

A DUAL RISK MODEL WITH ADDITIVE AND PROPORTIONAL GAINS: RUIN PROBABILITY AND DIVIDENDS

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Abstract

We consider a dual risk model with constant expense rate and i.i.d. exponentially distributed gains C_i (i = 1, 2, ...) that arrive according to a renewal process with general interarrival times. We add to this classical dual risk model the proportional gain feature; that is, if the surplus process just before the *i*th arrival is at level *u*, then for a > 0 the capital jumps up to the level $(1 + a)u + C_i$. The ruin probability and the distribution of the time to ruin are determined. We furthermore identify the value of discounted cumulative dividend payments, for the case of a Poisson arrival process of proportional gains. In the dividend calculations, we also consider a random perturbation of our basic risk process modeled by an independent Brownian motion with drift.

Keywords: Dual risk model; ruin probability; time to ruin; dividend

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1. Introduction

We consider a dual risk model with constant expense rate normalized to 1. Gains arrive according to a renewal process { $N(t), t \ge 0$ } with independent and identically distributed (i.i.d.) interarrival times $T_{i+1} - T_i$ having distribution $F(\cdot)$, density $f(\cdot)$, and Laplace–Stieltjes transform (LST) $\phi(\cdot)$. If the surplus process just before the *i*th arrival is at level *u*, then the capital jumps up to the level $(1 + a)u + C_i, i = 1, 2, ...$, where a > 0 and $C_1, C_2, ...$ are i.i.d. exponentially distributed random variables with mean $1/\mu$. Let U(t) be the surplus process, with U(0) = x > 0; then we can write

$$U(t) = x - t + \sum_{i=1}^{N(t)} (C_i + aU(T_i -)), \ t \ge 0.$$
(1.1)

Taking a = 0 yields a classical dual risk model, while $C_i \equiv 0$ yields a dual risk model with proportional gains. U(t) can also represent the workload in an M/G/1 queue or the inventory

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level in a storage model or dam model with a constant demand rate and occasional inflow that depends proportionally (apart from independent upward jumps) on the current amount of work in the system. We also give some results on a generalization of the model of (1.1) where at the *i*th jump epoch the jump has size $au + C_i$ with probability p, and has size D_i with probability 1 - p, where D_1, D_2, \ldots are independent, $\exp(\delta)$ -distributed random variables, independent of C_1, C_2, \ldots .

In this paper we are interested (i) in exactly identifying the Laplace transforms of the ruin probability and the ruin time, and (ii) in approximating the value function, which is the cumulative discounted amount of dividends paid up to the ruin time under a fixed barrier strategy. To find this value function we solve a two-sided exit problem for the risk process (1.1), which seems to be interesting in itself. In the discounted dividend case we also add to the risk process (1.1) a perturbation modeled by a Brownian motion X(t) with drift; that is, we replace the negative drift -t by the process X(t).

More formally, we start from the analysis of the ruin probability

$$R(x) := \mathbb{P}_x(\tau_x < \infty), \tag{1.2}$$

where $\mathbb{P}_{x}(\cdot) := \mathbb{P}(\cdot | U(0) = x)$ and the ruin time is defined as the first time the surplus process equals zero:

$$\tau_x = \inf\{t \ge 0 : U(t) = 0\}.$$
(1.3)

Our method of analyzing R(x) is based on a one-step analysis where the process under consideration is viewed at successive claim times. We obtain the Laplace transform (with respect to initial capital) of the ruin probability for the risk process (1.1). We also analyze the double Laplace transform of the ruin time (with respect to initial capital and time).

Another quantity of interest for companies is the expected cumulative and discounted amount of dividend payments calculated under a barrier strategy. To approach the dividend problem for the barrier strategy with barrier b, we consider the controlled surplus process U^b satisfying

$$U^{b}(t) = x - t + \sum_{i=1}^{N(t)} \left(C_{i} + aU^{b}(T_{i}) \right) - L^{b}(t), \qquad (1.4)$$

where the cumulative amount of dividends $L^{b}(t)$ paid up to time t comes from paying all the overflow above a fixed level b as dividends to shareholders. The object of interest is the average value of the cumulative discounted dividends paid up to the ruin time:

$$v(x) := \mathbb{E}_{x}\left[\int_{0}^{\tau_{x}^{b}} e^{-qt} dL^{b}(t)\right], \qquad (1.5)$$

where $\tau_x^b := \inf\{t \ge 0 : U^b(t) = 0\}$ is the ruin time and $q \ge 0$ is a given discount rate. Here we adopt the convention that \mathbb{E}_x is the expectation with respect to \mathbb{P}_x . We derive a differential-delay equation for v(x). However, such differential-delay equations are notoriously difficult to solve, and we have not been able to solve our equation. Hence we have developed the following approach. Under the additional assumption that $\{N(t), t \ge 0\}$ is a Poisson process and all C_i equal zero, we find the expected cumulative discounted dividends

$$\nu_N(x) := \mathbb{E}_x \left[\int_0^{\tau_x^b(N)} \mathrm{e}^{-qt} \mathrm{d}L^b(t) \right], \tag{1.6}$$

paid under the barrier strategy until the process reaches $\frac{b}{(a+1)^N}$, that is, up to $\tau_x^b(N) := \inf\{t \ge 0 : U^b(t) = \frac{b}{(a+1)^N}\}$. By taking *N* sufficiently large we can approximate the value function v(x) closely by $v_N(x)$. To find $v_N(x)$ we first develop a method for solving a two-sided exit problem. Defining d_n as the first time that U^b reaches (down-crosses) $\frac{b}{(a+1)^n}$ and u_n as the first time U^b up-crosses $\frac{b}{(a+1)^n}$, we determine (with 1. denoting an indicator function)

$$\rho_N(x) := \mathbb{E}_x \Big[e^{-qd_N} \mathbf{1}_{d_N < u_0} \Big], \tag{1.7}$$

which seems to be of interest in its own right. We then use a very similar method to find $v_N(x)$, and to also solve a second two-sided exit problem, determining

$$\mu_N(x) := \mathbb{E}_x \Big[e^{-qu_0} \mathbf{1}_{u_0 < d_N} \Big].$$
(1.8)

This method of approximating a function $v(\cdot)$ by a sequence of functions $v_N(\cdot)$ with a certain recursive structure may be applicable in quite a few other settings, in particular in dam and storage models. In Section 5 we perform a similar analysis for the risk process (1.1) perturbed by an independent Brownian motion. There we also use the fluctuation theory of spectrally negative Lévy processes, expressing the exit identities in terms of so-called scale functions, as presented for example in Kyprianou [17].

As Avanzi et al. in [5] point out, 'Whereas [a classical model] is appropriate for an insurance company, [a dual model] seems to be natural for companies that have occasional gains. For companies such as pharmaceutical or petroleum companies, the jump should be interpreted as the net present value of future income from an invention or discovery. Other examples are commission-based businesses, such as real estate agent offices or brokerage firms that sell mutual funds or insurance products with a front-end load.' Avanzi et al. [5] also suggest possible applications in modeling an annuity or pension fund, where the risk consists of survival and the event of death leads to gains. More precisely, in the model (1.1), a company which continuously pays expenses relevant to research or labor and operational costs occasionally gains some random income from selling a product, invention, or discovery; see e.g. [4, 3, 8, 19, 24, 25]. Lately, the budgets of many start-ups or e-companies have shown a different feature. Namely, their gains are not additive but depend strongly on the amount of investments, which usually are so huge that they are proportional to the value of the company. Then the arrival gain is proportional not only to the investments but also to the value of the company. Perhaps the most transparent case is the example of CD Projekt, one of the biggest Polish companies producing computer games. The issuance of new editions of its most famous game, Witcher, produces jumps in the value of the company (which are translated into jumps of asset value), and these jumps are proportional to the prior jump position of the value process; see Figure 1.

Related literature. Not many papers consider the ruin probability for the classical dual risk process (without proportional gain mechanism), but it corresponds to the first busy period in a single-server queue with initial workload *x*, and consequently we can refer to [14, 21]. If the interarrival time has an exponential distribution, then one can apply fluctuation theory of Lévy processes to identify the Laplace transform of the ruin time as well; see e.g. Kyprianou [17]. Albrecher *et al.* [2] study the ruin probability in the dual risk model under a loss-carryforward tax system and assuming exponentially distributed jump sizes. Palmowski *et al.* [20] focus on a discrete-time set-up and study the finite-time ruin probability. In terms of analysis technique, the approach in Sections 2 and 3 bears similarities to the approach used in [12, 10, 11, 23] to

CD Projekt SA

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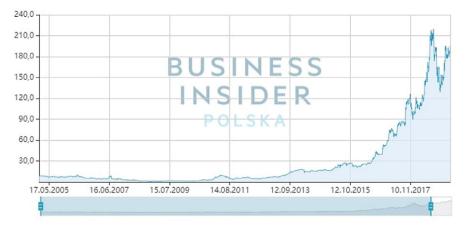


FIGURE 1. CD Projekt asset value; see https://businessinsider.com.pl.

study Lindley-type recursions $W_{n+1} = \max(0, aW_n + X_n)$, where a = 1 in the classical setting of a single-server queue with W_n the waiting time of the *n*th customer.

There is a good deal of work on dividend barriers in the dual model. All of those papers assume that the cost function is constant, and gains are modeled by a compound Poisson process. Avanzi et al. [5] consider cases where profits or gains follow an exponential distribution or a mixture of exponential distributions, and they derive explicit formulas for the expected discounted dividend value; see also Afonso et al. [1]. Avanzi and Gerber [4] use the Laplace transform method to study a dual model perturbed by a diffusion. Bayraktar et al. [8] and Avanzi et al. [6] employ fluctuation theory to prove the optimality of a barrier strategy for all spectrally positive Lévy processes and express the value function in terms of scale functions. Yin et al. [24, 25] consider terminal costs and dividends that are paid continuously at a constant rate (that might be bounded from above) when the surplus is above that barrier; see also Ng [19] for similar considerations. Albrecher *et al.* [2] examine a dual risk model in the presence of tax payments. Marciniak and Palmowski [18] consider a more general dual risk process where the rate of the costs depends on the present amount of reserves. Boxma and Frostig [9] consider the time to ruin and the expected discounted dividends for a different dividend policy, where a certain part of the gain is paid as dividends if upon arrival the gain finds the surplus above a barrier b or if it would bring the surplus above that level.

Organization of the paper. Section 2 is devoted to the determination of the ruin probability, while Section 3 considers the law of the ruin time. Section 4 considers two-sided exit problems that allow one to find the ruin probability and the total discounted dividend payments for the special case that the only capital growth is proportional growth. In Section 5 we handle the Brownian perturbation of the risk process (1.1). Section 6 contains suggestions for further research.

2. The ruin probability

In this section we determine the Laplace transform of the ruin probability R(x) when starting in *x*, as defined in (1.2). By distinguishing the two cases in which no jump up occurs before

x (hence ruin occurs at time *x*) and in which a jump up occurs at some time $t \in (0, x)$, we can write

$$R(x) = 1 - F(x) + \int_{t=0}^{x} \int_{y=0}^{\infty} R((1+a)(x-t) + y)\mu e^{-\mu y} dy dF(t).$$
(2.1)

Letting the Laplace transform

$$\rho(s) := \int_{x=0}^{\infty} e^{-sx} R(x) dx,$$

we have

$$\rho(s) = \frac{1 - \phi(s)}{s} + \int_{x=0}^{\infty} e^{-sx} \int_{t=0}^{x} \int_{z=(1+a)(x-t)}^{\infty} R(z)\mu e^{-\mu z} e^{\mu(1+a)(x-t)} dz dF(t) dx.$$
(2.2)

The triple integral in the right-hand side of (2.2), I(s), can be rewritten as follows:

$$I(s) = \int_{t=0}^{\infty} e^{-st} \int_{x=t}^{\infty} e^{-s(x-t)} e^{\mu(1+a)(x-t)} \int_{z=(1+a)(x-t)}^{\infty} \mu e^{-\mu z} R(z) dz dx dF(t)$$

= $\phi(s) \int_{v=0}^{\infty} e^{-sv + \mu(1+a)v} \int_{z=(1+a)v}^{\infty} \mu e^{-\mu z} R(z) dz dv$
= $\phi(s) \int_{z=0}^{\infty} \mu e^{-\mu z} R(z) \frac{e^{(\mu(1+a)-s)\frac{z}{1+a}} - 1}{\mu(1+a) - s} dz.$

Hence

$$\rho(s) = \frac{1 - \phi(s)}{s} + \phi(s) \frac{\mu}{\mu(1 + a) - s} \left[\rho\left(\frac{s}{1 + a}\right) - \rho(\mu) \right].$$
 (2.3)

Introducing

$$H(s) := \frac{1 - \phi(s)}{s} - \phi(s) \frac{\mu}{\mu(1+a) - s} \rho(\mu), \ J(s) := \phi(s) \frac{\mu}{\mu(1+a) - s}, \tag{2.4}$$

we rewrite (2.3) as

$$\rho(s) = J(s)\rho\left(\frac{s}{1+a}\right) + H(s).$$
(2.5)

Thus $\rho(s)$ is expressed in terms of $\rho\left(\frac{s}{1+a}\right)$, and after N-1 iterations this results in

$$\rho(s) = \sum_{k=0}^{N-1} \prod_{j=0}^{k-1} J\left(\frac{s}{(1+a)^j}\right) H\left(\frac{s}{(1+a)^k}\right) + \rho\left(\frac{s}{(1+a)^N}\right) \prod_{j=0}^{N-1} J\left(\frac{s}{(1+a)^j}\right), \quad (2.6)$$

with an empty product being equal to 1. Observe that, for large k, $H\left(\frac{s}{(1+a)^k}\right)$ approaches some constant and $J\left(\frac{s}{(1+a)^k}\right)$ approaches

$$\frac{\phi(0)}{1+a} = \frac{1}{1+a} < 1.$$

Hence the $\sum_{k=0}^{N-1} \prod_{j=0}^{k-1}$ term in (2.6) converges geometrically fast, and we obtain

$$\rho(s) = \sum_{k=0}^{\infty} \prod_{j=0}^{k-1} J\left(\frac{s}{(1+a)^j}\right) H\left(\frac{s}{(1+a)^k}\right).$$
(2.7)

Note that $\rho(\mu)$, featuring in the expression for H(s), is still unknown. Taking $s = \mu$ in (2.7) gives

$$\rho(\mu) = \sum_{k=0}^{\infty} \left(\prod_{j=0}^{k-1} J\left(\frac{\mu}{(1+a)^{j}}\right) \right) \left[\frac{1 - \phi\left(\frac{\mu}{(1+a)^{k}}\right)}{\frac{\mu}{(1+a)^{k}}} - \phi\left(\frac{\mu}{(1+a)^{k}}\right) \frac{\mu}{\mu(1+a) - \frac{\mu}{(1+a)^{k}}} \rho(\mu) \right],$$

and hence

$$\rho(\mu) = \frac{\sum_{k=0}^{\infty} \left(\prod_{j=0}^{k-1} J\left(\frac{\mu}{(1+a)^{j}}\right) \right)^{\frac{1-\phi\left(\frac{\mu}{(1+a)^{k}}\right)}{\frac{\mu}{(1+a)^{k}}}}{1 + \sum_{k=0}^{\infty} \left(\prod_{j=0}^{k-1} J\left(\frac{\mu}{(1+a)^{j}}\right) \right) \phi\left(\frac{\mu}{(1+a)^{k}}\right)^{\frac{(1+a)^{k}}{(1+a)^{k+1}-1}}.$$
(2.8)

We can sum up our analysis in the following first main result.

Theorem 2.1. The Laplace transform of the ruin probability, $\rho(s) = \int_0^\infty e^{-sx} \mathbb{P}_x(\tau_x < \infty) dx$, is given in (2.7) with H and J given in (2.4), where $\rho(\mu)$ is identified in (2.8).

Remark 2.1. It should be noticed that $R(x) \equiv 1$ satisfies Equation (2.1), but this trivial solution is not always the ruin probability. In fact, defining $X_n := U(T_n -)$, the surplus just before the *n*th jump epoch, the discrete Markov chain $\{X_n, n \ge 1\}$ satisfies the affine recursion

$$X_n = (a+1)X_{n-1} + (C_n - (T_n - T_{n-1})).$$

If a > 0, then from [13, Theorem 2.1.3, p. 13] we have that with a strictly positive probability X_n tends to $+\infty$. Thus R(x) < 1 if a > 0. If a = 0 then we are facing a G/M/1 queue, whose busy period ends with probability one if and only if $-\phi'(0) \ge \frac{1}{\mu}$.

Remark 2.2. Both H(s) and J(s) have a singularity at $s = \mu(1 + a)$, which suggests that the expression for $\rho(s)$ in (2.7) has a singularity for every $s = \mu(1 + a)^{j+1}$, j = 0, 1, ... However, $s = \mu(1 + a)$ is a removable singularity, as is already suggested by the form of (2.3), where $s = \mu(1 + a)$ also is a removable singularity. To verify formally that $s = \mu(1 + a)$ is not a singularity of (2.7), we proceed as follows (the same procedure can be applied for $s = \mu(1 + a)^{j+1}$, j = 1, 2, ...). Isolate the coefficients of the factor $\frac{1}{\mu(1+a)-s}$ in (2.7). Their sum C(s) equals

$$C(s) := -\phi(s)\mu\rho(\mu) + \phi(s)\mu \sum_{k=1}^{\infty} \prod_{j=1}^{k-1} J\left(\frac{s}{(1+a)^j}\right) H\left(\frac{s}{(1+a)^k}\right).$$

Introducing $k_1 := k - 1$ and $j_1 := j - 1$, and using (2.7), we readily see that $C(\mu(1 + a)) = 0$. **Remark 2.3.** For the case of Poisson arrivals, taking $F(x) = 1 - e^{-\lambda x}$, one gets a specific form for $\rho(s)$, indicating that R(x) is a weighted sum of exponential terms.

More precisely, in this case $\phi(s) = \frac{\lambda}{\lambda + s}$ and then

$$J(s) = \frac{1}{1+a} \frac{\lambda}{\lambda+s} \frac{\mu(1+a)}{\mu(1+a)-s}, \qquad H(s) = \frac{1}{\lambda+s} - \frac{\lambda}{\lambda+s} \frac{\mu(1+a)}{\mu(1+a)-s} \frac{\rho(\mu)}{1+a}.$$

From (2.5) it follows that

$$\rho(s) = H(s) + J(s)H\left(\frac{s}{1+a}\right) + J(s)J\left(\frac{s}{1+a}\right)H\left(\frac{s}{(1+a)^2}\right)$$
$$+ \dots + \prod_{j=0}^{k-1} J\left(\frac{s}{(1+a)^j}\right)H\left(\frac{s}{(1+a)^k}\right) + \dots$$
(2.9)

Observe that for $k \in \mathbb{N}$,

$$\prod_{j=0}^{k-1} J\left(\frac{s}{(1+a)^j}\right) = \frac{1}{(1+a)^k} \prod_{j=0}^{k-1} \frac{\lambda(1+a)^j}{\lambda(1+a)^j + s} \prod_{j=0}^{k-1} \frac{\mu(1+a)^{j+1}}{\mu(1+a)^{j+1} - s}.$$
 (2.10)

Note the following:

- the first term $\frac{1}{(1+a)^k}$ on the right-hand side of (2.10) gives a geometric decay in (2.9);
- the first product on the right-hand side of (2.10) is the LST of the sum $E_1 + E_2 + ... + E_k$ of independent random variables E_j (j = 1, 2, ..., k) with exponential distributions with parameter $\lambda(1 + a)^j$; that is, it is the LST of a hyper-exponential distribution;
- the second product represents the LST of $-(F_1 + F_2 + ... + F_k)$, where F_j , j = 1, 2, ..., k, are independent random variables with exponential distributions with parameter $\mu(1+a)^{j+1}$;
- hence $\prod_{j=0}^{k-1} J\left(\frac{s}{(1+a)^j}\right)$ describes the LST of the sum of the above random variables weighted by the geometric term.

Moreover, the term $H\left(\frac{s}{(1+a)^k}\right)$ appearing in (2.9) is the LST of

$$(1+a)^k \mathbb{P}(E_k > t) - \frac{\rho(\mu)}{1+a}\tilde{f}(t),$$

where $\tilde{f}(t)$ is the density of $E_k - F_k$, with E_k and F_k being independent.

In fact, this can be further generalized to general interarrival times using the representation (2.9) of $\rho(s)$. Indeed, let *T* be a generic interarrival time between jumps. Then $J\left(\frac{s}{(1+a)^j}\right)$ is the LST of the density of $\frac{T}{(1+a)^j} - F_j$ multiplied by $\frac{1}{1+a}$, where *T* and F_j are independent. Similarly, $H\left(\frac{s}{(1+a)^k}\right)$ is the LST of

$$(1+a)^k \mathbb{P}\left(\frac{T}{(1+a)^k} > t\right) - \frac{\rho(\mu)}{1+a}\tilde{f}(t),$$

where $\tilde{f}(t)$ is the density of the residual of $\frac{T}{(1+a)^k}$ minus F_k :

$$\tilde{f}(t) = \phi(\mu(1+a))\mu(1+a)^{k+1}e^{\mu(1+a)^{k+1}t}\mathbf{1}_{t<0} + \int_0^\infty \mu(1+a)e^{-\mu(1+a)y}f(y+t/(1+a)^k)dy\mathbf{1}_{t>0}$$

To get this observation we use the identity

$$\begin{split} \phi\bigg(\frac{s}{(1+a)^k}\bigg) & \frac{\mu(1+a)^{k+1}}{\mu(1+a)^{k+1}-s} = \phi(\mu(1+a))\frac{\mu(1+a)^{k+1}}{\mu(1+a)^{k+1}-s} \\ & + \mu(1+a)\frac{\phi\bigg(\frac{s}{(1+a)^k}\bigg) - \phi(\mu(1+a))}{\mu(1+a) - s/(1+a)^k} \end{split}$$

and note that taking the LST of $\phi(\mu(1+a))\mu(1+a)^{k+1}e^{\mu(1+a)^{k+1}t}\mathbf{1}_{t<0}$ gives the first increment on the right-hand side of the above equality, while taking the LST of $\int_0^\infty \mu(1+a)e^{-\mu(1+a)y}f(y+t/(1+a)^k)dy\mathbf{1}_{t>0}$ produces the second increment there.

Remark 2.4. We now make a few comments about possible generalizations. We could allow a hyper-exponential-*K* distribution for *C*, leading to *K* unknowns $\rho(\mu_1), \ldots, \rho(\mu_K)$ which can be found by taking $s = \mu_1, \ldots, s = \mu_K$.

We could also consider the following generalization of the model (1.1): when the *i*th jump upwards occurs while $U(T_i-) = u$, that jump has size $au + C_i$ with probability p, and has size D_i with probability 1 - p, where D_1, D_2, \ldots are independent, $\exp(\delta)$ -distributed random variables, independent of C_1, C_2, \ldots . By taking p = 1 we get the old model, while a = 0, $\mu = \delta$ gives a classical dual risk model. It is readily verified that, for this generalized model, (2.3) becomes

$$\rho(s) = \frac{1 - \phi(s)}{s} + p\phi(s)\frac{\mu}{\mu(1+a) - s} \left[\rho\left(\frac{s}{1+a}\right) - \rho(\mu)\right]$$
$$+ (1-p)\phi(s)\frac{\delta}{\delta - s}[\rho(s) - \rho(\delta)]. \tag{2.11}$$

Introducing

$$H_{1}(s) := \frac{\frac{1-\phi(s)}{s} - p\phi(s)\frac{\mu}{\mu(1+a)-s}\rho(\mu) - (1-p)\phi(s)\frac{\delta}{\delta-s}\rho(\delta)}{1 - (1-p)\frac{\delta}{\delta-s}\phi(s)} \\ J_{1}(s) := \frac{p\phi(s)\frac{\mu}{\mu(1+a)-s}}{1 - (1-p)\frac{\delta}{\delta-s}\phi(s)},$$

we rewrite (2.11) as

$$\rho(s) = J_1(s)\rho\left(\frac{s}{1+a}\right) + H_1(s),$$

resulting in

$$\rho(s) = \sum_{k=0}^{\infty} \prod_{j=0}^{k-1} J_1\left(\frac{s}{(1+a)^j}\right) H_1\left(\frac{s}{(1+a)^k}\right).$$
(2.12)

Finally, $\rho(\mu)$ and $\rho(\delta)$ have to be determined. One equation is supplied by substituting $s = \mu$ in (2.12) (just as was done below (2.7)). For a second equation we invoke Rouché's theorem, which implies that, for any $p \in (0, 1)$, the equation $\delta - s - (1 - p)\delta\phi(s) = 0$ has exactly one zero, say s_1 , in the right half *s*-plane. Observing that $\rho(s)$ is analytic in that half-plane, so that $\rho(s_1)$ is finite, it follows from (2.11) that

$$\frac{1-\phi(s_1)}{s_1} + p\phi(s_1)\frac{\mu}{\mu(1+a) - s_1} \left[\rho\left(\frac{s_1}{1+a}\right) - \rho(\mu)\right] - (1-p)\frac{\delta}{\delta - s_1}\phi(s_1)\rho(\delta) = 0.$$

While this provides a second equation, it also introduces a third unknown, viz., $\rho(\frac{s_1}{1+a})$. However, substituting $s = \frac{s_1}{1+a}$ in (2.12) expresses $\rho(\frac{s_1}{1+a})$ using $\rho(\mu)$ and $\rho(\delta)$, thus providing a third equation.

3. The time to ruin

In this section we study the distribution of τ_x , the time to ruin when starting at level *x*, as defined in (1.3). Following a similar approach as in the previous section, again distinguishing between the first upward jump occurring before or after *x*, we can write

$$\mathbb{E}[e^{-\alpha\tau_x}] = e^{-\alpha x}(1 - F(x)) + \int_{t=0}^x \int_{y=0}^\infty e^{-\alpha y} \mathbb{E}[e^{-\alpha\tau_{(1+a)(x-t)+y}}]\mu e^{-\mu y} dy dF(t).$$

Remark 3.1. Taking $\alpha = 0$ yields the ruin probability R(x). In that respect, it would not have been necessary to present a separate analysis of R(x); however, to improve the readability of the paper, we have chosen to demonstrate the analysis technique first for the easier case of R(x).

Introducing the Laplace transform

$$\tau(s,\alpha) := \int_{x=0}^{\infty} e^{-sx} \mathbb{E}\left[e^{-\alpha\tau_x}\right] \mathrm{d}x,$$

and using calculations very similar to those leading to (2.3), we obtain

$$\tau(s,\alpha) = \frac{1-\phi(s+\alpha)}{s+\alpha} + \phi(s+\alpha)\frac{\mu}{\mu(1+\alpha)-s} \left[\tau\left(\frac{s}{1+\alpha},\alpha\right) - \tau(\mu,\alpha)\right].$$
 (3.1)

Introducing

$$H_{1}(s,\alpha) := \frac{1-\phi(s+\alpha)}{s+\alpha} - \phi(s+\alpha)\frac{\mu}{\mu(1+a)-s}\tau(\mu,\alpha), \ J_{1}(s,\alpha) := \phi(s+\alpha)\frac{\mu}{\mu(1+a)-s}$$
(3.2)

we rewrite (3.1) as

$$\tau(s,\alpha) = J_1(s,\alpha)\tau\left(\frac{s}{1+a},\alpha\right) + H_1(s,\alpha),$$

which after N - 1 iterations yields the following (an empty product being equal to 1):

$$\tau(s,\alpha) = \sum_{k=0}^{N-1} \prod_{j=0}^{k-1} J_1\left(\frac{s}{(1+a)^j},\alpha\right) H_1\left(\frac{s}{(1+a)^k},\alpha\right) + \tau\left(\frac{s}{(1+a)^N},\alpha\right) \prod_{j=0}^{N-1} J_1\left(\frac{s}{(1+a)^j},\alpha\right).$$
(3.3)

Observe that, for large k, $H_1\left(\frac{s}{(1+a)^k}, \alpha\right)$ approaches some function of α and $J_1\left(\frac{s}{(1+a)^k}, \alpha\right)$ approaches $\frac{\phi(\alpha)}{1+a} < 1$. Hence the $\sum_{k=0}^{N-1} \prod_{j=0}^{k-1}$ term in (3.3) converges geometrically fast, and we obtain

$$\tau(s,\alpha) = \sum_{k=0}^{\infty} \prod_{j=0}^{k-1} J_1\left(\frac{s}{(1+a)^j},\alpha\right) H_1\left(\frac{s}{(1+a)^k},\alpha\right).$$
 (3.4)

Now, $\tau(\mu, \alpha)$, featuring in the expression for $H_1(s, \alpha)$, is still unknown. Taking $s = \mu$ in (3.4) gives

$$\tau(\mu, \alpha) = \sum_{k=0}^{\infty} \left(\prod_{j=0}^{k-1} J_1\left(\frac{\mu}{(1+a)^j}, \alpha\right) \right)$$
$$\left[\frac{1 - \phi\left(\frac{\mu}{(1+a)^k} + \alpha\right)}{\frac{\mu}{(1+a)^k} + \alpha} - \phi\left(\frac{\mu}{(1+a)^k} + \alpha\right) \frac{\mu}{\mu(1+a) - \frac{\mu}{(1+a)^k}} \tau(\mu, \alpha) \right],$$

and hence

$$\tau(\mu, \alpha) = \frac{\sum_{k=0}^{\infty} \left(\prod_{j=0}^{k-1} J_1\left(\frac{\mu}{(1+a)^j}, \alpha\right)\right) \frac{1-\phi\left(\frac{\mu}{(1+a)^k} + \alpha\right)}{\frac{\mu}{(1+a)^k} + \alpha}}{1 + \sum_{k=0}^{\infty} \left(\prod_{j=0}^{k-1} J_1\left(\frac{\mu}{(1+a)^j}\right)\right) \phi\left(\frac{\mu}{(1+a)^k} + \alpha\right) \frac{(1+a)^k}{(1+a)^{k+1} - 1}}.$$
(3.5)

Thus we have proved the following, which is the second main result of this paper.

Theorem 3.1. The double Laplace transform (with respect to time and initial conditions) $\tau(s, \alpha) = \int_{x=0}^{\infty} e^{-sx} \mathbb{E}[e^{-\alpha \tau_x}] dx$ is given in (3.4) with H_1 and J_1 given in (3.2), with $\tau(\mu, \alpha)$ identified in (3.5).

4. Exit problems, ruin, and barrier dividend value function

In this section we consider the same model as in the previous sections, but with the restriction that the only growth is a proportional growth occurring according to a Poisson process at rate λ ; throughout this section we further assume that $C_i \equiv 0$. In other words,

$$U(t) = x - t + a \sum_{i=1}^{N(t)} U(T_i), \ t \ge 0,$$

and N(t) is a Poisson process with intensity $\lambda > 0$. In Subsection 4.1 we solve the two-sided downward exit problem for this model. In Subsection 4.2 we use a similar method to determine the discounted cumulative dividend payments paid up to the ruin time under the barrier strategy with barrier *b* and with a discount rate *q*. But first we briefly discuss an alternative, more straightforward, approach to determining the discounted cumulative dividend payments, pointing out why this approach does not work. We start from the observation that for x > b we have

$$v(x) = v(b) + x - b.$$
(4.1)

We now focus on $x \le b$. One-step analysis based on the first arrival epoch gives

$$v(x) = \int_0^x \lambda e^{-(\lambda+q)t} v((x-t)(1+a)) dt, \ 0 \le x < b,$$
(4.2)

and by taking z := x - t and taking the derivative with respect to x we end up with the equation

$$v'(x) + (\lambda + q)v(x) = \lambda v(x(1+a)), \qquad x \le \frac{b}{1+a}.$$
 (4.3)

Moreover, from (4.1) we have

$$v'(x) + (\lambda + q)v(x) = \lambda(x(1 + a) - b + v(b)), \qquad \frac{b}{1 + a} < x \le b,$$

which can easily be solved. Unfortunately, differential-delay equations like (4.3) seem hard to solve explicitly; cf. [15]. As an alternative, one could try to solve the equation numerically.

Instead, we adopt a different approach, in which we distinguish levels $L_n := \frac{b}{(a+1)^n}$ and assume that ruin occurs when the level $\frac{b}{(a+1)^N}$ is reached for some value of N. When N is large, the expected amount of discounted cumulative dividends closely approximates the expected amount until ruin at zero occurs. The above choice of levels is suitable because each proportional jump upward brings the process from a value in (L_{n+1}, L_n) to a value in (L_n, L_{n-1}) .

4.1. Two-sided downward exit problem and ruin time

Recall that d_n is the first time that U reaches (down-crosses) $\frac{b}{(a+1)^n}$, and u_n is the first time the risk process up-crosses $\frac{b}{(a+1)^n}$. Hence u_0 is the time until level b is first up-crossed, i.e., a dividend is being paid. For an integer N we aim to obtain the Laplace transform $\rho_N(x) = \mathbb{E}_x \left[e^{-qd_N} \mathbf{1}_{d_N < u_0} \right]$ defined in (1.7). For large N, $\rho_N(x)$ approximates the LST of the time until ruin in the event that no dividend is ever paid. If q = 0, then $\rho_N(x)$ approximates the probability that ruin occurs before dividends are paid, that is, before reaching b. If q = 0 and N and b are both tending to infinity, then $\rho_N(x)$ approximates the ruin probability R(x) defined in (1.2). For $1 \le n \le N$ let

$$\rho_{n,N} := \rho_N\left(\frac{b}{(a+1)^n}\right).$$

To simplify notation we define $\rho_n := \rho_{n,N}$. Clearly, $\rho_N = 1$. As mentioned above, we consider levels

$$L_n := \frac{b}{(a+1)^n}, \quad n = 0, 1, \dots, N;$$

let

$$\mathcal{N}:=\{L_1,\ldots,L_N\}.$$

To determine $\rho_N(x)$ when $x \in \left(\frac{b}{(a+1)^n}, \frac{b}{(a+1)^{n-1}}\right] = (L_n, L_{n-1}]$, note that, because $d_n < u_0$, there must be a down-crossing of level L_n before level b is up-crossed. We can now distinguish n different possibilities: when starting at x, the surplus process first decreases through L_n , or it first increases via jumps above level L_{n-j} before there is a first down-crossing through that same level L_{n-j} , $j = 1, \ldots, n-1$. Denoting the time for the former event by

$$T_{n,0} := d_n \mathbf{1}_{d_n < u_{n-1}},$$

and the times for the latter n - 1 events by

$$T_{n,j} := u_{n-1} \mathbf{1}_{u_{n-1} < d_n} + u_{n-2} \mathbf{1}_{u_{n-2} < d_{n-1}} + \ldots + u_{n-j} \mathbf{1}_{u_{n-j} < d_{n-j+1}} + d_{n-j} \mathbf{1}_{d_{n-j} < u_{n-j-1}},$$

for j = 1, ..., n - 1, we can derive the following representation of ρ_N . It will turn out to be useful to introduce

$$\overline{G}_c(t) = e^{-(\lambda+q)ct}$$
 and $g_c(t) = \lambda c e^{-(\lambda+q)ct}$.

Theorem 4.1. For $\frac{b}{(a+1)^n} < x \le \frac{b}{(a+1)^{n-1}}$,

$$\rho_{N}(x) = \bar{G}_{1}\left(x - \frac{b}{(a+1)^{n}}\right)\rho_{n}$$

$$+ \bar{G}_{(a+1)} \circledast g_{1}\left(x - \frac{b}{(a+1)^{n}}\right)\rho_{n-1}$$

$$+ \bar{G}_{(a+1)^{2}} \circledast g_{(a+1)} \circledast g_{1}\left(x - \frac{b}{(a+1)^{n}}\right)\rho_{n-2}$$

$$+ \dots$$

$$+ \bar{G}_{(a+1)^{n-1}} \circledast g_{(a+1)^{n-2}} \circledast \dots \circledast g_{1}\left(x - \frac{b}{(a+1)^{n}}\right)\rho_{1}, \qquad (4.4)$$

where \circledast denotes convolution.

Proof. Let $\frac{b}{(a+1)^n} < x \le \frac{b}{(a+1)^{n-1}}$. The *n* terms in the right-hand side of (4.4) represent the *n* disjoint possibilities where $d_n < u_0$. Notice that $T_{n,j}$ is the first time that the process *U* reaches a level in \mathcal{N} via a down-crossing, by reaching $\frac{b}{(a+1)^{n-j}}$. Furthermore, ρ_{n-j} is the Laplace transform of the time to reach $\frac{b}{(a+1)^N}$ starting at $\frac{b}{(a+1)^{n-j}}$. Now first considering $T_{n,0}$, we have

$$E_{x}\left[e^{-qd_{n}}1_{d_{n}< u_{n-1}}\right] = e^{-(q+\lambda)\left(x-\frac{b}{(a+1)^{n}}\right)} = \bar{G}_{1}\left(x-\frac{b}{(a+1)^{n}}\right).$$

By the strong Markov property, considering $T_{n,j}$, we derive

$$\mathbb{E}_{x}\left[e^{-q\left(u_{n-1}+\dots+u_{n-j}+d_{n-j}\right)}1_{u_{n-1}< d_{n},\dots,u_{n-j}< d_{n-j+1},d_{n-j}< u_{n-j-1}}\right] = \int_{t_{1}=0}^{A_{n,0}}\int_{t_{2}=0}^{A_{n,1}}\dots\int_{t_{j}=0}^{A_{n,j-1}}\bar{G}_{1}(A_{n,j})g_{1}(t_{j})g_{1}(t_{j-1})\dots g_{1}(t_{1})dt_{j}\dots dt_{1},$$

where

$$A_{n,0} := x - \frac{b}{(a+1)^n},$$

and for k = 1, 2, ..., n,

$$A_{n,k} := (a+1)(A_{n,k-1} - t_k), \tag{4.5}$$

with t_k an integration variable. By the change of variables $y_j = t_j/(a+1)^{j-1}$, we obtain that

$$\mathbb{E}_{x}\left[e^{-q\left(u_{n-1}+\dots+u_{n-j}+d_{n-j}\right)}1_{u_{n-1}< d_{n},\dots,u_{n-j}< d_{n-j+1},d_{n-j}< u_{n-j-1}}\right] = \bar{G}_{(a+1)^{j}} \circledast g_{(a+1)^{j-1}} \circledast \dots \circledast g_{1}\left(x-\frac{b}{(a+1)^{n}}\right),$$

and the theorem follows.

From Theorem 4.1 it follows that to obtain ρ_n we need to solve the following N-1 equations for n = 1, 2, ..., N-1:

$$\begin{split} \rho_n &= \bar{G}_1 \left(\frac{b}{(a+1)^n} - \frac{b}{(a+1)^{n+1}} \right) \rho_{n+1} \\ &+ \bar{G}_{(a+1)} \circledast g_1 \left(\frac{b}{(a+1)^n} - \frac{b}{(a+1)^{n+1}} \right) \rho_n \\ &+ \bar{G}_{(a+1)^2} \circledast g_{(a+1)} \circledast g_1 \left(\frac{b}{(a+1)^n} - \frac{b}{(a+1)^{n+1}} \right) \rho_{n-1} \\ &+ \dots \\ &+ \bar{G}_{(a+1)^n} \circledast g_{(a+1)^{n-1}} \circledast \dots \circledast g_1 \left(\frac{b}{(a+1)^n} - \frac{b}{(a+1)^{n+1}} \right) \rho_1. \end{split}$$

Defining

$$\gamma_0(x) := \bar{G}_1(x),$$

and for $n \ge 1$,

$$\gamma_n(x) := \overline{G}_{(a+1)^n} \circledast g_{(a+1)^{n-1}} \circledast \ldots \circledast g_1(x)$$

by the formula for convolution of exponentials given in [22, Chapter 5] we have for $n \ge 1$

$$\gamma_n(x) = \left(\frac{\lambda}{\lambda+q}\right)^n \sum_{i=0}^n \frac{e^{-(\lambda+q)(a+1)^i x}}{\prod_{j\neq i} ((a+1)^j - (a+1)^i)}.$$

Thus we have the following set of linear equations for n = 1, 2, ..., N - 1:

$$\rho_n = \sum_{j=0}^n \gamma_j \left(\frac{b}{(a+1)^n} \left(1 - \frac{1}{a+1} \right) \right) \rho_{n+1-j}.$$

Notice that $\rho_N = 1$. The above formula can be rewritten as follows:

$$\rho_n = \sum_{j=0}^n \gamma_{j,n} \rho_{n+1-j} = \sum_{j=1}^{n+1} \gamma_{n+1-j,n} \rho_j, \qquad (4.6)$$

where

$$\gamma_{j,n} := \gamma_j \left(\frac{b}{(a+1)^n} \left(1 - \frac{1}{a+1} \right) \right).$$

Introducing the $(N-1) \times (N-1)$ matrix Γ , with as its *n*th row $(\gamma_{n,n}, \gamma_{n-1,n}, \dots, \gamma_{0,n}, 0, \dots, 0)$, and the column vector $\rho := (\rho_1, \dots, \rho_{N-1})^T$, we can write the set of equations (4.6) as

$$\rho = \Gamma \rho + Z,$$

where $Z = (0, ..., 0, \gamma_{0,N-1})^T$. Hence, with *I* the $(N-1) \times (N-1)$ matrix with ones on the diagonal and zeroes at all other positions,

$$\rho = (I - \Gamma)^{-1} Z.$$

4.2. Expected discounted dividends

Recall that $v_N(x)$ defined in (1.6) is the expected discounted dividends under the barrier strategy until reaching $\frac{b}{(a+1)^N}$, that is up to $\tau_x^b(N)$ for the regulated process $U^b(t)$ defined in (1.4). Note that

$$v(x) = \lim_{N \to +\infty} v_N(x)$$

for v(x) defined in (1.5). Let

$$v_n := v_N \left(\frac{b}{(a+1)^n}\right) \quad \text{for } n = 0, 1, \dots, N-1.$$
 (4.7)

The next theorem identifies $v_N(x)$.

Theorem 4.2. For $\frac{b}{(a+1)^n} < x \le \frac{b}{(a+1)^{n-1}}$,

$$v_{N}(x) = \bar{G}_{1}\left(x - \frac{b}{(a+1)^{n}}\right)v_{n}$$

$$+ \bar{G}_{(a+1)} \circledast g_{1}\left(x - \frac{b}{(a+1)^{n}}\right)v_{n-1}$$

$$+ \bar{G}_{(a+1)^{2}} \circledast g_{(a+1)} \circledast g_{1}\left(x - \frac{b}{(a+1)^{n}}\right)v_{n-2}$$

$$+ \dots$$

$$+ \bar{G}_{(a+1)^{n-1}} \circledast g_{(a+1)^{n-2}} \circledast \dots \circledast g_{1}\left(x - \frac{b}{(a+1)^{n}}\right)v_{1}$$

$$+ 1 \circledast g_{(a+1)^{n-1}} \circledast \dots \circledast g_{1}\left(x - \frac{b}{(a+1)^{n}}\right)v(b)$$

$$+ (a+1)^{n} \mathcal{Q} \circledast g_{(a+1)^{n-1}} \circledast \dots \circledast g_{1}\left(x - \frac{b}{(a+1)^{n}}\right), \qquad (4.8)$$

where Q(x) = x and \circledast again denotes convolution.

Proof. The proof follows exactly the same reasoning as the proof of Theorem 4.1: we consider again the disjoint events in which the first down-crossing of a level from \mathcal{N} occurs at $L_n - j$, $j = 0, \ldots, n - 1$. However, we now do not exclude the possibility that $L_0 = b$ is upcrossed before level L_N is reached. This gives rise to the last two lines of (4.8). More precisely, let $L_n < x \le L_{n-1}$ and let \mathcal{A}_n be the event that level $L_0 = b$ is upcrossed before one of the levels L_j , $j = 1, \ldots, N$, is downcrossed. This occurs when each of the subsequent *n* jumps occurs before down-crossing L_{n-j+1} , $j = 0, \ldots, n$, i.e. when $u_{n-j} < d_{n+1-j}$, $j = 1, \ldots, n$. For $L_n < x \le L_{n-1}$, the time to this event is

$$\Upsilon_n(x) := u_{n-1} \mathbf{1}_{u_{n-1} < d_n} + u_{n-2} \mathbf{1}_{u_{n-2} < d_{n-1}} + \ldots + u_1 \mathbf{1}_{u_1 < d_2} + u_0 \mathbf{1}_{u_0 < d_1}.$$

The Laplace transform of Υ_n (or the discounted time until A_n occurs), starting at *x* with $L_n < x \le L_{n-1}$, can be obtained by arguments similar to those leading to (4.6); that is,

$$\mathbb{E}_{x}\left[e^{-q\left(u_{n-1}+\dots+u_{0}\right)}1_{u_{n-1}< d_{n},\dots,u_{0}< d_{1}}\right]$$

= 1 * g_{(a+1)ⁿ⁻¹} * ... * g₁\left(x-\frac{b}{(a+1)^{n}}\right).

Once the process up-crosses b, a dividend is paid and the process restarts at level b. Thus the second-to-last line of (4.8) is the expected discounted dividends paid until ruin starting at time u_0 at level b (not including the dividends paid at this time). The amount of the expected discounted dividends paid at $u_{n-1} + \cdots + u_0$ is

$$\int_{t_1=0}^{A_{n,0}} \int_{t_2=0}^{A_{n,1}} \dots \int_{t_n=0}^{A_{n,n-1}} A_{n,n} g_1(t_n) g_1(t_{n-1}) \dots g_1(t_1) dt_n \dots dt_1$$

= $(a+1)^n \mathcal{Q} \circledast g_{(a+1)^{n-1}} \circledast \dots \circledast g_1 \left(x - \frac{b}{(a+1)^n} \right),$

where $A_{n,n}$ is defined in (4.5) and the last equality is obtained by change of variables $y_j = \frac{t_j}{(1+a)^{j-1}}, j = 1, 2, ..., n$.

It remains to determine $v_0 = v(b), v_1, \dots, v_{N-1}$, since then from Theorem 4.2 we have $v_N(x)$ for all $x \in (0, b]$. Notice that $v_N = 0$. We first derive an equation for v(b). Let

$$\delta_n := (a+1)^n \mathcal{Q} \circledast g_{(a+1)^{n-1}} \circledast \ldots \circledast g_1 \left(\frac{b}{(a+1)^n} \left(1 - \frac{1}{a+1} \right) \right)$$

and let

$$\omega_n := 1 \circledast g_{(a+1)^{n-1}} \circledast \dots \circledast g_1\left(\frac{b}{(a+1)^{n-1}}\left(1-\frac{1}{a+1}\right)\right)$$

We distinguish between the following cases, when starting from b: (i) level $L_1 = \frac{b}{a+1}$ is reached before b is up-crossed again; this gives rise to the first term in the right-hand side of (4.9) below; (ii) b is up-crossed before level L_1 is reached. Thus

$$v(b) = \bar{G}_1 \left(b - \frac{b}{a+1} \right) v \left(\frac{b}{a+1} \right) + \frac{\lambda}{\lambda+q} \left(1 - \bar{G}_1 \left(b - \frac{b}{a+1} \right) \right) v(b)$$
$$+ (a+1)\mathcal{Q} \circledast g_1 \left(b - \frac{b}{a+1} \right).$$
(4.9)

Notice that

$$\frac{\lambda}{\lambda+q}\left(1-\bar{G}_1\left(b-\frac{b}{a+1}\right)\right)=1 \circledast g_1\left(b-\frac{b}{a+1}\right)=\omega_1.$$

Hence

$$v_0 = \gamma_0 \left(b - \frac{b}{a+1} \right) v_1 + \omega_1 v_0 + \delta_1.$$

By taking $x = \frac{b}{(a+1)^{n-1}}$ in Theorem 4.2, we get for n = 1, ..., N-1,

$$v_n = \sum_{j=0}^n \gamma_j \left(\frac{b}{(a+1)^{n-1}} \left(1 - \frac{1}{a+1} \right) \right) v_{n+1-j} + \omega_{n+1} v_0 + \delta_{n+1}.$$
(4.10)

Introducing the column vector $V := (v_0, ..., v_{N-1})^T$, we can write the equation for v_0 and the set of equations (4.10) together as

$$V = \Psi V + \Delta,$$

where Ψ is an $N \times N$ matrix with row $n, n = 0, \dots, N-1$, equal to $(\omega_{n+1}, \gamma_{n,n}, \dots, \gamma_{0,n}, 0, \dots, 0)$. Notice that row N-1 is $(\omega_N, \gamma_{N-1,N-1}, \dots, \gamma_{1,N-1})$. Hence

$$V = (I - \Psi)^{-1} \Delta.$$

Remark 4.1. Our analysis can be used to solve the two-sided upward exit problem for our risk process as well. We recall that $\mu_n = \mu_N \left(\frac{b}{(a+1)^n}\right)$ and it is defined in (1.8). Then by the same arguments as those leading to (4.8) and (4.10), for $\frac{b}{(a+1)^n} < x \le \frac{b}{(a+1)^{n-1}}$, we obtain that

$$\begin{split} \mu_N(x) &= \bar{G}_1 \left(x - \frac{b}{(a+1)^n} \right) \mu_n \\ &+ \bar{G}_{(a+1)} \circledast g_1 \left(x - \frac{b}{(a+1)^n} \right) \mu_{n-1} \\ &+ \bar{G}_{(a+1)^2} \circledast g_{(a+1)} \circledast g_1 \left(x - \frac{b}{(a+1)^n} \right) \mu_{n-2} \\ &+ \dots \\ &+ \bar{G}_{(a+1)^{n-1}} \circledast g_{(a+1)^{n-2}} \circledast \dots \circledast g_1 \left(x - \frac{b}{(a+1)^n} \right) \mu_1 \\ &+ 1 \circledast g_{(a+1)^{n-1}} \circledast \dots \circledast g_1 \left(x - \frac{b}{(a+1)^n} \right). \end{split}$$

Similarly to the equation for v_0 and Equation (4.10), we then obtain barriers that can inform plans that

$$\mu_0 = \gamma_0 \left(b - \frac{b}{a+1} \right) \mu_1 + \omega_1, \tag{4.11}$$
$$\mu_n = \sum_{j=0}^n \gamma_j \left(\frac{b}{(a+1)^{n-1}} \left(1 - \frac{1}{a+1} \right) \right) \mu_{n+1-j} + \omega_{n+1}, \ n = 1, \dots, N-1.$$

Moreover, observe that $\mu_0 = \mu_N(b) = 1$.

5. Exit times and barrier dividends value function with Brownian perturbation

In this section we extend the model of Section 4 by allowing small perturbations between jumps. These perturbations are modeled by a Brownian motion X(t) with drift η and variance σ^2 ; that is,

$$X(t) = \eta t + \sigma B(t), \tag{5.1}$$

for a standard Brownian motion B(t). Hence our risk process is formally defined as

$$U(t) = x + X(t) + \sum_{i=1}^{N(t)} aU(T_i), \ t \ge 0,$$

where N(t) is a Poisson process with intensity $\lambda > 0$. We apply the fluctuation theory of onesided Lévy processes to solve the two-sided exit problem (Subsection 5.1) and to obtain the expected discounted barrier dividends (Subsection 5.2). The key functions for this fluctuation theory are the scale functions; see [17]. To introduce these functions let us first define the Laplace exponent of X(t):

$$\psi(\theta) := \frac{1}{t} \log \mathbb{E}_x \left[e^{\theta X_t} \right] = \eta \theta + \frac{\sigma^2}{2} \theta^2.$$

This function is strictly convex, is differentiable, equals zero at zero, and tends to infinity at infinity. Hence its right inverse $\Phi(q)$ exists for $q \ge 0$. The first scale function $W^{(q)}(x)$ is the unique right-continuous function disappearing on the negative half-line whose Laplace transform is

$$\int_0^\infty e^{-\theta x} W^{(q)}(x) \mathrm{d}x = \frac{1}{\psi(\theta) - q},$$

for $\theta > \Phi(q)$. With the first scale function we can associate a second scale function via $Z^{(q)}(x) := 1 - q \int_0^x W^{(q)}(y) dy$. In the case of linear Brownian motion as defined in (5.1), the (first) scale function for a Brownian motion with drift η and variance σ^2 equals (cf. [16])

$$W^{(q)}(x) = \frac{1}{\sqrt{\eta^2 + 2q\sigma^2}} \left[e^{\left(\sqrt{\eta^2 + 2q\sigma^2} - \eta\right)\frac{x}{\sigma^2}} - e^{-\left(\sqrt{\eta^2 + 2q\sigma^2} + \eta\right)\frac{x}{\sigma^2}} \right].$$

Let $\alpha < x < \beta$ and

$$d^X_{\alpha} = \min\{t : X(t) = \alpha\}$$
 and $u^X_{\beta} = \min\{t : X(t) = \beta\}$

Throughout this section we use the following three facts given in Theorem 8.1 and Theorem 8.7 of Kyprianou [17]:

1.

$$\mathbb{E}_{x}\left[e^{-qu_{\beta}^{X}}\mathbf{1}_{u_{\beta}^{X} < d_{\alpha}^{X}}\right] = \frac{W^{(q)}(x-\alpha)}{W^{(q)}(\beta-\alpha)}.$$
(5.2)

2.

$$\mathbb{E}_{x}\left[e^{-qd_{\alpha}^{X}}\mathbf{1}_{d_{\alpha}^{X} < u_{\beta}^{X}}\right] = Z^{(q)}(x-\alpha) - \frac{W^{(q)}(x-\alpha)}{W^{(q)}(\beta-\alpha)}Z^{(q)}(\beta-\alpha).$$
(5.3)

3. Let \mathcal{E}_q be an exponentially distributed random variable with parameter *q* independent of the process *X*. Then for $\alpha < x < \beta$,

$$\frac{\mathbb{P}_{x}\left(X(\mathcal{E}_{q})\in(y,\,y+\mathrm{d}y),\,\mathcal{E}_{q}< u_{\beta}^{X}\wedge d_{\alpha}^{X}\right)}{q\mathrm{d}y}=u_{\alpha,\beta}^{(q)}(x,\,y)$$
$$=\frac{W^{(q)}(x-\alpha)}{W^{(q)}(\beta-\alpha)}W^{(q)}(\beta-y)-W^{(q)}(x-y).$$
(5.4)

5.1. Downward exit problem and ruin time

In this subsection we obtain

$$\rho_N(x) = \mathbb{E}_x \Big[e^{-qd_N} \mathbf{1}_{d_N < u_0} \Big].$$
(5.5)

This is done in three steps. In Step 1 we determine the LST of the time needed, starting from some $x \in (L_n, L_{n-1})$, to reach a level in \mathcal{N} by down-crossing L_{n-k} , $k = 0, 1, \ldots, n-1$. In Step 2 we determine the LST of the time needed, starting from some $x \in (L_n, L_{n-1})$, to reach a level in \mathcal{N} by up-crossing L_{n-k} , $k = 1, 2, \ldots, n$. In Step 3 we express $\rho_N(x)$ in ρ_1, \ldots, ρ_N , with ρ_n the LST of the time needed to down-cross L_N , starting from L_n , and before up-crossing L_0 . We construct a system of linear equations in these ρ_n , with the LSTs of Steps 1 and 2 featuring as coefficients in the equations.

Step 1: The time until the first down-crossing of L_{n-k} Let $L_n < x < L_{n-1}$, and let d_n^X and u_{n-1}^X respectively denote the times at which the X process first down-crosses L_n and first up-crosses L_{n-1} , when starting from x. By (5.3) we have

$$\begin{aligned} \xi_n(x - L_n) &:= \mathbb{E}_x \Big[e^{-qd_n^X} \mathbf{1}_{d_n^X < u_{n-1}^X \land \mathcal{E}_\lambda} \Big] \\ &= Z^{(q+\lambda)}(x - L_n) - \frac{W^{(q+\lambda)}(x - L_n)}{W^{(q+\lambda)}(L_{n-1} - L_n)} Z^{(q+\lambda)}(L_{n-1} - L_n). \end{aligned}$$

Denoting by $\tau_{L_{n-k}}^{-}$ the first time that *U* hits a level in \mathcal{N} and that this is done by down-crossing L_{n-k} , we derive

$$\tau_{L_{n-k}}^{-} = \mathcal{E}_{1,\lambda} \mathbf{1}_{\mathcal{E}_{1,\lambda} < u_{n-1}^{X} \land d_{n}^{X}} + \mathcal{E}_{2,\lambda} \mathbf{1}_{\mathcal{E}_{2,\lambda} < u_{n-2}^{X} \land d_{n-1}^{X}} + \dots + \mathcal{E}_{k,\lambda} \mathbf{1}_{\mathcal{E}_{k,\lambda} < u_{n-k}^{X} \land d_{n+1-k}^{X}} + d_{n-k}^{X} \mathbf{1}_{d_{n-k}^{X} < \mathcal{E}_{k+1,\lambda} \land u_{n-k-1}^{X}}$$

where $\mathcal{E}_{k,\lambda}$, k = 1, ..., N are i.i.d. distributed as \mathcal{E}_{λ} . Let

$$r_{n,n-k}(x) := \mathbb{E}_{x} \left[e^{-q\tau_{L_{n-k}}^{-}} \right]$$

= $\mathbb{E}_{x} \left[e^{-q(\sum_{i=1}^{k} \mathcal{E}_{i,\lambda} + d_{n-k}^{X})} 1_{\mathcal{E}_{1,\lambda} < d_{n}^{X} \land u_{n-1}^{X}} 1_{\mathcal{E}_{2,\lambda} < d_{n-1}^{X} \land u_{n-2}^{X}} \right]$
 $\cdots 1_{\mathcal{E}_{k,\lambda} < d_{n-k+1}^{X} \land u_{n-k}^{X}} 1_{d_{n-k}^{X} < \mathcal{E}_{k+1,\lambda} \land u_{n-k-1}^{X}} \right].$

Observe that $r_{n,n-k}(x)$ is the partial LST of the time needed to reach L_{n-k} from above before reaching any other level in \mathcal{N} . Clearly,

$$r_{n,n}(x) = \xi_n(x - L_n).$$
 (5.6)

Applying (5.4) and (5.6) yields

$$r_{n,n-1}(x) = \lambda \int_{L_n}^{L_{n-1}} u_{L_n,L_{n-1}}^{(q+\lambda)}(x, y) r_{n-1,n-1}((a+1)y) dy$$

= $\lambda \frac{W^{(q+\lambda)}(x-L_n)}{W^{(q+\lambda)}(L_{n-1}-L_n)} \int_{L_n}^{L_{n-1}} W^{(q+\lambda)}(L_{n-1}-y) \xi_{n-1}((a+1)y-L_{n-1}) dy$
 $-\lambda \int_{L_n}^x W^{(q+\lambda)}(x-y) \xi_{n-1}((a+1)y-L_{n-1}) dy.$ (5.7)

Note that

$$\int_{L_n}^{x} W^{(q+\lambda)}(x-y)\xi_{n-1}((a+1)y-L_{n-1})dy$$

= $\int_{0}^{x-L_n} W^{(q+\lambda)}(z)\xi_{n-1}((a+1)(x-z-L_n))dz$
= $W^{(q+\lambda)} \circledast \xi_{n-1,a+1}(x-L_n),$

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where \circledast again denotes convolution and $\xi_{n,(a+1)^k}(x) := \xi_n((a+1)^k x), k \in \mathbb{N}$. Thus,

$$r_{n,n-1}(x) = \lambda \frac{W^{(q+\lambda)}(x-L_n)}{W^{(q+\lambda)}(L_{n-1}-L_n)} W^{(q+\lambda)} \circledast \xi_{n-1,a+1}(L_{n-1}-L_n) -\lambda W^{(q+\lambda)} \circledast \xi_{n-1,a+1}(x-L_n).$$
(5.8)

Define

$$A_{0,n,n-1} := \lambda \frac{W^{(q+\lambda)} \circledast \xi_{n-1,a+1}(L_{n-1} - L_n)}{W^{(q+\lambda)}(L_{n-1} - L_n)} \quad \text{and} \quad A_{1,n,n-1} := \lambda.$$

Then

$$r_{n,n-1}(x) = A_{0,n,n-1} W^{(q+\lambda)}(x-L_n) - A_{1,n,n-1} W^{(q+\lambda)} \circledast \xi_{n-1,a+1}(x-L_n).$$

We next obtain $r_{n,n-2}(x)$. Applying (5.4) and (5.8), we have

$$\begin{split} r_{n,n-2}(x) &= \lambda \int_{L_n}^{L_{n-1}} u_{L_n,L_{n-1}}^{(q+\lambda)}(x,y) r_{n-1,n-2}((a+1)y) dy \\ &= \lambda \frac{W^{(q+\lambda)}(x-L_n)}{W^{(q+\lambda)}(L_{n-1}-L_n)} \int_{L_n}^{L_{n-1}} W^{(q+\lambda)}(L_{n-1}-y) \cdot \\ & \left(A_{0,n-1,n-2} W^{(q+\lambda)}((a+1)y-L_{n-1})\right) \\ & -A_{1,n-1,n-2} W^{(q+\lambda)} \circledast \xi_{n-2,a+1}((a+1)y-L_{n-1})\right) dy \\ & -\lambda \int_{L_n}^x W^{(q+\lambda)}(x-y) \cdot \\ & \left(A_{0,n-1,n-2} W^{(q+\lambda)}((a+1)y-L_{n-1})\right) \\ & -A_{1,n-1,n-2} W^{(q+\lambda)}((a+1)y-L_{n-1}) \\ & -A_{1,n-1,n-2} W^{(q+\lambda)} \circledast \xi_{n-2,a+1}((a+1)y-L_{n-1})\right) dy. \end{split}$$

Similarly as before, observe that

$$\int_{L_n}^{x} W^{(q+\lambda)}(x-y) W^{(q+\lambda)}((a+1)y - L_{n-1}) dy = W^{(q+\lambda)} \circledast W^{(q+\lambda)}_{a+1}(x-L_n),$$
(5.9)

where $W_{(a+1)^{k}}^{(q+\lambda)}(x) := W^{(q+\lambda)}((a+1)^{k}x), k \in \mathbb{N}$, and

$$\int_{L_n}^{x} W^{(q+\lambda)}(x-y) W^{(q+\lambda)} \circledast \xi_{n-2,a+1} ((a+1)y - L_{n-1}) dy$$

= $(a+1) W^{(q+\lambda)} \circledast W^{(q+\lambda)}_{a+1} \circledast \xi_{n-2,(a+1)^2} (x - L_n).$

Define

$$\begin{split} A_{0,n,n-2} &:= \frac{\lambda}{W^{(q+\lambda)}(L_{n-1}-L_n)} \left(A_{0,n-1,n-2} W^{(q+\lambda)} \circledast W^{(q+\lambda)}_{a+1}(L_{n-1}-L_n) \right. \\ &- A_{1,n-1,n-2}(a+1) W^{(q+\lambda)} \circledast W^{(q+\lambda)}_{a+1} \circledast \xi_{n-2,(a+1)^2}(L_{n-1}-L_n) \right) \\ &= \frac{1}{W^{(q+\lambda)}(L_{n-1}-L_n)} \cdot \left(A_{1,n,n-2} W^{(q+\lambda)} \circledast W^{(q+\lambda)}_{a+1}(L_{n-1}-L_n) - A_{2,n,n-2} W^{(q+\lambda)} \circledast W^{(q+\lambda)}_{a+1} \circledast \xi_{n-2,(a+1)^2}(L_{n-1}-L_n) \right) , \\ A_{1,n,n-2} &:= \lambda A_{0,n-1,n-2}, \qquad A_{2,n,n-2} := \lambda (a+1) A_{1,n-1,n-2}. \end{split}$$

Then

$$r_{n,n-2}(x) = A_{0,n,n-2} W^{(q+\lambda)}(x - L_n) - A_{1,n,n-2} W^{(q+\lambda)} \circledast W^{(q+\lambda)}_{a+1}(x - L_n) + A_{2,n,n-2} W^{(q+\lambda)} \circledast W^{(q+\lambda)}_{a+1} \circledast \xi_{n-2,(a+1)^2}(x - L_n).$$

The general case for k = 2, ..., n - 1 is given in the following proposition.

Proposition 5.1. *For* $L_n < x < L_{n-1}$ *and* k = 2, ..., n - 1*,*

$$r_{n,n-k}(x) = \sum_{j=0}^{k-1} (-1)^j A_{j,n,n-k} \circledast_{i=0}^j W_{(a+1)^i}^{(q+\lambda)}(x-L_n) + (-1)^k A_{k,n,n-k} \circledast_{i=0}^{k-1} W_{(a+1)^i}^{(q+\lambda)} \circledast \xi_{n-k,(a+1)^k}(x-L_n),$$
(5.10)

where $A_{j,n,n-k}$, j = 0, ..., k, are coefficients which are obtained recursively.

Proof. The proof is by induction on k. Clearly, (5.10) holds for k = 2. Assume it holds for $k - 1 \ge 2$. By the induction hypothesis we have

$$r_{n,n-(k-1)}(x) = \sum_{j=0}^{k-2} (-1)^j A_{j,n,n-(k-1)} \circledast_{i=0}^j W_{(a+1)^i}^{(q+\lambda)}(x-L_n)$$

$$+ (-1)^{k-1} A_{k-1,n,n-(k-1)} \circledast_{i=0}^{k-2} W_{(a+1)^i}^{(q+\lambda)} \circledast \xi_{n-(k-1),(a+1)^{k-1}}(x-L_n).$$
(5.11)

Using (5.4) and (5.11), we have

$$\begin{split} r_{n,n-k}(x) &= \lambda \int_{L_n}^{L_{n-1}} u_{L_n,L_{n-1}}^{(q+\lambda)}(x,y) r_{n-1,n-k}((a+1)y) dy \\ &= \lambda \frac{W^{(q+\lambda)}(x-L_n)}{W^{(q+\lambda)}(L_{n-1}-L_n)} \int_{L_n}^{L_{n-1}} W^{(q+\lambda)}(L_{n-1}-y) \cdot \\ &\left(\sum_{j=0}^{k-2} (-1)^j A_{j,n-1,n-k} \circledast_{i=0}^j W^{(q+\lambda)}_{(a+1)^i}((a+1)y-L_{n-1}) \right) \\ &+ (-1)^{k-1} A_{k-1,n-1,n-k} \circledast_{i=0}^{k-2} W^{(q+\lambda)}_{(a+1)^i} \circledast \xi_{n-k,(a+1)^{k-1}}((a+1)y-L_{n-1}) \right) dy \end{split}$$

568

A dual risk model with additive and proportional gains

$$\begin{split} &-\lambda \int_{L_n}^{x} W^{(q+\lambda)}(x-y) \cdot \\ &\left(\sum_{j=0}^{k-2} (-1)^j A_{j,n-1,n-k} \circledast_{i=0}^j W^{(q+\lambda)}_{(a+1)^i} ((a+1)y - L_{n-1}) \right. \\ &+ (-1)^{k-1} A_{k-1,n-1,n-k} \circledast_{i=0}^{k-2} W^{(q+\lambda)}_{(a+1)^i} \circledast \xi_{n-k,(a+1)^{k-1}} ((a+1)y - L_{n-1}) \right) dy. \end{split}$$

Note that (5.9) holds, and for $j \ge 1$,

$$\int_{L_n}^{x} W^{(q+\lambda)}(x-y) \circledast_{i=0}^{j} W^{(q+\lambda)}_{(a+1)^i}((a+1)y - L_{n-1}) dy$$

= $(a+1) \circledast_{i=0}^{j+1} W^{(q+\lambda)}_{(a+1)^i}(x - L_n).$

If we choose

$$\begin{aligned} A_{1,n,n-k} &:= \lambda A_{0,n-1,n-k}, \\ A_{j+1,n,n-k} &:= \lambda(a+1)A_{j,n-1,n-k}, \\ A_{0,n,n-k} &:= \frac{1}{W^{(q+\lambda)}(L_{n-1}-L_n)} \left(\sum_{j=1}^{k-1} (-1)^{j-1}A_{j,n,n-k} \circledast_{i=0}^{j} W_{(a+1)^{i}}^{(q+\lambda)}(L_{n-1}-L_n) \right. \\ &+ (-1)^{k-1}A_{k,n,n-k} \circledast_{i=0}^{k-1} W_{(a+1)^{i}}^{(q+\lambda)} \circledast \xi_{n-k,(a+1)^{k}}(L_{n-1}-L_n) \right), \end{aligned}$$

then (5.10) holds true, which completes the proof.

Step 2: The time until the first up-crossing of L_{n-k} Let $L_n < x < L_{n-1}$, and let $\tau_{L_{n-k}}^+$ be the first time that *U* reaches a level in \mathcal{N} and that this is done by up-crossing L_{n-k} by the Brownian motion. Note that

$$\tau_{L_{n-k}}^{+} = \mathcal{E}_{1,\lambda} \mathbf{1}_{\mathcal{E}_{1,\lambda} < u_{n-1}^{X} \land d_{n}^{X}} + \mathcal{E}_{2,\lambda} \mathbf{1}_{\mathcal{E}_{2,\lambda} < u_{n-2}^{X} \land d_{n-1}^{X}} + \dots + \mathcal{E}_{k-1,\lambda} \mathbf{1}_{\mathcal{E}_{k-1,\lambda} < u_{n-k+1}^{X} \land d_{n+2-k}^{X}} + u_{n-k}^{X} \mathbf{1}_{u_{n-k}^{X} < \mathcal{E}_{k,\lambda} \land d_{n-k+1}^{X}}.$$

For $k = 1, \ldots, n$ we define

$$\omega_{n,n-k}(x) := \mathbb{E}_{x} \left[e^{-q\tau_{L_{n-k}}^{+}} \right]$$
$$= \mathbb{E}_{x} \left[e^{-q \left(\sum_{j=1}^{k-1} \mathcal{E}_{j,\lambda} + u_{n-k}^{\chi} \right)} \mathbf{1}_{\mathcal{E}_{1,\lambda} < u_{n-1}^{\chi} \land d_{n}^{\chi}} \cdots \mathbf{1}_{\mathcal{E}_{k-1,\lambda} < u_{n-k+1}^{\chi} \land d_{n+2-k}^{\chi}} \mathbf{1}_{u_{n-k}^{\chi} < d_{n-k+1}^{\chi} \land \mathcal{E}_{k,\lambda}} \right].$$

Applying (5.2) we have

$$\Omega_{n-1}(x-L_n) := \omega_{n,n-1}(x) = \mathbb{E}_x \left(e^{-q\tau_{L_{n-1}}^+} \mathbf{1}_{u_{n-1}^X < \mathcal{E}_{1,\lambda} \land d_n^X} \right) = \frac{W^{(q+\lambda)}(x-L_n)}{W^{(q+\lambda)}(L_{n-1}-L_n)}.$$
 (5.12)

Further, using (5.4) and (5.12), observe that

$$\begin{split} \omega_{n,n-2}(x) &= \lambda \int_{L_n}^{L_{n-1}} u_{L_n,L_{n-1}}^{(q+\lambda)}(x,y) \omega_{n-1,n-2}((a+1)y) dy \\ &= \lambda \frac{W^{(q+\lambda)}(x-L_n)}{W^{(q+\lambda)}(L_{n-1}-L_n)} \int_{L_n}^{L_{n-1}} W^{(q+\lambda)}(L_{n-1}-y) \Omega_{n-2}((a+1)y-L_{n-1}) dy \\ &\quad -\lambda \int_{L_n}^x W^{(q+\lambda)}(x-y) \Omega_{n-2}((a+1)y-L_{n-1}) dy \\ &= \lambda \frac{W^{(q+\lambda)}(x-L_n)}{W^{(q+\lambda)}(L_{n-1}-L_n)} W^{(q+\lambda)} \circledast \Omega_{n-2,a+1}(L_{n-1}-L_n) \\ &\quad -\lambda W^{(q+\lambda)} \circledast \Omega_{n-2,a+1}(x-L_n), \end{split}$$

where $\Omega_{n,(a+1)^k}(x) = \Omega_n((a+1)^k x), k \in \mathbb{N}$. Let

$$B_{1,n,n-2} := \lambda$$
 and $B_{0,n,n-2} := \lambda \frac{W^{(q+\lambda)} \circledast \Omega_{n-2,a+1}(L_{n-1} - L_n)}{W^{(q+\lambda)}(L_{n-1} - L_n)}$

Then

$$\omega_{n,n-2}(x) = B_{0,n,n-2}W^{(q+\lambda)}(x-L_n) - B_{1,n,n-2}W^{(q+\lambda)} \circledast \Omega_{n-2,a+1}(x-L_n).$$

The next proposition gives a general expression for $\omega_{n,n-k}(x)$.

Proposition 5.2. For $k = 2, \ldots, n$ we have

$$\omega_{n,n-k}(x) = \sum_{j=0}^{k-2} (-1)^j B_{j,n,n-k} \circledast_{i=0}^j W_{(a+1)^i}^{(q+\lambda)}(x-L_n) + (-1)^{k-1} B_{k-1,n,n-k} \circledast_{i=0}^{k-2} W_{(a+1)^i}^{(q+\lambda)} \circledast \Omega_{n-k,(a+1)^{k-1}}(x-L_n),$$

where $B_{j,n,n-k}$, j = 0, ..., k - 1, are coefficients which are obtained recursively.

Proof. The proof is similar to the proof of Proposition 5.1. The proposition clearly holds for k = 2. Assume it holds for $k - 1 \ge 2$. Applying (5.4) we obtain that

$$\begin{split} \omega_{n,n-k}(x) &= \lambda \int_{L_n}^{L_{n-1}} u_{L_n,L_{n-1}}^{(q+\lambda)}(x,y) \omega_{n-1,n-k}((a+1)y) dy \\ &= \lambda \frac{W^{(q+\lambda)}(x-L_n)}{W^{(q+\lambda)}(L_{n-1}-L_n)} \int_{L_n}^{L_{n-1}} W^{(q+\lambda)}(L_{n-1}-y) \cdot \\ &\left(\sum_{j=0}^{k-3} (-1)^j B_{j,n-1,n-k} \circledast_{i=0}^j W_{(a+1)^i}^{(q+\lambda)}((a+1)y-L_{n-1}) \right) \\ &+ (-1)^{k-2} B_{k-2,n-1,n-k} \circledast_{i=0}^{k-3} W_{(a+1)^i}^{(q+\lambda)} \circledast \Omega_{n-k,(a+1)^{k-2}}((a+1)y-L_{n-1}) \right) dy \end{split}$$

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A dual risk model with additive and proportional gains

$$- \lambda \int_{L_n}^{x} W^{(q+\lambda)}(x-y) \cdot \left(\sum_{j=0}^{k-3} (-1)^j B_{j,n-1,n-k} \circledast_{i=0}^j W^{(q+\lambda)}_{(a+1)^i} ((a+1)y - L_{n-1}) + (-1)^{k-2} B_{k-2,n-1,n-k} \circledast_{i=0}^{k-3} W^{(q+\lambda)}_{(a+1)^i} \circledast \Omega_{n-k,(a+1)^{k-2}} ((a+1)y - L_{n-1}) \right) dy.$$

Taking

$$B_{1,n,n-k} := \lambda B_{0,n-1,n-k}, \quad B_{j+1,n,n-k} := (a+1)\lambda B_{j,n-1,n-k}, \quad j = 1, \dots, k-2,$$

$$B_{0,n,n-k} := \frac{1}{W^{(q+\lambda)}(L_{n-1} - L_n)} \cdot \left(\sum_{j=1}^{k-1} (-1)^j B_{j,n,n-k} \circledast_{i=0}^j W^{(q+\lambda)}_{(a+1)^i}(L_{n-1} - L_n) + (-1)^{k-1} B_{k-1,n,n-k} \circledast_{i=0}^{k-2} W^{(q+\lambda)}_{(a+1)^i} \circledast \Omega_{n-k,(a+1)^{k-1}}(L_{n-1} - L_n)\right)$$

completes the proof of this proposition.

Step 3: Determination of the exit/ruin time transform $\rho_N(x)$

To find $\rho_N(x)$ we start from the key observation that for $L_n < x < L_{n-1}$ we have

$$\rho_N(x) = r_{n,n}(x)\rho_n + \sum_{j=1}^{n-1} (r_{n,n-j}(x) + \omega_{n,n-j}(x))\rho_{n-j},$$
(5.13)

where

$$\rho_n := E_{L_n} \left[e^{-qd_N} \mathbf{1}_{d_N < u_0} \right].$$

In the next step we construct a system of linear equations to find ρ_n , n = 1, 2, ..., N. Clearly, $\rho_0 = 0$ and $\rho_N = 1$. Moreover,

$$\rho_{1} = \left(Z^{(q+\lambda)}(L_{1} - L_{2}) - \frac{W^{(q+\lambda)}(L_{1} - L_{2})}{W^{(q+\lambda)}(L_{0} - L_{2})} Z^{(q+\lambda)}(L_{0} - L_{2}) \right) \rho_{2} + \left(\lambda \int_{L_{2}}^{L_{1}} u^{(q+\lambda)}_{L_{2},L_{0}}(L_{1}, y) r_{1,1}((a+1)y) dy \right) \rho_{1}.$$

The term in the first set of parentheses is the Laplace transform of the time to down-cross L_2 before \mathcal{E}_{λ} and before L_0 is reached; cf. (5.3). The second term is the Laplace transform of \mathcal{E}_{λ} , where the exponential time expires when $U \in (y, y + dy)$ is between L_2 and L_1 before reaching L_2 or L_0 , and then the time to reach L_1 from above. Similarly, note that

$$\rho_2 = \left(Z^{(q+\lambda)}(L_2 - L_3) - \frac{W^{(q+\lambda)}(L_2 - L_3)}{W^{(q+\lambda)}(L_1 - L_3)} Z^{(q+\lambda)}(L_1 - L_3) \right) \rho_3$$
(5.14)

$$+\frac{W^{(q+\lambda)}(L_2-L_3)}{W^{(q+\lambda)}(L_1-L_3)}\rho_1$$
(5.15)

O. BOXMA ET AL.

+
$$\lambda \int_{L_3}^{L_2} u_{L_3,L_1}^{(q+\lambda)}(L_2, y) \left(r_{2,2}((a+1)y)\rho_2 + (r_{2,1}((a+1)y) + \omega_{2,1}((a+1)y))\rho_1 \right) dy$$
 (5.16)

+
$$\left(\lambda \int_{L_2}^{L_1} u_{L_3,L_1}^{(q+\lambda)}(L_2, y) r_{1,1}((a+1)y) dy\right) \rho_1.$$
 (5.17)

The term in the parentheses in (5.14) is the expected discounted time to reach L_3 before a jump and before up-crossing L_1 . The factor in (5.15) is the expected discounted time to reach L_1 before a jump and before down-crossing L_3 (cf. (5.2)). The expressions (5.16) and (5.17) describe the expected discounted time until a jump when a jump occurs before reaching L_1 or L_3 , and then the expected discounted time until the process reaches one of the levels L_j for $j \le 2$. The expression (5.16) describes the case where just before a jump U is between L_3 and L_2 , and (5.17) describes the case where just before a jump U is between L_2 and L_1 . By rearranging (5.14)–(5.17), we get

$$\begin{split} \rho_2 &= \left(Z^{(q+\lambda)}(L_2 - L_3) - \frac{W^{(q+\lambda)}(L_2 - L_3)}{W^{(q+\lambda)}(L_1 - L_3)} Z^{(q+\lambda)}(L_1 - L_3) \right) \rho_3 \\ &+ \left(\lambda \int_{L_3}^{L_2} u_{L_3,L_1}^{(q+\lambda)}(L_2, y) r_{2,2}((a+1)y) dy \right) \rho_2 \\ &+ \left(\frac{W^{(q+\lambda)}(L_2 - L_3)}{W^{(q+\lambda)}(L_1 - L_3)} + \lambda \int_{L_3}^{L_2} u_{L_3,L_1}^{(q+\lambda)}(L_2, y) (r_{2,1}((a+1)y) + \omega_{2,1}((a+1)y)) dy \right) \\ &+ \lambda \int_{L_2}^{L_1} u_{L_3,L_1}^{(q+\lambda)}(L_2, y) r_{1,1}((a+1)y) dy \right) \rho_1. \end{split}$$

Using similar arguments, we can show that generally, for $1 < n \le N - 1$,

$$\begin{split} \rho_n &= \left(Z^{(q+\lambda)}(L_n - L_{n+1}) - \frac{W^{(q+\lambda)}(L_n - L_{n+1})}{W^{(q+\lambda)}(L_{n-1} - L_{n+1})} Z^{(q+\lambda)}(L_{n-1} - L_{n+1}) \right) \rho_{n+1} \\ &+ \frac{W^{(q+\lambda)}(L_n - L_{n+1})}{W^{(q+\lambda)}(L_{n-1} - L_{n+1})} \rho_{n-1} \\ &+ \lambda \int_{L_{n+1}}^{L_n} u^{(q+\lambda)}_{L_{n+1},L_{n-1}}(L_n, y) \left(r_{n,n}((a+1)y)\rho_n \\ &+ \sum_{k=1}^{n-1} \left(r_{n,n-k}((a+1)y) + \omega_{n,n-k}((a+1)y) \right) \rho_{n-k} \right) dy \\ &+ \lambda \int_{L_n}^{L_{n-1}} u^{(q+\lambda)}_{L_{n+1},L_{n-1}}(L_n, y) \left(r_{n-1,n-1}((a+1)y)\rho_{n-1} \\ &+ \sum_{k=1}^{n-2} \left(r_{n-1,n-1-k}((a+1)y) + \omega_{n-1,n-1-k}((a+1)y)\rho_{n-1-k} \right) dy, \end{split}$$

which is equivalent to

$$\rho_n = \left(Z^{(q+\lambda)}(L_n - L_{n+1}) - \frac{W^{(q+\lambda)}(L_n - L_{n+1})}{W^{(q+\lambda)}(L_{n-1} - L_{n+1})} Z^{(q+\lambda)}(L_{n-1} - L_{n+1}) \right) \rho_{n+1}$$
$$+ \left(\lambda \int_{L_{n+1}}^{L_n} u_{L_{n+1},L_{n-1}}^{(q+\lambda)}(L_n, y) r_{n,n}((a+1)y) dy \right) \rho_n$$

A dual risk model with additive and proportional gains

$$+ \left(\lambda \int_{L_{n+1}}^{L_n} u_{L_{n+1},L_{n-1}}^{(q+\lambda)} (L_n, y)(r_{n,n-1}((a+1)y) + \omega_{n,n-1}((a+1)y))dy + \frac{W^{(q+\lambda)}(L_n - L_{n+1})}{W^{(q+\lambda)}(L_{n-1} - L_{n+1})} + \lambda \int_{L_n}^{L_{n-1}} u_{L_{n+1},L_{n-1}}^{(q+\lambda)} (L_n, y)r_{n-1,n-1}((a+1)y)dy \right) \rho_{n-1} + \lambda \sum_{k=2}^{n-1} \left(\int_{L_{n+1}}^{L_n} u_{L_{n+1},L_{n-1}}^{(q+\lambda)} (L_n, y)(r_{n,n-k}((a+1)y) + \omega_{n,n-k}((a+1)y))dy + \int_{L_n}^{L_{n-1}} u_{L_{n+1},L_{n-1}}^{(q+\lambda)} (L_n, y)(r_{n-1,n-1-(k-1)}((a+1)y) + \omega_{n,n-k}((a+1)y))dy + \omega_{n-1,n-1-(k-1)}((a+1)y))dy \right) \rho_{n-k}.$$

$$(5.18)$$

Thus we have proved the following main result.

Theorem 5.1. The two-sided downward exit time transform $\rho_N(x)$ defined in (5.5) is given in (5.13) with $r_{n,n-k}$ identified in (5.6), (5.7), and Proposition 5.1; $\omega_{n,n-k}$ identified in (5.12) and Proposition 5.2; and ρ_k given via the system of equations (5.18).

5.2. Expected discounted dividends until ruin

In this section we obtain $v_N(x)$ —the expected discounted dividends obtained up to the time when the process reaches L_N , starting at x. Let $L_n < x < L_{n-1}$ and

$$\mathcal{T}_{n,0}(x) := \mathbb{E}_{x} \Big[e^{-q\mathcal{S}_{n}} \mathbf{1}_{\mathcal{E}_{1,\lambda} < u_{n-1}^{\chi} \wedge d_{n}^{\chi}} \mathbf{1}_{\mathcal{E}_{2,\lambda} < u_{n-2}^{\chi} \wedge d_{n-1}^{\chi}} \cdots \mathbf{1}_{\mathcal{E}_{n,\lambda} < u_{0}^{\chi} \wedge d_{1}^{\chi}} \Big],$$

where

$$S_n := \sum_{i=1}^n \mathcal{E}_{i,\lambda}$$

Thus $\mathcal{T}_{n,0}(x)$ is the expected discounted time until up-crossing L_0 by a jump when this occurs before reaching any level in \mathcal{N} . Also for $L_n < x < L_{n-1}$ let

$$v_n^J(x) := \mathbb{E}_x \Big[e^{-qS_n} \mathbf{1}_{\mathcal{E}_{1,\lambda} < u_{n-1}^X \land d_n^X} \mathbf{1}_{\mathcal{E}_{2,\lambda} < u_{n-2}^X \land d_{n-1}^X} \dots \mathbf{1}_{\mathcal{E}_{n,\lambda} < u_0^X \land d_1^X} ((a+1)U(S_n) - L_0) \Big].$$
(5.19)

Note that $v_n^J(x)$ is the expected discounted overflow above $L_0 = b$ when this occurs before reaching any level in \mathcal{N} . First consider $v_1^J(x)$, so take $L_1 < x < L_0$. Applying (5.4) we get

$$v_1^J(x) = \lambda \int_{L_1}^{L_0} u_{L_1,L_0}^{(q+\lambda)}(x, y)((a+1)y - L_0) dy$$

$$= \lambda \left(\frac{W^{(q+\lambda)}(x - L_1)}{W^{(q+\lambda)}(L_0 - L_1)} \int_{L_1}^{L_0} W^{(q+\lambda)}(L_0 - y)((a+1)y - L_0) dy - \int_{L_1}^x W^{(q+\lambda)}(x - y)((a+1)y - L_0) dy \right).$$
(5.20)

Thus,

$$v_1^J(x) = A_{1,f,0} W^{(q+\lambda)}(x-L_1) - A_{1,f,1} W^{(q+\lambda)} \circledast \mathcal{Q}_{a+1}(x-L_1),$$

where $Q_{(a+1)^k}(x) = Q((a+1)^k x) = (a+1)^k x, k \in \mathbb{N}$, and

$$A_{1,f,0} := \lambda \frac{W^{(q+\lambda)} \circledast \mathcal{Q}_{a+1}(L_0 - L_1)}{W^{(q+\lambda)}(L_0 - L_1)} \text{ and } A_{1,f,1} := \lambda.$$

Similarly, replacing $(a + 1)y - L_0$ by 1 in (5.20), we obtain that the expected discounted time until a jump above $L_0 = b$ is

$$\mathcal{T}_{1,0}(x) = A_{1,\mathcal{T},0} W^{(q+\lambda)}(x-L_1) - A_{1,\mathcal{T},1} \overline{W}^{(q+\lambda)}(x-L_1),$$

where

$$A_{1,\mathcal{T},0} = \lambda \frac{\overline{W}^{(q+\lambda)}(L_0 - L_1)}{W^{(q+\lambda)}(L_0 - L_1)} \quad \text{and} \quad A_{1,\mathcal{T},1} = \lambda,$$
(5.21)

and $\overline{W}^{(q)}(x) = \int_0^x W^{(q)}(y) dy$. Next consider $v_2^J(x)$, so take $L_2 < x < L_1$. Then

$$\begin{split} v_2^J(x) &= \lambda \int_{L_2}^{L_1} u_{L_2,L_1}^{(q+\lambda)}(x, y) v_1^J((a+1)y) \mathrm{d}y \\ &= \lambda \frac{W^{(q+\lambda)}(x-L_2)}{W^{(q+\lambda)}(L_1-L_2)} \int_{L_2}^{L_1} W^{(q+\lambda)}(L_1-y) \left(A_{1,f,0} W^{(q+\lambda)}((a+1)y-L_1)\right) \\ &\quad -A_{1,f,1} W^{(q+\lambda)} \circledast \mathcal{Q}_{a+1}((a+1)y-L_1) \right) \mathrm{d}y \\ &\quad -\lambda \int_{L_2}^x W^{(q+\lambda)}(x-y) \left(A_{1,f,0} W^{(q+\lambda)}((a+1)y-L_1)\right) \\ &\quad -A_{1,f,1} W^{(q+\lambda)} \circledast \mathcal{Q}_{a+1}((a+1)y-L_1) \right) \mathrm{d}y. \end{split}$$

Thus,

$$v_{2}^{J}(x) = A_{2,f,0}W^{(q+\lambda)}(x-L_{2}) - A_{2,f,1}W^{(q+\lambda)} \circledast W_{a+1}^{(q+\lambda)}(x-L_{2}) + A_{2,f,2}W^{(q+\lambda)} \circledast W_{a+1}^{(q+\lambda)} \circledast \mathcal{Q}_{(a+1)^{2}}(x-L_{2}),$$

where

$$\begin{split} A_{2,f,1} &:= \lambda A_{1,f,0}, \qquad A_{2,f,2} := \lambda(a+1)A_{1,f,1}, \\ A_{2,f,0} &:= \frac{\lambda A_{1,f,0} W^{(q+\lambda)} \circledast W^{(q+\lambda)}_{a+1}(L_1 - L_2)}{W^{(q+\lambda)}(L_1 - L_2)} \\ &- \frac{\lambda A_{1,f,1}(a+1) W^{(q+\lambda)} \circledast W^{(q+\lambda)}_{a+1} \circledast \mathcal{Q}_{(a+1)^2}(L_1 - L_2)}{W^{(q+\lambda)}(L_1 - L_2)} \\ &= \frac{A_{2,f,1} W^{(q+\lambda)} \circledast W^{(q+\lambda)}_{a+1}(L_1 - L_2) - A_{2,f,2} W^{(q+\lambda)} \circledast W^{(q+\lambda)}_{a+1} \circledast \mathcal{Q}_{(a+1)^2}(L_1 - L_2)}{W^{(q+\lambda)}(L_1 - L_2)}. \end{split}$$

Similarly,

$$\mathcal{T}_{2,0}(x) = A_{2,\mathcal{T},0} W^{(q+\lambda)}(x-L_2) - A_{2,\mathcal{T},1} W^{(q+\lambda)} \circledast W^{(q+\lambda)}_{a+1}(x-L_2) + A_{2,\mathcal{T},2} W^{(q+\lambda)} \circledast \overline{W}_{a+1}(x-L_2),$$

574

where

$$A_{2,\mathcal{T},1} := \lambda A_{1,\mathcal{T},0} \text{ and } A_{2,\mathcal{T},2} := \lambda A_{1,\mathcal{T},1},$$
 (5.22)

$$A_{2,\mathcal{T},0} := \frac{A_{2,\mathcal{T},1} W^{(q+\lambda)} \circledast W^{(q+\lambda)}_{a+1} (L_1 - L_2) - A_{2,\mathcal{T},2} W^{(q+\lambda)} \circledast \overline{W}^{(q+\lambda)}_{a+1} (L_1 - L_2)}{W^{(q+\lambda)} (L_1 - L_2)}.$$
 (5.23)

Using arguments similar to those in the proof of Proposition 5.2, one can derive the following result.

Proposition 5.3. For $L_n < x < L_{n-1}$ we have

$$v_n^J(x) = \sum_{j=0}^{n-1} (-1)^j A_{n,f,j} \circledast_{i=0}^j W_{(a+1)^i}^{(q+\lambda)}(x-L_n) + (-1)^n A_{n,f,n} \circledast_{i=0}^{n-1} W_{(a+1)^i}^{(q+\lambda)} \circledast \mathcal{Q}_{(a+1)^n}(x-L_n),$$

where the $A_{n,f,j}$ are obtained recursively as follows:

$$\begin{split} A_{n,f,1} &:= \lambda A_{n-1,f,0}, \\ A_{n,f,j} &:= \lambda (a+1) A_{n-1,f,j-1}, \quad 2 \le j \le n, \\ A_{n,f,0} &:= \frac{\sum_{j=1}^{n-1} (-1)^{j-1} A_{n,f,j} \circledast_{i=0}^{j} W_{(a+1)^{i}}^{(q+\lambda)} (L_{n-1} - L_{n})}{W^{(q+\lambda)} (L_{n-1} - L_{n})} \\ &+ (-1)^{n-1} \frac{A_{n,f,n} \circledast_{i=0}^{n-1} W_{(a+1)^{i}}^{(q+\lambda)} \circledast \mathcal{Q}_{(a+1)^{n}} (L_{n-1} - L_{n})}{W^{(q+\lambda)} (L_{n-1} - L_{n})} \end{split}$$

Similarly, for $L_n < x < L_{n-1}$,

$$\mathcal{T}_{n,0}(x) = \sum_{j=0}^{n-1} (-1)^j A_{n,\mathcal{T},j} \circledast_{i=0}^{j-1} W_{(a+1)^i}^{(q+\lambda)}(x-L_n) + (-1)^n A_{n,\mathcal{T},n} \circledast_{i=0}^{n-1} W_{(a+1)^i}^{(q+\lambda)} \circledast \overline{W}_{(a+1)^n}^{(q+\lambda)}(x-L_n),$$

where $A_{n,T,j}$, n = 1, 2, j = 0, 1, 2, are as in (5.21) and (5.22)–(5.23). For $n \ge 3$, the $A_{n,T,j}$ are obtained recursively as follows:

$$\begin{split} A_{n,\mathcal{T},1} &\coloneqq \lambda A_{n-1,\mathcal{T},0}, \\ A_{n,\mathcal{T},j} &\coloneqq \lambda(a+1)A_{n-1,\mathcal{T},j-1}, \quad 2 \le j \le n, \\ A_{n,\mathcal{T},0} &\coloneqq \frac{\sum_{j=1}^{n-1} (-1)^{j-1}A_{n,\mathcal{T},j} \circledast_{i=0}^{j} W_{(a+1)^{i}}^{(q+\lambda)}(L_{n-1} - L_{n})}{W^{(q+\lambda)}(L_{n-1} - L_{n})} \\ &+ (-1)^{n-1} \frac{A_{n,\mathcal{T},n} \circledast_{i=0}^{n-1} W_{(a+1)^{i}}^{(q+\lambda)} \circledast \overline{W}_{(a+1)^{n}}^{(q+\lambda)}(L_{n-1} - L_{n})}{W^{(q+\lambda)}(L_{n-1} - L_{n})}. \end{split}$$

Recall that $v_N(x)$ gives the expected discounted dividends upon reaching L_N , having started at state x, and (cf. (4.7)) $v_n = v_N(L_n)$. Observe that for $L_n < x < L_{n-1}$ we have

$$v_N(x) = \sum_{k=0}^{n-1} r_{n,n-k}(x)v_{n-k} + \sum_{k=1}^{n-1} \omega_{n,n-k}(x)v_{n-k} + (\omega_{n,0}(x) + \mathcal{T}_{n,0})v_0 + v_n^J(x).$$
(5.24)

The first two terms in the right-hand side of (5.24) correspond to cases in which a level from \mathcal{N} is reached before $L_0 = b$ is reached or up-crossed. The v_0 term covers the two cases in which level L_0 is reached ($\omega_{n,0}(x)$) is the expected discounted time to reach level L_0 before reaching any other level in \mathcal{N}) and level L_0 is up-crossed by a jump ($\mathcal{T}_{n,0}(x)$) is the expected discounted time until up-crossing L_0 before reaching any other level in \mathcal{N}). Finally $v_n^J(x)$ is the expected discounted overflow above L_0 by a jump, when it occurs before reaching any level in \mathcal{N} .

We now derive a system of equations identifying all v_n . Clearly $v_N = 0$. Let us set an equation for v_0 . Assume that $U(0) = b = L_0$. Let $\overline{X}(t) := \sup_{0 \le s \le t} \{X(s)\}, V(t) := (\overline{X}(t) - b)_+$, and let R(t) = X(t) - V(t), where $y_+ = \max(y, 0)$. Observe that V(t) is the cumulative amount of dividends obtained up to time t only via the process X(t). From Theorem 8.11 in Kyprianou [17] we have, with $d_{\alpha}^R = \min\{t : R(t) = \alpha\}$,

$$\frac{\mathbb{P}_{x}(R(\mathcal{E}_{q}) \in (y, y + \mathrm{d}y), \mathcal{E}_{q} < d_{\alpha}^{R})}{q\mathrm{d}y} = \mu^{(q)}(x, y) = \frac{W^{(q)}(x - \alpha)}{W^{(q)'}(b - \alpha)}W^{(q)'}(b - y) - W^{(q)}(x - y),$$

where $W^{(q)'}(x)$ is the derivative of $W^{(q)}(x)$ with respect to x. Moreover, from [7] we know that the expected discounted dividends paid until $d_{L_1}^R \wedge \mathcal{E}_{\lambda}$ starting at b equal

$$\eta\left(b,\frac{b}{a+1}\right) := \mathbb{E}_b\left[\int_0^\infty e^{-qt} \mathbf{1}_{t < d_{L_1}^R \wedge \mathcal{E}_\lambda} \mathrm{d}V(t)\right] = \frac{W^{(q+\lambda)}\left(b - \frac{b}{a+1}\right)}{W^{(q+\lambda)'}\left(b - \frac{b}{a+1}\right)}.$$

Additionally, from Theorem 8.10(i) in Kyprianou [17] with $\theta = 0$ we have

$$\mathbb{E}_{x}\left[e^{-qd_{\alpha}^{R}}\mathbf{1}_{d_{\alpha}^{R}<\mathcal{E}_{\lambda}}\right] = Z^{(q+\lambda)}(x-\alpha) - (q+\lambda)\frac{W^{(q+\lambda)}(b-\alpha)}{W^{(q+\lambda)'}(b-\alpha)}W^{(q+\lambda)}(x-\alpha).$$
(5.25)

Therefore,

$$v_{0} = \eta \left(b, \frac{b}{a+1} \right) + \lambda \int_{L_{1}}^{L_{0}} \mu^{(q+\lambda)}(b, y)((a+1)y - b + v_{0}) dy + \left(Z^{(q+\lambda)}(L_{0} - L_{1}) - (q+\lambda) \frac{W^{(q+\lambda)}(L_{0} - L_{1})}{W^{(q+\lambda)'}(L_{0} - L_{1})} W^{(q+\lambda)}(L_{0} - L_{1}) \right) v_{1}.$$
(5.26)

The second term is the expected discounted dividends due to a jump that occurs at time \mathcal{E}_{λ} before down-crossing L_1 . The last term equals

$$\mathbb{E}_{L_0}\left[e^{-qd_{L_1}^R}\mathbf{1}_{d_{L_1}^R} < \mathcal{E}_{\lambda}\right] v_1$$

and hence is the expected discounted dividends when L_1 is down-crossed before the exponential time \mathcal{E}_{λ} has expired. Further, we have

$$v_1 = \left(Z^{(q+\lambda)}(L_1 - L_2) - \frac{W^{(q+\lambda)}(L_1 - L_2)}{W^{(q+\lambda)}(L_0 - L_2)} Z^{(q+\lambda)}(L_0 - L_2) \right) v_2$$
(5.27)

+
$$\left(\lambda \int_{L_2}^{L_1} u_{L_2,L_0}^{(q+\lambda)}(L_1, y) r_{1,1}((a+1)y) \mathrm{d}y\right) v_1$$
 (5.28)

$$+\left(\frac{W^{(q+\lambda)}(L_1-L_2)}{W^{(q+\lambda)}(L_0-L_2)}\right)$$
(5.29)

A dual risk model with additive and proportional gains

$$+\lambda \int_{L_2}^{L_1} u_{L_2,L_0}^{(q+\lambda)}(L_1, y) \left(\mathcal{T}_{1,0}((a+1)y) + \omega_{1,0}((a+1)y) \right) dy \right) v_0$$
(5.30)

$$+\lambda \int_{L_1}^{L_0} u_{L_2,L_0}^{(q+\lambda)}(L_1, y)((a+1)y - b) \mathrm{d}y$$
(5.31)

$$+\lambda \int_{L_2}^{L_1} u_{L_2,L_0}^{(q+\lambda)}(L_1, y) v_1^J((a+1)y) \mathrm{d}y.$$
(5.32)

The term in the parentheses in (5.27) is the expected discounted time to reach L_2 before a jump and before reaching L_0 . The term that multiplies v_1 in (5.28) is the expected discounted time to reach L_1 before any other level in \mathcal{N} is reached. The term (5.29) is the expected discounted time to reach L_0 by the Brownian motion before down-crossing L_2 and before the exponential time has expired. The first term in (5.30) is the expected discounted time to jump above b when this jump occurs before the exponential time has expired, and the second term is the expected discounted time to reach b by the Brownian motion. The term (5.31) is the expected discounted dividends due to a jump when the exponential time has expired while the process is in (L_1, L_0) , and (5.32) is the expected discounted dividends due to a jump when the exponential time has expired while the process is in (L_2, L_1) . Similarly, we can observe that

$$v_2 = \left(Z^{(q+\lambda)}(L_2 - L_3) - \frac{W^{(q+\lambda)}(L_2 - L_3)}{W^{(q+\lambda)}(L_1 - L_3)}Z^{(q+\lambda)}(L_1 - L_3)\right)v_3$$
(5.33)

$$+ \left(\lambda \int_{L_3}^{L_2} u_{L_3,L_1}^{(q+\lambda)}(L_2, y) r_{2,2}((a+1)y) \mathrm{d}y\right) v_2$$
(5.34)

+
$$\left(\frac{W^{(q+\lambda)}(L_2-L_3)}{W^{(q+\lambda)}(L_1-L_3)} + \lambda \int_{L_3}^{L_2} u^{(q+\lambda)}_{L_3,L_1}(L_2, y)(r_{2,1}((a+1)y) + \omega_{2,1}((a+1)y))dy\right)$$
 (5.35)

$$+ \lambda \int_{L_2}^{L_1} u_{L_3,L_1}^{(q+\lambda)}(L_2, y) r_{1,1}((a+1)y) dy \bigg) v_1$$
(5.36)

+
$$\left(\lambda \int_{L_3}^{L_2} u_{L_3,L_1}^{(q+\lambda)}(L_2, y)(\mathcal{T}_{2,0}((a+1)y) + \omega_{2,0}((a+1)y))dy\right)$$
 (5.37)

$$+ \lambda \int_{L_2}^{L_1} u_{L_3,L_1}^{(q+\lambda)}(L_2, y)(\mathcal{T}_{1,0}((a+1)y) + \omega_{1,0}(a+1)y))dy \bigg) v_0$$
(5.38)

$$+\lambda \left(\int_{L_3}^{L_2} u_{L_3,L_1}^{(q+\lambda)}(L_2, y) v_2^J((a+1)y) \mathrm{d}y + \int_{L_2}^{L_1} u_{L_3,L_1}^{(q+\lambda)}(L_2, y) v_1^J((a+1)y) \mathrm{d}y \right).$$
(5.39)

Indeed, (5.33) is the expected discounted dividends when level L_3 is reached before any other level in \mathcal{N} and before a jump. The term (5.34) is the expected discounted dividends when a jump occurs before reaching L_1 or L_3 , and just before the jump U is between L_2 and L_3 . Similarly, (5.35) and (5.36) are the expected discounted dividends when a jump occurs before reaching L_1 or L_3 and after this jump the first level that is reached is L_1 . Additionally, (5.37) and (5.38) are the expected discounted dividends when a jump occurs before reaching L_1 or L_3 and after this jump the first level that is reached is L_0 . Moreover, (5.39) is the expected discounted dividends due to overflow above *b* when a jump occurs before reaching L_1 or L_3 and after this jump the first level that is reached is $L_0 = b$ due to a dividend payment after up-crossing *b* by a jump. Using similar arguments we can conclude that for $2 \le n \le N - 1$ we have

$$v_n = \left(Z^{(q+\lambda)}(L_n - L_{n+1}) - \frac{W^{(q+\lambda)}(L_n - L_{n+1})}{W^{(q+\lambda)}(L_{n-1} - L_{n+1})} Z^{(q+\lambda)}(L_{n-1} - L_{n+1}) \right) v_{n+1}$$
(5.40)

$$+ \left(\lambda \int_{L_{n+1}}^{L_n} u_{L_{n+1},L_{n-1}}^{(q+\lambda)}(L_n, y) r_{n,n}((a+1)y) \mathrm{d}y\right) v_n$$
(5.41)

+
$$\left(\lambda \int_{L_{n+1}}^{L_n} u_{L_{n+1},L_{n-1}}^{(q+\lambda)}(L_n, y)(r_{n,n-1}((a+1)y) + \omega_{n,n-1}((a+1)y))dy\right)$$
 (5.42)

$$+\frac{W^{(q+\lambda)}(L_n-L_{n+1})}{W^{(q+\lambda)}(L_{n-1}-L_{n+1})} + \lambda \int_{L_n}^{L_{n-1}} u_{L_{n+1},L_{n-1}}^{(q+\lambda)}(L_n,y)r_{n-1,n-1}((a+1)y)dy v_{n-1}$$
(5.43)

$$+\lambda \sum_{k=2}^{n-1} \left(\int_{L_{n+1}}^{L_n} u_{L_{n+1},L_{n-1}}^{(q+\lambda)}(L_n, y)(r_{n,n-k}((a+1)y) + \omega_{n,n-k}((a+1)y)) dy \right)$$
(5.44)

$$+ \int_{L_n}^{L_{n-1}} u_{L_{n+1},L_{n-1}}^{(q+\lambda)} (L_n, y) (r_{n-1,n-1-(k-1)}((a+1)y)$$
(5.45)

$$+ \omega_{n-1,n-1-(k-1)}((a+1)y))dy \bigg) v_{n-k} \\ + \lambda \bigg(\int_{L_{n+1}}^{L_n} u_{L_{n+1},L_{n-1}}^{(q+\lambda)}(L_n, y) \omega_{n,0}((a+1)y)dy$$
(5.46)

$$\int_{L_{n+1}}^{L_{n+1}} u_{L_{n+1},L_{n-1}}^{(q+\lambda)}(L_n, y)\omega_{n-1,0}((a+1)y)dy + \int_{L_{n+1}}^{L_n} u_{L_{n+1},L_{n-1}}^{(q+\lambda)}(L_n, y)\mathcal{T}_{n,0}((a+1)y)dy$$
(5.47)

$$+ \int_{L_{n}}^{L_{n-1}} u_{L_{n+1},L_{n-1}}^{(q+\lambda)}(L_{n}, y) \mathcal{T}_{n-1,0}((a+1)y) dy \bigg) v_{0} \\ + \lambda \bigg(\int_{L_{n+1}}^{L_{n}} u_{L_{n+1},L_{n-1}}^{(q+\lambda)}(L_{n}, y) v_{n}^{J}((a+1)y) dy \\ + \int_{L_{n}}^{L_{n-1}} u_{L_{n+1},L_{n-1}}^{(q+\lambda)}(L_{n}, y) v_{n-1}^{J}((a+1)y) dy \bigg).$$
(5.48)

The terms (5.40)–(5.45) are obtained by the same arguments as those leading to (5.18). The terms (5.46)–(5.47) describe the expected discounted time to reach L_0 by the Brownian motion when it is the first level reached in \mathcal{N} . In (5.46), level L_0 is reached by the Brownian motion, and in (5.47) it is reached immediately after a jump above L_0 . Finally, (5.48) describes the

expected discounted dividends paid due to up-crossing of L_0 when it occurs before any other level in \mathcal{N} has been reached. Finally, notice that

$$v_N = 0.$$
 (5.49)

To sum up, we have the following main result.

Theorem 5.2. The value function $v_N(x)$, defined formally in (1.6), is given by (5.24), with $r_{n,n-k}$, $\omega_{n,n-k}$ identified in Propositions 5.1 and 5.2; with $v_n^J(x)$, $\mathcal{T}_{n,0}(x)$ identified in Proposition 5.3; and with v_n solving the system of linear equations (5.26)–(5.49).

6. Suggestions for further research

The present study could serve as a first step towards the analysis of more general classes of dual risk processes with proportional gains. Below we suggest a few topics for further research:

- (i) One could consider more general jumps up from level *u*, possibly of the form $u + \zeta(u) + C_i$, where $\zeta(u)$ is a subordinator.
- (ii) In Sections 4 and 5 we have considered proportional growth at jump epochs, assuming that $C_i \equiv 0$. It would be interesting to remove the latter assumption.
- (iii) Another interesting research topic is an exact analysis of the value function v(x) defined in (1.5), without recourse to the approximation approach with levels L_0, \ldots, L_N . One would then have to solve the differential-delay equation (4.2).

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Competing interests

There were no competing interests to declare which arose during the preparation or publication process of this article.

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