

## ON THE DISTRIBUTION OF SUM OF INDEPENDENT POSITIVE BINOMIAL VARIABLES

BY  
J. C. AHUJA

1. **Introduction.** Let  $X_1, X_2, \dots, X_n$  be  $n$  independent and identically distributed random variables having the positive binomial probability function

$$(1) \quad f(x; p) = \binom{N}{x} p^x (1-p)^{N-x} / 1 - (1-p)^N, \quad x \in T$$

where  $0 < p < 1$ , and  $T = \{1, 2, \dots, N\}$ . Define their sum as  $Y = X_1 + X_2 + \dots + X_n$ . The distribution of the random variable  $Y$  has been obtained by Malik [2] using the inversion formula for characteristic functions. It appears that his result needs some correction. The purpose of this note is to give an alternative derivation of the distribution of  $Y$  by applying one of the results, established by Patil [3], for the generalized power series distribution. The distribution function of  $Y$  is also found in an explicit form in terms of a linear combination of the incomplete beta functions.

2. **Distribution of sum.** If we take  $p = \theta / (1 + \theta)$ , the probability function (1) may be written as

$$(2) \quad f(x; \theta) = \binom{N}{x} \theta^x / g(\theta), \quad x \in T$$

where  $g(\theta) = (1 + \theta)^N - 1$ . The binomial expansion of  $[g(\theta)]^n$  in powers of  $\theta$  gives

$$\begin{aligned} [g(\theta)]^n &= [(1 + \theta)^N - 1]^n \\ &= \sum_{r=0}^n (-1)^{n-r} \binom{n}{r} (1 + \theta)^{Nr} \\ &= \sum_{r=0}^n (-1)^{n-r} \binom{n}{r} \sum_{y=0}^{Nr} \binom{Nr}{y} \theta^y \end{aligned}$$

which, after changing the order of summation, becomes

$$(3) \quad [g(\theta)]^n = \sum_{y=0}^{Nn} \left[ \sum_{r=0}^n (-1)^{n-r} \binom{n}{r} \binom{Nr}{y} \right] \theta^y$$

where the terms in the second summation are zero for  $r < y/N$ . Using the binomial coefficient identity (12.17) given by Feller [1, p. 65], it can be easily verified that

$$\sum_{r=0}^n (-1)^{n-r} \binom{n}{r} \binom{Nr}{y} = 0$$

for  $y=0, 1, \dots, n-1$ , so that (3) reduces to

$$(4) \quad [g(\theta)]^n = \sum_{y=n}^{Nn} \left[ \sum_{r=0}^n (-1)^{n-r} \binom{n}{r} \binom{Nr}{y} \right] \theta^y.$$

We now recall that the probability function (2) is a special case of the generalized power series distribution as defined by Patil [3] with range  $T$  and the series function  $g(\theta)=(1+\theta)^N-1$ . So the result (7) of Patil [3] and the series expansion (4) provide us the distribution of  $Y$  as

$$(5) \quad h(y) = \sum_{r=1}^n (-1)^{n-r} \binom{n}{r} \binom{Nr}{y} \theta^y / [(1+\theta)^N - 1]^n$$

for  $y=n, n+1, \dots, Nn$ , since the term in the summation is zero for  $r=0$ . Taking  $\theta=p/1-p$  in (5), we get the distribution of  $Y$  in the form

$$(6) \quad h(y) = \sum_{r=1}^n (-1)^{n-r} \binom{n}{r} \binom{Nr}{y} p^y (1-p)^{Nn-y} / [1 - (1-p)^N]^n$$

for  $y=n, n+1, \dots, Nn$ . Further, it may be easily seen that the distribution function of  $Y$  is obtained as

$$(7) \quad \begin{aligned} F(y) &= 1 - \sum_{x=y+1}^{Nn} \left\{ \sum_{r=1}^n (-1)^{n-r} \binom{n}{r} \binom{Nr}{x} p^x (1-p)^{Nn-x} / [1 - (1-p)^N]^n \right\} \\ &= 1 - [1 - (1-p)^N]^{-n} \sum_{r=1}^n (-1)^{n-r} \binom{n}{r} (1-p)^{N(n-r)} I_p(y+1, Nr-y) \end{aligned}$$

where  $I_p(y+1, Nr-y)$  is the incomplete beta function tabulated by Pearson [4].

It may be remarked that the distribution of  $Y$  obtained by Malik [2, p. 335] should read as

$$(8) \quad f(y) = \sum_{r=0}^{n-1} b_r \binom{N(n-r)}{y} p^y q^{Nn-y}$$

for  $y=n, n+1, \dots, Nn$ , where  $p+q=1$ , and  $b_r = \binom{n}{r} (-1)^r / (1-q^N)^n$ , which shows that (6) and (8) are identical.

REFERENCES

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3. G. P. Patil, *Minimum variance unbiased estimation and certain problems of additive number theory*, *Ann. Math. Statist.* **34** (1963), 1050-1056.
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PORTLAND STATE UNIVERSITY,  
PORTLAND, OREGON