



# Adolescent vegetable consumption: the role of socioemotional family characteristics

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## Abstract

**Objective:** To describe associations between adolescents' frequency of vegetable consumption, food parenting practices and socioemotional family characteristics, and to explore potential mediated relationships that may contribute to an understanding of the family processes involved.

**Design:** Cross-sectional survey among adolescents aged 13–15 years.

**Setting:** A survey questionnaire including self-report measures on adolescents' frequency of vegetable consumption, perceived food parenting practices (i.e. family dinner frequency, maternal/paternal healthy eating guidance (HEG), maternal/paternal social support for vegetable consumption) and socioemotional family characteristics (i.e. general family functioning and level of cohesion and conflict within the family) was distributed in a convenience sample of secondary school students.

**Participants:** Four hundred forty students from five secondary schools in eastern Norway completed the questionnaire.

**Results:** Results from multiple linear regression analysis revealed positive and statistically significant associations between adolescents' frequency of vegetable consumption, maternal HEG and family cohesion. A partial indirect (mediated) association between family cohesion and adolescents' frequency of vegetable consumption, working through maternal HEG, was also found.

**Conclusions:** Results from the present study suggest that perceived family cohesion may influence adolescents' frequency of vegetable consumption both directly and indirectly. However, there is a need for continued investigation of family-related factors influencing adolescent eating. In particular, the role of socioemotional family characteristics should be further scrutinised in future studies.

## Keywords

Vegetable consumption  
Adolescents  
Socioemotional family characteristics  
Family cohesion  
Food parenting practices  
Healthy eating guidance

An insufficient intake of fruit and vegetables (FV) is found to be among the leading risk factors of the global burden of non-communicable diseases<sup>(1)</sup>. Therefore, it is a concern that most adolescents in the Nordic countries and elsewhere have a lower consumption of FV, particularly vegetables, than recommended by the authorities<sup>(2–4)</sup>. Adolescence is known as a critical period for the development of dietary behaviours<sup>(5)</sup>, and since such behaviours are likely to track into adult life<sup>(6)</sup>, it is important to increase FV consumption among adolescents to reduce morbidity and mortality from non-communicable diseases. Thus, continued research aiming to reveal key influences on adolescent FV consumption and to develop effective interventions tailored at this group of the population seems imperative.

In Norway, which is the setting of the present study, the Norwegian School Fruit Scheme was launched in 2007 as part of a Norwegian governmental initiative to promote and increase the consumption of FV among children and adolescents. The Norwegian School Fruit Scheme provided students in all secondary schools (grades 8–10) and all combined schools (grades 1–10) with a free piece of fruit or vegetable every school day. The programme lasted for 7 years and resulted in an increased fruit consumption among adolescents regardless of gender and socio-economic status. However, the same positive effect was not found for vegetables<sup>(7–9)</sup>. One obvious reason is that the programme primarily delivered fruits to the students. Thus, the potential for increasing vegetable consumption

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through the Norwegian School Fruit Scheme was limited. Correspondingly, a review by Evans *et al.*<sup>(10)</sup> found that school-based interventions moderately improved fruit intake but had minimal impact on vegetable intake.

The influence of the traditional Norwegian meal pattern, which typically includes one hot meal (dinner) and two or three cold meals<sup>(11)</sup>, must be taken into consideration in the assessment of Norwegian adolescents' vegetable consumption. The cold meals usually consist of bread or cereals. Fruit is more practical to eat with these cold meals and in between meals than vegetables, as they come in convenient portion sizes, 'in their own package' and need little treatment prior to eating compared with vegetables<sup>(12)</sup>. Consequently, in Norway, vegetables are mostly eaten at dinner<sup>(13)</sup>, which most children and adolescents share with their families<sup>(14,15)</sup>. The importance of family meals for more healthful food choices has been stated in reviews by Berge<sup>(16)</sup> and Fulkerson *et al.*<sup>(17)</sup>, and the presence of at least one parent at meals has been associated with a higher FV consumption<sup>(17,18)</sup>. Moreover, irregular family meals (breakfasts and dinners) have been associated with less vegetable consumption in a recent study by Totland *et al.*<sup>(19)</sup>. Hence, a relevant approach for understanding predictors of adolescents' vegetable consumption might be recognising family-related factors, besides shared family dinners, that could influence this behaviour.

Since the family environment has been acknowledged as a fundamental context for the development of eating behaviours<sup>(20)</sup>, several studies have addressed factors such as socio-economic position<sup>(2)</sup> and various food parenting practices, including the arrangement of family meals<sup>(21–23)</sup>. However, less research has focused on fundamental, socioemotional family characteristics such as general family functioning and level of family cohesion or conflict as correlates of adolescent eating<sup>(24)</sup>. Previous research has linked socioemotional family characteristics to social and emotional outcomes in youth<sup>(25–27)</sup>, and Kitzmann and Beech<sup>(28)</sup> have accentuated the importance of exploring these fundamental family characteristics in relation to (un)healthy eating among adolescents. Moreover, these features of the family environment have been suggested as contexts that may enhance or limit the effectiveness of family-based interventions<sup>(28,29)</sup>. Thus, it seems relevant to scrutinise the role of socioemotional family characteristics as potential determinants of adolescent vegetable consumption.

The present study is part of the Family and Dietary Habits project and is based on a framework constructed to describe various family environmental levels and constructs included in questionnaires developed for this project<sup>(30)</sup>. The Family and Dietary Habits framework constitutes an ecological model emphasising factors within the family environment that may contribute to explain dietary behaviours in adolescents. The factors included in this framework are organised in three levels: an individual level (adolescent eating and personal characteristics), a parental level (parenting style and food parenting practices, including parents' arrangement of shared family meals) and a

family level (fundamental socioemotional family characteristics), with factors interacting within and across levels. Adolescent vegetable consumption was the *individual level* factor of interest in the present study and served as the dependent variable in model analyses. *Parental level* factors hypothesised to be related to adolescent vegetable consumption were food parenting practices such as family dinner frequency, healthy eating guidance (HEG) and positive encouragement for vegetable consumption. *Family level* factors hypothesised to be related to adolescent vegetable consumption were fundamental socioemotional family characteristics such as general family functioning, family cohesion and family conflict.

Based on this introductory section, the objectives of the present study were to: (1) describe associations between adolescent vegetable consumption and the parental and family level factors presented above and (2) explore potential mediated relationships that may contribute to an increased understanding of the family processes involved.

## Methods

### *Participants and procedures*

Secondary school students aged 13–15 years were recruited through a convenience sample of five public schools in eastern Norway. Since the students were underaged, a parental consent form including questions assessing household educational level was distributed to the students' parents. After receiving written consent from parents, students who agreed to participate were asked to complete a web-based questionnaire during school hours. Of the 1136 students invited to take part in this cross-sectional study, 440 (39%) completed the questionnaire.

### *Questionnaire*

The students spent between 25 and 45 min on completing the questionnaire, which consisted of 141 questions assessing dietary intake (vegetables and sugar-sweetened beverages), accessibility and availability (of vegetables and sugar-sweetened beverages), personal characteristics, family meals, parenting styles, food parenting practices, socioemotional family characteristics and sociodemographic factors<sup>(30)</sup>. All measures in the questionnaire, including parental and family level factors, were assessed from the perspective of the students. The subset of measures used in the current study is presented in the following paragraphs.

### *Frequency of vegetable consumption*

Vegetable consumption was assessed using frequency measures reproduced from Lien *et al.*<sup>(31)</sup>. Students were asked two questions to report their usual intake of cold (raw) and heated (boiled, fried, roasted, etc.) vegetables, respectively, on an eight-point frequency scale (1 = never/seldom, 2 = less than once a week, 3 = 1–2 times



a week, 4 = 3–4 times a week, 5 = 5–6 times a week, 6 = once a day, 7 = 2 times a day, 8 = 3 times a day or more). Vegetable juices were not included in this measure. As suggested by Andersen *et al.*<sup>(32)</sup>, the response categories were recoded to reflect vegetable consumption in times/week prior to data analyses (0 = never/seldom; 0.5 = less than once a week; 1.5 = 1–2 times a week; 3.5 = 3–4 times a week; 5.5 = 5–6 times a week; 7 = once a day; 14 = twice a day; 21 = 3 times a day). Consequently, all response categories had a common denominator (times a week), which improved the readability of the results and increased comparability with studies using similar measures<sup>(31–34)</sup>. Total frequency of vegetable consumption was calculated by adding up the consumption of cold and heated vegetables.

### **Family dinner frequency**

Family dinner frequency was measured with one item: 'How often does your mother and/or father usually sit down and eat dinner with you?'. Response alternatives were given on an eight-point frequency scale (1 = never/seldom; 2 = once a week; 3 = twice a week; 4 = 3 times a week; 5 = 4 times a week; 6 = 5 times a week; 7 = 6 times a week; 8 = 7 times a week). This variable was not normally distributed as most of the adolescents ate dinner together with their parent(s) 6 or 7 times/week (80.5%). Therefore, responses were dichotomised into '0–5 times a week' and '6–7 times a week'.

### **Healthy eating guidance**

The concept of HEG was developed by Haszard *et al.*<sup>(35)</sup> and includes food parenting practices like teaching about nutrition, modelling healthy eating, encouraging a balanced and varied diet and making healthy foods and beverages accessible in the home. In the present study, perceived maternal and paternal HEG was measured separately for mothers and fathers by the nine-item HEG subscale adapted from Haszard *et al.*'s<sup>(35)</sup> five-factor version of Musher-Eizenman and Holub's<sup>(36)</sup> Comprehensive Feeding Practices Questionnaire. To the authors' knowledge, the Comprehensive Feeding Practices Questionnaire has previously been used to assess food parenting practices from a parental perspective only. Thus, the items had to be slightly modified to represent the perspective of adolescents in the current research. For example, 'My mother/father discusses with me why it is important to eat healthy foods' (see Appendix 1 for a complete list of HEG items). The HEG items were scored on a five-point Likert scale ranging from 1 (disagree) to 5 (agree), where the sum of scores was divided by 9 to give a total average score ranging from 1.0 to 5.0. Higher scores indicate higher levels of HEG. Haszard *et al.*<sup>(35)</sup> reported good internal consistency reliability for the HEG subscale with an  $\alpha$  coefficient of 0.82. Also, construct validity was supported by Haszard *et al.*<sup>(35)</sup>, as parents with concern for child overweight, and parents who rated a healthy diet as very important for their child,

were found to report higher levels of HEG. Previous testing of the Comprehensive Feeding Practices Questionnaire with parents in a Norwegian setting indicated that this instrument is also a valid tool for measuring multiple parental feeding practices with parents of 10–12-year-olds<sup>(37)</sup>.

### **Positive encouragement for vegetable consumption**

Parents' encouragement of healthy eating behaviours has been associated with positive outcomes<sup>(38)</sup>. For example, Melbye, Øgaard and Øverby<sup>(39)</sup> found a positive association between parental encouragement of a balanced and varied diet and vegetable intake in 10–12-year-olds. Furthermore, Young, Fors and Hayes<sup>(40)</sup> found that perceived parental support for FV consumption was a significant predictor of FV consumption in adolescents. In the current research, perceived maternal and paternal positive encouragement was measured by the five-item Positive Encouragement Subscale (PES) adapted from Dave *et al.*'s<sup>(38)</sup> Emotional Social Support Scale for FV intake. Since vegetable consumption was the dependent variable of interest in the present study, the PES was modified to cover intake of vegetables only: 'How often, during the past month, did your mother/father (1) compliment you for your vegetable consumption; (2) encourage you to eat vegetables when you were tempted not to; (3) discuss your vegetable consumption with you; (4) remind you to eat vegetables; and (5) asked you on ideas on how you could eat more vegetables'. The PES items were scored on a five-point frequency scale ranging from 1 (never) to 5 (very often), where the sum of scores was divided by 5 to give a total average score ranging from 1.0 to 5.0. Higher scores indicate higher levels of positive encouragement. Dave *et al.*<sup>(38)</sup> observed good internal consistency for the PES with an  $\alpha$  coefficient of 0.82. They also found evidence of construct validity, as the PES correlated with related measures such as reinforcement, availability and accessibility.

### **General family functioning**

General family functioning includes structural, organisational and interactional patterns of the family as described by Bowen's<sup>(41)</sup> Family Systems Theory. According to this theory, the interactions that occur within a family are reciprocal in that each member of the family is being shaped by other family members' behaviours. These mutual influences may provide insight into behaviours that ultimately determine health outcomes in individual family members<sup>(24)</sup>. In the current research, perceived family functioning was measured with the General Functioning Scale, which is a twelve-item subscale extracted from the McMaster Family Assessment Device<sup>(42)</sup>. For example, 'Planning family activities is difficult because we misunderstand each other' (see Appendix 1 for a complete list of General Functioning Scale items). The General Functioning Scale response categories ranged from 1 (strongly agree) to 4 (strongly disagree), where the sum of scores was divided by 12 to give a total average score





ranging from 1.0 to 4.0. A higher score (i.e.  $\geq 2.0$ ) indicates poorer family functioning<sup>(43,44)</sup>. The General Functioning Scale has demonstrated good psychometric properties with adolescents in various cultural contexts<sup>(45–47)</sup>.

### **Family cohesion and family conflict**

Family cohesion has been defined as the degree of perceived commitment, support and help family members provide for each other, or as the emotional connection between family members<sup>(48)</sup>. Cohesion is recognised as an important influence on children's development and functioning<sup>(49,50)</sup> and has been shown to affect adolescents' feeling of control over their own health<sup>(51)</sup>. Interestingly, previous studies have suggested a link between family cohesion and healthy dietary behaviours among adolescents<sup>(50,52–54)</sup>. Family conflict has been defined as the degree of perceived aggression and conflict among family members<sup>(55)</sup>, and in contrast to cohesion, it has been associated with negative outcomes in children and adolescents<sup>(55–58)</sup>. For example, in a study by Schuetzmann *et al.*<sup>(59)</sup>, conflict and rejection were linked to deviant eating behaviour in preadolescents. Furthermore, family conflict has been associated with unhealthy eating in high-school students<sup>(54)</sup>. In the present study, family cohesion and family conflict were measured by items derived from the Cohesion (9 items) and Conflict (8 items) subscales included in the Family Environment Scale (FES)<sup>(48)</sup>. For example, cohesion: 'Family members really help and support one another'. For example, conflict: 'We fight a lot in our family' (see Appendix 1 for a complete list of items measuring cohesion and conflict). The cohesion and conflict items were scored on a four-point scale ranging from 1 (true) to 4 (false). As for other rating scale variables in this study, averaged sum scores for the cohesion and conflict subscales were calculated. Higher scores indicate higher levels of cohesion and conflict, respectively<sup>(60)</sup>. Previous research assessing the psychometric properties of the FES, from which the cohesion and conflict scales are derived, has shown inconsistent results. In terms of internal consistency reliability for the FES subscales, the originally reported  $\alpha$  coefficients varied between 0.64 and 0.79, with an acceptable benchmark to be above 0.60 (this value was justified by the emphasis placed on the breadth of the measured constructs)<sup>(48)</sup>. In a study by Charalampous, Kokkinos and Panayiutou<sup>(60)</sup>, where the validity and reliability of the FES were tested with individuals aged 16–60 years, the cohesion and conflict scales emerged as unidimensional, supporting the convergent validity of the scales. The  $\alpha$  coefficients were found to be similar to the  $\alpha$  coefficients originally reported by Moos and Moos<sup>(48)</sup>, 0.74 and 0.64 for cohesion and conflict, respectively. Charalampous, Kokkinos and Panayiutou<sup>(60)</sup> endorsed the strong theoretical basis and predictive utility of the scales which make them fruitful for examining the family environment. Furthermore, in a study by Kalavana, Maes and de Gucht<sup>(54)</sup>, where the FES was administered to senior high-school students (mean age 16.6 years,

SD = 4.8), the construct validity of the cohesion and conflict subscales was supported and both factors had an acceptable internal consistency with  $\alpha$  of 0.76 (cohesion) and 0.74 (conflict). The internal consistency for the cohesion and conflict subscales was also found to be acceptable in a more recent study on adolescents aged 11–18 years, with reported  $\alpha$  of 0.80 (cohesion) and 0.75 (conflict)<sup>(61)</sup>.

### **Sociodemographic factors**

Sociodemographic factors are well-known, inflexible correlates of dietary behaviours and were included as covariates in the current study. Data from the parent with the longest education were used as a measure of highest household educational level and were classified as 'less than or equal to 12 years', 'between 13 and 16 years' and 'more than 16 years'. Gender was classified as 'boy' or 'girl'. Family structure was classified as 'living with both parents' *v.* 'other living arrangements'.

### **Statistical analyses**

#### *Initial analyses*

The SPSS statistical software package version 25 (SPSS Inc.) was used for statistical analyses. Initial analyses included frequencies for categorical variables, and mean scores, standard deviations, skewness, kurtosis, Cronbach's  $\alpha$  and intra-class correlation coefficients for rating scale variables. As suggested by Kline<sup>(62)</sup>, we applied cut-off values of 3.0 and 8.0 for skewness and kurtosis, respectively. Cronbach's  $\alpha$  was used to assess internal consistency reliability for all rating scale variables and was classified as  $>0.70$  = 'acceptable' and  $>0.80$  = 'preferable'<sup>(63)</sup>. Intra-class correlation coefficients were used to assess test-retest reliability in a subsample of adolescents ( $n$  54) and were classified as  $\geq 0.81$  = 'excellent', 0.61–0.80 = 'moderate' and  $\leq 0.40$  = 'poor'<sup>(64)</sup>. Prior to regression analyses, bivariate correlations were run to test for multicollinearity between independent variables. We applied a cut-off value of 0.80 or greater for multicollinearity, as suggested by Haerens *et al.*<sup>(53)</sup>.

#### *Model analyses*

Two different regression strategies were applied to address the research objectives. First, a multiple linear regression analysis was run to describe associations between adolescents' frequency of vegetable consumption (individual level), perceived food parenting practices (parental level) and socioemotional family characteristics (family level) derived from the Family and Dietary Habits framework. Sociodemographic factors were also included and treated as covariates. Next, based on results from the multiple linear regression, potential mediated relationships were explored. The analytical strategy applied to test for mediation was based on Hayes<sup>(65)</sup> modelling tool PROCESS, version 3. This tool includes bootstrapping resampling techniques resulting in more robust results than standard

**Table 1** Frequencies for categorical variables (*n* 440)

Variable	%
Adolescent gender	
Female	52
Male	48
Household educational level	
≤12 years	34
13–16 years	39
>16 years	27
Family structure	
Living with both parents	69
Other living arrangements	31
Family dinner frequency	
0–5 times/week	19
6–7 times/week	81

methods relying on parametric assumptions<sup>(66)</sup>. Since approximately 5000 bootstrap samples are considered sufficient for most applications<sup>(67)</sup>, and since it is the current PROCESS default, we generated 5000 bootstrap samples for the mediation analysis by resampling with replacement from the original sample. Associations between predictor and mediator, and between mediator and outcome variables, were reported in traditional manner by unstandardised coefficients and associated *P*-values. This was also the case for total and direct associations between predictor and outcome variables. Since *P*-values for indirect (mediated) effects are not displayed in the PROCESS output, the indirect effect was reported by unstandardised regression coefficients with 95 % CI. Conforming to the bootstrapping approach, an indirect effect which CI did not include zero was considered statistically significant<sup>(66)</sup>.

## Results

### Initial analyses

Frequencies for categorical variables are presented in Table 1. As can be seen from this table, the sample consisted of 52 % girls and 48 % boys. Most adolescents (74 %) came from highly educated households (34 % with 13–16 years of education, 27 % with more than 16 years of education), and a large proportion (69 %) lived together with both parents. Mean scores, standard deviations, skewness, kurtosis, Cronbach's  $\alpha$  and intra-class correlation coefficients for rating scale variables are presented in Table 2. As depicted in this table, all variables had values within the range of chosen cut-offs for skewness and kurtosis, Cronbach's  $\alpha$  were satisfactory to preferable (range: 0.72–0.89) and intra-class correlation coefficients were good to excellent (range: 0.68–0.83). Finally, no multicollinearities were found for the independent variables to be included in subsequent model analyses (range: 0.01–0.33).

### Model analyses

The multiple linear regression analysis resulted in positive and statistically significant associations between adolescents'

frequency of vegetable consumption and perceived maternal HEG ( $\beta = 0.22$ ,  $P = 0.04$ ), family cohesion ( $\beta = 0.21$ ,  $P = 0.02$ ) and household educational level ( $\beta = 0.11$ ,  $P = 0.04$ ), explaining 9 % of the variance in vegetable consumption (Table 3). Based on these results, a single-mediator model including these variables was tested to explore the potential processes involved.

Since a causal model is the theoretical basis for the examination of potential mediating mechanisms, the temporal order assumption of a causal model was taken into account when specifying this model. To be more specific, in a three-variable mediation model, the independent variable *X* is hypothesised to precede (and cause) mediator *M*, which, in turn, precedes (and causes) dependent variable *Y*, such that accounting for the effect of *X* on *M* and of *M* on *Y* explains, in part or in whole, the influence of *X* on *Y*<sup>(68)</sup>. Following from this, it seems reasonable that family cohesion (which is a fundamental family characteristic) temporally precedes and thus may have the potential to influence, adolescents' perceptions or acknowledgement, of maternal HEG (which is a context-specific, food-related behaviour). The opposite (perceived maternal HEG influencing family cohesion) seems less likely. Likewise, it seems reasonable that maternal HEG temporally precedes and may have the potential to influence adolescents' frequency of vegetable consumption. Consequently, family cohesion was included as the predictor (*X*), while maternal HEG was included as the potential mediator (*M*) of the association between family cohesion and adolescents' frequency of vegetable consumption (*Y*). Household educational level was included as a covariate. Results from mediation analysis showed that family cohesion was significantly and positively associated with maternal HEG ( $\beta = 0.65$ ,  $P < 0.001$ ) and that maternal HEG was significantly and positively associated with adolescents' frequency of vegetable consumption ( $\beta = 0.90$ ,  $P < 0.05$ ). Positive, statistically significant total ( $\beta = 2.63$ ,  $P < 0.001$ ) and direct ( $\beta = 2.04$ ,  $P < 0.01$ ) associations were also found between family cohesion and adolescents' frequency of vegetable consumption. Finally, a statistically significant indirect (mediating) effect of perceived maternal HEG on the association between family cohesion and adolescents' frequency of vegetable consumption ( $\beta = 0.58$ , 95 % CI 0.11, 1.15) was found. This effect accounted for about 22 % of the total effect of family cohesion on adolescent vegetable consumption (i.e. ratio of indirect to total effect,  $P_M = 0.22$ ) and thus represents a partial mediation (Fig. 1).

## Discussion

The present study aimed to: (1) describe associations between adolescents' frequency of vegetable consumption, selected food parenting practices and socioemotional family characteristics; and (2) explore potential mediated relationships that may contribute to an increased understanding of the family processes involved. Family cohesion and maternal HEG was found to be the most important

**Table 2** Means, standard deviations, skewness, kurtosis, Cronbach's  $\alpha$  and intra-class correlations (ICC) for rating scale variables

Variable/scale (number of items)	Mean*	SD*	Skewness*	Kurtosis*	$\alpha^*$	ICC†
Adolescent vegetable consumption‡ (2)	9.46	6.37	1.36	2.47	-	0.69
Family functioning (12)	1.72	0.46	0.56	0.36	0.85	0.83
Family cohesion (9)	3.37	0.45	-0.88	0.54	0.83	0.82
Family conflict (8)	1.89	0.50	0.45	-0.11	0.72	0.73
HEG, mothers (9)	2.16	0.83	-0.81	0.39	0.87	0.71
HEG, fathers (9)	2.31	0.88	-0.68	0.09	0.89	0.68
PEV, mothers (5)	2.41	0.98	0.24	-0.63	0.87	0.72
PEV, fathers (5)	2.24	0.98	0.45	-0.45	0.88	0.68

HEG, healthy eating guidance; PEV, positive encouragement for vegetable consumption.

\*n 440.

†n 54.

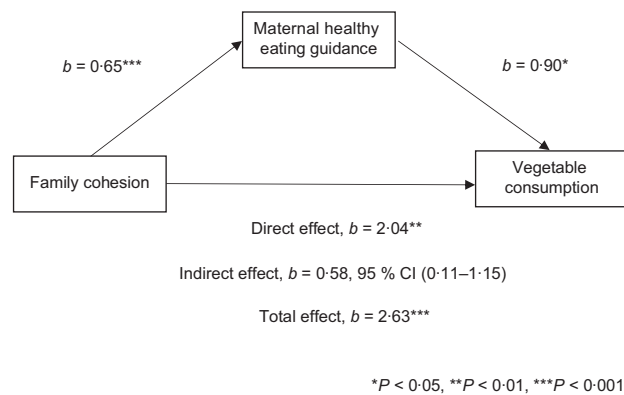
‡Times/week.

**Table 3** Regression coefficients ( $\beta$ ) and variance explained ( $R^2$ ) for multiple linear regression on vegetable consumption frequency

Independent variables	$\beta$
Sociodemographic factors (covariates)	
Adolescent gender	-0.03
Household educational level	0.11*
Family structure	-0.04
Socioemotional family characteristics	
Family functioning	-0.14
Family cohesion	0.21*
Family conflict	-0.07
Sociocultural factors	
Family dinner frequency	0.07
HEG, mothers	0.22*
HEG, fathers	-0.11
PEV, mothers	0.00
PEV, fathers	0.00
$R^2$	0.09

HEG, healthy eating guidance; PEV, positive encouragement for vegetable consumption.

\* $P < 0.05$ .



**Fig. 1.** Path diagram for modelling family cohesion as a predictor of adolescents' frequency of vegetable consumption, partly mediated by maternal healthy eating guidance (HEG). Coefficient estimates ( $b$ ) and statistical significance tests ( $P$ -values and CI) were obtained using the PROCESS script for SPSS

correlates of vegetable consumption frequency in the multiple linear regression model, while household educational level appeared as a weaker correlate. When testing

for mediated relationships (adjusting for household educational level), maternal HEG was found to act as a partial mediator of the positive association between family cohesion and adolescents' frequency of vegetable consumption.

The finding of a positive association between family cohesion and adolescents' frequency of vegetable consumption is supported by previous research by Franko *et al.*<sup>(52)</sup>, which indicates that family cohesion may be linked to healthy eating in numerous ways: First, a cohesive family may be a family that explicitly promotes healthy behaviours: for example, parents encourage healthy eating, and adolescents who feel a high level of connectedness with their parents may be more inclined to follow their suggestions. Second, cohesion has been linked to psychological health, which may have a direct effect on the development of healthy attitudes and behaviours (including healthy eating) in children and adolescents<sup>(52,53,69)</sup>.

The positive association between maternal HEG and adolescents' frequency of vegetable consumption is in line with former studies on social influences postulating that the influence of important others is an essential element in explaining child and adolescent eating behaviours. For example, in a cohort study among children aged 6–11 years and their parents, Couch *et al.*<sup>(70)</sup> found that food parenting practices such as encouragement, modelling and family rules showed strong positive relationships with child FV intake. Positively framed practices such as these were also associated with increased consumption of vegetables and decreased consumption of sugar-sweetened beverages in a Norwegian study on 10–12-year-olds<sup>(71)</sup>. Mothers are of special interest because their food intake has been shown to be related to that of their children, presumably due to their role as 'gatekeepers' of food in the household<sup>(72)</sup>. Conforming to this, results from a study by Pinard *et al.*<sup>(73)</sup>, where home environmental contributors to obesity among children and adolescents aged 5–17 years were explored, indicated that mothers provide much of both the physical (availability/accessibility) and social (role modelling/policies/feeding styles) context in which child and adolescent food choices are made. Furthermore, the finding of perceived maternal HEG as a mediator of the relationship between family cohesion and adolescent vegetable



consumption frequency is in line with the above-mentioned suggestions by Franko *et al.*<sup>(52)</sup>, as a high extent of family cohesion (as perceived by adolescents) may prepare the ground for effective maternal HEG. This could possibly reflect a mechanism where adolescents' perception and appreciation of a cohesive family environment make them more open and responsive to maternal advice and guidance which, in turn, has a favourable effect on adolescent vegetable consumption. Hence, the findings from the present study shed light on potential mechanisms involved in the dynamic relationships between different family environmental levels and adolescent eating behaviours.

No associations were found between adolescents' frequency of vegetable consumption and the parental level factors family dinner frequency and positive encouragement for vegetable consumption (PEV). The lack of association between vegetable consumption frequency and family dinner frequency may be explained by: (1) the fact that in Norway, vegetables are mostly eaten at dinner<sup>(13)</sup>; and (2) the limited variation in family dinner frequency in the population of interest<sup>(14,15)</sup> (the latter was confirmed in the present study). The lack of association seen for PEV could possibly be explained by a lack of parental encouragement specifically targeting vegetable consumption. Another possible explanation may be the adolescents' lack of recognition of such encouragement. The relatively low mean values for perceived maternal and paternal PEV in the present sample (see Table 2) support this line of reasoning. It is worth noting that the way children perceive getting support from their parents and the way parents perceive offering their support may be very different from each other<sup>(38)</sup>.

The lack of associations seen for the family level factors general family functioning and family conflict may be due to these factors' 'distance' to the behaviour of interest, as more proximal environmental (e.g. home availability/accessibility of vegetables), individual (e.g. taste preferences) and social (e.g. peer influence) factors may play a greater role in influencing adolescent dietary behaviours<sup>(74)</sup>. Nevertheless, both general family functioning and conflict may have an impact on the relationships between the more proximal factors and adolescent dietary behaviours, even if we were not able to detect it with the measures and analyses applied in the present study. For example, previous research has indicated that family conflict can significantly predict unhealthy dietary behaviours in adolescents<sup>(54,75)</sup>.

### **Strengths and limitations**

There has been a call for research relating fundamental socioemotional family characteristics to adolescent eating<sup>(24)</sup>. Furthermore, research applying ecological models to increase the understanding of how processes within the family may influence adolescent dietary behaviours have been requested<sup>(76)</sup>. Thus, one strength of the present study is that it adds to the current literature by its ecological approach in assessing influences of the family environment

on adolescent vegetable consumption, thereby acknowledging the dynamic interplay of various factors and levels of the home food environment. More specifically, this work combines well-researched food parenting practices with less explored fundamental socioemotional family characteristics to uncover family environmental influences on adolescent vegetable consumption.

Among the limitations of the current work is the study's cross-sectional nature, which hampers causal inferences. The self-report on all study variables is another limitation, increasing the risk of social desirability responses and common methods bias. The application of a frequency measure for vegetable consumption may also be considered a limitation because of its limited accuracy regarding the amount of vegetables ingested. However, such accuracy was not a key issue in the present work where the intention was to rank individuals according to their usual consumption of vegetables in terms of frequency (i.e. times/week). The lower respondent burden of frequency measures compared with more accurate methods such as repeated 24-h recalls or food diaries, and their ability to capture long-term dietary intake<sup>(77)</sup>, was also reasons for choosing this approach. Frequency measures appear to be feasible instruments in survey research aiming at exploring associations between dietary habits and a wide range of potential determinants without wearing out the respondents. Also, the use of a convenience sample with a large proportion of adolescents from highly educated households may limit the generalisability of our findings. Moreover, the relatively low explanatory power of the multiple linear regression model may be considered a limitation. However, the objective of this study was not to adapt models with the greatest possible explanatory power, but to describe associations between adolescents' frequency of vegetable consumption, food parenting practices and socioemotional family characteristics – and to explore the potential processes involved.

### **Conclusions**

The lower than recommended vegetable consumption in adolescents calls not only for studies and actions tailored directly towards this group of the population. Results from the present study suggest that perceived family cohesion may influence adolescent vegetable consumption both directly and indirectly (through maternal HEG), indicating that research and development of interventions directed towards the socioemotional aspects of the family environment may also be relevant. The large number of studies stating the importance of family meals suggests that developing interventions aimed at increasing the frequency of family meals could be a first step. Based on the findings from the current study, we suggest that a possible second step could be to provide parents with knowledge about how to create a socioemotional family environment that



prepares the ground for positively framed food parenting practices and favourable eating behaviours. However, since knowledge is a prerequisite, but by itself not sufficient to induce behaviour change, parent and adolescent empowerment and other underexplored factors and processes that may help explain adolescent eating behaviours in general, and vegetable consumption in particular, should also be included in future research. Ultimately, understanding the factors and mechanisms at play is essential for the development, implementation and success of any intervention.

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### Supplementary material

For supplementary material accompanying this paper visit <https://doi.org/10.1017/S1368980021001658>

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