 1983). Magical ideation is similar to many of the superstitious beliefs and cognitive distortions characteristic of DG (Goodie & Fortune, 2013; Leonard & Williams, 2018; Toneatto, 1999; Zack & Poulos, 2009). Magical ideation is elevated among gamblers, particularly those who prefer games of chance, such as lotteries and electronic gaming machines (Savage et al., 2014). There is currently mixed evidence concerning an association between the SCZ PRS and measures of positive schizotypy or magical ideation (Hatzimanolis et al., 2017; Isvoranu et al., 2019; van Os et al., 2017).

The SCZ PRS has been shown to be associated with various forms of substance use and dependence (Carey et al., 2016; Hartz et al., 2017; Reginsson et al., 2017). The current findings extend this work, demonstrating that polygenic risk for SCZ associates with a behavioral addiction, DG, that is highly comorbid with substance use disorders in the general population (Kessler et al., 2008; Petry et al., 2005) and among individuals with psychosis (Fortgang et al., 2018). Together, findings to date might indicate that common genetic risk for SCZ confers a general liability to impaired control over rewarding behaviors rather than modulating response to a particular drug or drug class. Alternatively, the findings are consistent with a process in which polygenic SCZ risk increases substance use, which may then serve as a mediator that causally influences downstream acquisition of DG.

Prominent theoretical accounts of DG etiology point to impulsivity and dysphoric mood as important liabilities (Blaszczynski & Nower, 2002; Sharpe, 2002). The SCZ PRS has been shown to be associated with negative symptoms, anxiety, depression, and externalizing traits and symptoms (Jansen et al., 2017; Jones et al., 2016; Nivard et al., 2017; Riglin et al., 2017; Sengupta et al., 2017; van Os et al., 2017). Thus, some of the theorized DG risk factors could conceivably arise from pleiotropic effects of genetic liability to SCZ.

Polygenic risk for BD was associated with decreased odds of ever having gambled. This finding was surprising because (a) diagnostic criteria for mania and hypomania include excessive involvement in pleasurable activities with a high potential for painful consequences and (b) higher scores on the BD PRS were (nonsignificantly) associated with increased odds of DG. If this apparent protective effect of BD genetic risk on lifetime gambling proves replicable, more research will be needed to account for this puzzling finding.¹

Several limitations should be considered. A simple two-item assessment was used to define lifetime gambling and DG status. Using formal diagnostic criteria or more nuanced multiitem screening instruments would have allowed us to probe whether particular clinical features of DG were associated with polygenic risk for various psychiatric disorders. Corrections for multiple testing were not used and the nominally significant effects in Table 1 would not remain if adjusted *p*-value thresholds were applied. This approach was adopted because (a) the current study was exploratory given the lack of available information concerning the molecular genetics of DG and (b) the analyses also generated effect size estimates that can be directly compared to other studies. Nonetheless, there is clearly a need for replication of the current findings in independent samples. As expected in a general population survey, the prevalence of DG was low, limiting statistical power. Further investigation using clinical or case-control samples featuring more affected individuals would be informative. The PRSs were computed based on the ~600,000 markers shared across 2 genotyping arrays. Imputation to a reference panel prior to PRS calculation might have improved the precision of the computed scores. Add Health did not include assessments of psychosis, so

we were not able to determine the prevalence of SCZ and related conditions. Thus, although we speculate that the findings are not attributable to DG being a direct cause or consequence of diagnosable SCZ spectrum disorders, these possibilities could not be ruled out empirically. The sample size of the discovery GWAS data set has implications for the predictive power of the PRS, with larger samples permitting more precise estimates of allelic effect sizes (Chatterjee et al., 2013). The SCZ PRS has the highest estimated predictive power relative to polygenic scores for other psychiatric traits, in part owing to the massive sample size of the SCZ GWAS (So & Sham, 2016). This may partly explain why a nominally significant association was only observed between DG and SCZ.

In conclusion, our study indicates that genetic predisposition to SCZ is associated with increased risk of problematic gambling in young adulthood. Although replication studies are needed, the current finding provides new clues about the biological bases of DG. The SCZ PRS is being widely used in studies of various neuro-behavioral domains. As evidence accumulates concerning the most robust correlates of polygenic risk for SCZ, the nature of the specific mechanism(s) conferring risk for DG may become more evident. Multitrait GWAS approaches that allow joint modeling of DG and SCZ might improve detection of DG-related loci and generate more informative PRSs for use in future research (Andreassen et al., 2013; Turley et al., 2018).

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Notes

1 We considered the possibility that failure to control for educational attainment in the analysis could produce this effect. Lower educational attainment has been associated with increased odds of ever gambling (Kessler et al., 2008), and higher scores on the BD PRS have been associated with completing more years of schooling and attainment of a university degree (Power et al., 2015). However, the association between BD PRS and lifetime gambling was essentially unchanged when controlling for either measured educational attainment or an educational attainment polygenic score (computed based upon Lee et al., 2018).

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