

# Simple proofs of Steck's determinantal expressions for probabilities in the Kolmogorov and Smirnov tests

E.J.G. Pitman

This paper gives simple proofs of two theorems of Steck concerning the distribution of sample distribution functions.

Theorems I and II below were stated and proved by Steck in two notable papers [2], [3]. As Steck showed, Theorem I enables us to determine:

- (i) the probability that the empirical distribution function lies between two other distribution functions;
- (ii) very general confidence regions for an unknown distribution function;
- (iii) the power of a test based on the empirical distribution function.

From Theorem II we can obtain the null distribution of the two-sample Smirnov statistics for arbitrary sample sizes. Steck's proofs were indirect, and somewhat complicated. Mohanty [1] gave a shorter proof of Theorem II. Here I give simple, direct proofs of both theorems.

LEMMA. *Let*

$$A_1, A_2, \dots, A_{m-1},$$

$$B_1, B_2, \dots, B_m$$

*denote events such that for any integer  $k$ , the sets*

---

Received 28 April 1972.

$$\{A_r, B_s ; r < k, s \leq k\} , \{A_r, B_s ; r > k, s > k\}$$

are independent, then

$$P(B_1 B_2 \dots B_m A_1 A_2 \dots A_{m-1}) = \Delta_m = \det(d_{ij}) , \quad 1 \leq i, j \leq m ,$$

where

$$\begin{aligned} d_{ij} &= 0 && \text{if } i > j + 1 , \\ &= 1 && \text{if } i = j + 1 , \\ &= P(B_i) && \text{if } i = j , \\ &= P(B_i B_{i+1} \dots B_j A_i^* A_{i+1}^* \dots A_{j-1}^*) && \text{if } i < j , \end{aligned}$$

and  $A_r^*$  is the complement of  $A_r$ .

Note that the conditions on the events make  $B_1, B_2, \dots, B_m$  independent. The events  $A_1, A_2, \dots$  are 1-dependent. Put

$$\bar{A}_r = A_1 A_2 \dots A_r , \quad \bar{B}_r = B_1 B_2 \dots B_r .$$

The lemma may be proved by use of the principle of inclusion and exclusion.

$$\begin{aligned} P(\bar{B}_m A_1 A_2 \dots A_{m-1}) &= P(\bar{B}_m) - \sum_{r < s} P(\bar{B}_m A_r^*) + \sum_{r < s} P(\bar{B}_m A_r^* A_s^*) \\ &\quad - \dots + (-1)^{m-1} P(\bar{B}_m A_1^* A_2^* \dots A_{m-1}^*) , \end{aligned}$$

which can be shown directly to be the expansion of  $\Delta_m$ . A proof by induction is shorter, and easier to print.

Assume the lemma true for  $m = n$ .

$$\Delta_n = P(\bar{B}_n \bar{A}_{n-1}) .$$

Consider  $\Delta_{n+1}$ . The elements of its last row are all zero except

$$d_{n+1,n} = 1 , \quad d_{n+1,n+1} = P(B_{n+1}) .$$

Therefore

$$\Delta_{n+1} = P(B_{n+1}) \Delta_n - \Delta'_n ,$$

where  $\Delta'_n$  differs from  $\Delta_n$  only in having  $B_n$  in the last column of  $\Delta_n$

replaced by  $B_n B_{n+1} A_n^*$ , which satisfies the same conditions relative to the other events appearing in  $\Delta'_n$  as does  $B_n$ . Therefore

$$\begin{aligned} \Delta'_n &= P(\overline{B}_n B_{n+1} A_n^* \overline{A}_{n-1}) = P(\overline{B}_{n+1} \overline{A}_{n-1} A_n^*) ; \\ P(B_{n+1}) \Delta_n &= P(B_{n+1}) P(\overline{B}_n \overline{A}_{n-1}) = P(\overline{B}_{n+1} \overline{A}_{n-1}) . \end{aligned}$$

Thus

$$\begin{aligned} \Delta_{n+1} &= P(\overline{B}_{n+1} \overline{A}_{n-1}) - P(\overline{B}_{n+1} \overline{A}_{n-1} A_n^*) \\ &= P(\overline{B}_{n+1} \overline{A}_{n-1} A_n) = P(\overline{B}_{n+1} \overline{A}_n) , \end{aligned}$$

and so the lemma is true for  $m = n + 1$ . It is easy to show that it is true for  $m = 2$ , and so it is true for all  $m$ .

**COROLLARY.** Taking the case where every  $P(B_i) = 1$ , we obtain the following result for the 1-dependent sequence of events  $A_1, A_2, \dots$ ,

$$P(A_1 A_2 \dots A_{m-1}) = \det(d_{ij}) ,$$

where

$$\begin{aligned} d_{ij} &= 0 && \text{if } i > j + 1 , \\ &= 1 && \text{if } i = j \text{ or } j + 1 , \\ &= P(A_i^* A_{i+1}^* \dots A_{j-1}^*) && \text{if } i < j . \end{aligned}$$

**THEOREM I.** Let

$$0 \leq u_1 \leq u_2 \leq \dots \leq u_m \leq 1 ,$$

$$0 \leq v_1 \leq v_2 \leq \dots \leq v_m \leq 1 ,$$

be given constants such that

$$u_i < v_i , \quad i = 1, 2, \dots, m .$$

If  $U_1, U_2, \dots, U_m$  are the order statistics (in ascending order) from a sample of  $m$  independent uniform random variables with range 0 to 1,

$$P(u_i \leq U_i \leq v_i , 1 \leq i \leq m) = m! \det \left[ (v_i - u_j)_+^{j-i-1} / (j-i-1)! \right] ,$$

where  $(x)_+ = \max(x, 0)$ , and it is understood that determinant elements

for which  $i > j + 1$  are all zero.

Proof. Let  $Y_1, Y_2, \dots, Y_m$  be independent random variables, each with a uniform distribution from 0 to 1. The required probability is equal to

$$(1) \quad m!P(u_i \leq Y_i \leq v_i, 1 \leq i \leq m; Y_1 \leq Y_2 \leq \dots \leq Y_m) .$$

Denote by  $B_i$  the event  $u_i \leq Y_i \leq v_i$ . Denote by  $A_i$  the event  $Y_i \leq Y_{i+1}$ , and by  $A_i^*$  the complement of  $A_i$ , that is, the event  $Y_i > Y_{i+1}$ . The events  $A_i, B_i$  satisfy the conditions of the lemma.

Hence

$$(2) \quad P(u_i \leq Y_i \leq v_i, 1 \leq i \leq m; Y_1 \leq Y_2 \leq \dots \leq Y_m) \\ = P(B_1 B_2 \dots B_m A_1^* A_2^* \dots A_{m-1}^*) = \det(d_{ij}) . \\ d_{ii} = P(B_i) = v_i - u_i .$$

If  $i < j$ ,  $d_{ij} = P(B_i B_{i+1} \dots B_j A_i^* A_{i+1}^* \dots A_{j-1}^*)$ . The event  $B_i B_{i+1} \dots B_j A_i^* A_{i+1}^* \dots A_{j-1}^*$  is

$$u_r \leq Y_r \leq v_r, \quad i \leq r \leq j, \\ Y_i > Y_{i+1} > \dots > Y_j .$$

This is equivalent to

$$v_i \geq Y_i > Y_{i+1} > \dots > Y_j \geq u_j ,$$

the probability of which is  $(v_i - u_j)_+^{j-i+1} / (j-i+1)!$ . The theorem then follows from (1) and (2).

**THEOREM II.** Let  $b_1 \leq b_2 \leq \dots \leq b_m$  and  $c_1 \leq c_2 \leq \dots \leq c_m$  be sequences of integers such that  $b_i < c_i$ . The number of sets of integers  $(R_1, R_2, \dots, R_m)$  such that

$$R_1 < R_2 < \dots < R_m, \\ b_i < R_i < c_i, \quad 1 \leq i \leq m$$

is the  $m$ -th order determinant  $\det(d_{ij})$ , where

$$d_{ij} = 0 \quad \text{if } i > j + 1 \text{ or if } c_i - b_j \leq 1,$$

$$= \begin{pmatrix} c_i - b_j + j - i - 1 \\ j - i + 1 \end{pmatrix} \quad \text{otherwise.}$$

Proof. Put  $Y_i = R_i - i$ ,  $u_i = b_i - i + 1$ ,  $v_i = c_i - i - 1$ . The conditions on the  $R_i$  are equivalent to

$$Y_i \text{ an integer, } u_i \leq Y_i \leq v_i; \quad 1 \leq i \leq m,$$

$$Y_1 \leq Y_2 \leq \dots \leq Y_m.$$

As before, denote by  $A_i$  the event  $Y_i \leq Y_{i+1}$ , and by  $A_i^*$  its complement, the event  $Y_i > Y_{i+1}$ . Put  $N_i = v_i - u_i + 1$ .

The required number is equal to

$$(3) \quad N_1 N_2 \dots N_m P(A_1 A_2 \dots A_{m-1})$$

when the  $Y_i$  are independent random variables, and  $Y_i$  has a uniform distribution over the integers from  $u_i$  to  $v_i$ . By the corollary to the lemma, this is

$$N_1 N_2 \dots N_m \det(d'_{ij}),$$

where

$$d'_{ij} = 0 \quad \text{if } i > j + 1,$$

$$= 1 \quad \text{if } i = j \text{ or } j + 1,$$

$$= P(A_i^* A_{i+1}^* \dots A_{j-1}^*) \quad \text{if } i < j;$$

$$P(A_i^* A_{i+1}^* \dots A_{j-1}^*) = P(v_i \geq Y_i > Y_{i+1} > \dots > Y_j \geq u_j)$$

as before. This is zero if  $v_i - u_j < j - i$ , that is if  $c_i - b_j \leq 1$ .

Otherwise it is equal to

$$\begin{pmatrix} v_i - u_j + 1 \\ j - i + 1 \end{pmatrix} / N_i N_{i+1} \dots N_j = \frac{d_{ij}}{N_i N_{i+1} \dots N_j}.$$

The numerator is the number of vectors of integers  $(y_i, y_{i+1}, \dots, y_j)$  satisfying  $v_i \geq y_i > y_{i+1} > \dots > y_j \geq u_j$ , and the denominator is the number of vectors satisfying  $u_r \leq y_r \leq v_r$ ,  $i \leq r \leq j$ .

Put  $M_0 = 1$ ,  $M_r = N_1 N_2 \dots N_r$ ; then in all cases

$$d'_{i,j} = d_{i,j} M_{i-1} / M_j.$$

The required number (3) is

$$M_m \det(d'_{i,j}) = M_m \det(d_{i,j} M_{i-1} / M_j) = \det(d_{i,j}),$$

as may be obtained by taking factors out of rows and out of columns. This proves the theorem.

#### References

- [1] S.G. Mohanty, "A short proof of Steck's result on two-sample Smirnov statistics", *Ann. Math. Statist.* 42 (1971), 413-414.
- [2] G.P. Steck, "The Smirnov two sample tests as rank tests", *Ann. Math. Statist.* 40 (1969), 1449-1466.
- [3] G.P. Steck, "Rectangle probabilities for uniform order statistics and the probability that the empirical distribution function lies between two distribution functions", *Ann. Math. Statist.* 42 (1971), 1-11.

301 Davey Street,  
Hobart,  
Tasmania 7000.