ON THE GENERALISATION OF SIDEL'NIKOV'S THEOREM TO q-ARY LINEAR CODES

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Abstract

We generalise Sidel'nikov's theorem from binary codes to *q*-ary codes for q > 2. Denoting by A(z) the cumulative distribution function attached to the weight distribution of the code and by $\Phi(z)$ the standard normal distribution function, we show that $|A(z) - \Phi(z)|$ is bounded above by a term which tends to 0 when the code length tends to infinity.

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1. Introduction

For the binary alphabet, it is well known that the cumulative distribution of linear codes can be approximated by a standard normal distribution. If \mathscr{C} is an [n, k, d] binary linear code with weight distribution (A_0, A_1, \ldots, A_n) , where A_j is the number of codewords in \mathscr{C} with weight j, we define $\mathbf{a} = (a_0, a_1, \ldots, a_n)$, where $a_j = A_j/2^k$. The mean and variance of \mathbf{a} are $\mu(\mathbf{a}) = \sum_{j=0}^n ja_j$ and $\sigma^2(\mathbf{a}) = \sum_{j=0}^n (\mu(\mathbf{a}) - j)^2 a_j$, respectively. The cumulative distribution function (cdf) associated with \mathbf{a} is $A(z) = \sum_{j\geq\mu(\mathbf{a})-\sigma(\mathbf{a})z}^n a_j$. Let

$$\Phi(z) = \frac{1}{\sqrt{2\pi}} \int_{-\infty}^{z} e^{-t^{2}/2} dt$$

be the cdf of the normal law and let d' be the minimum distance of the dual code \mathscr{C}^{\perp} . Sidel'nikov [4] proved that $|A(z) - \Phi(z)| = O(1/\sqrt{d'})$ for *n* large when $d' \ge 3$, which means that $\Phi(z)$ can be regarded as an asymptotic approximation of A(z).

For *q*-ary alphabets, Delsarte [1] showed that the cdf A(z) of linear codes can still be approximated by $\Phi(z)$. However, he did not provide a detailed proof. In this paper, we consider the situation of *q*-ary linear codes again, and rigorously prove that the asymptotic relation between A(z) and $\Phi(z)$ still holds. More specifically, we derive the bound $|A(z) - \Phi(z)| \le C/\sqrt[6]{d'}$, where *C* is a constant that depends only on *q*.

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2. Preliminaries

For any real vector $\mathbf{v} = (v_0, v_1, \dots, v_n)$ with $v_j \ge 0$ and $\sum_{j=0}^n v_j = 1$, the mean and variance of \mathbf{v} are defined by

$$\mu(\mathbf{v}) = \sum_{j=0}^{n} j v_j \quad \text{and} \quad \sigma^2(\mathbf{v}) = \sum_{j=0}^{n} (\mu(\mathbf{v}) - j)^2 v_j.$$

The *s*th central moment of **v** is

$$\mu_s(\mathbf{v}) = \sum_{j=0}^n \left(\frac{\mu(\mathbf{v}) - j}{\sigma(\mathbf{v})}\right)^s v_j.$$

Let \mathscr{C} be an [n, k, d] *q*-ary linear code and A_j the number of codewords in \mathscr{C} with weight *j*. Define $a_j = A_j/q^k$ so that $\mathbf{a} = (a_0, a_1, \ldots, a_n)$ satisfies the conditions above. Hence $\mu(\mathbf{a}), \sigma^2(\mathbf{a})$ and $\mu_s(\mathbf{a})$ can be defined. The cdf of \mathbf{a} is given by

$$A(z) = \sum_{j \ge \mu(\mathbf{a}) - \sigma(\mathbf{a})z}^{n} a_j.$$

Let $\mathbf{b} = (b_0, b_1, \dots, b_n)$, where $b_j = q^{-n} {n \choose j} (q-1)^j$. Then $\mathbf{b} \sim B(n, 1-1/q)$, **b** satisfies the conditions above,

$$\mu(\mathbf{b}) = \frac{(q-1)n}{q}, \quad \sigma^2(\mathbf{b}) = \frac{(q-1)n}{q^2}$$

and

$$\mu_{s}(\mathbf{b}) = \sum_{j=0}^{n} \left(\frac{\mu(\mathbf{b}) - j}{\sigma(\mathbf{b})}\right)^{s} b_{j} = \frac{q^{-n}}{\sqrt{(q-1)n^{s}}} \sum_{j=0}^{n} \left[(q-1)n - qj\right]^{s} {\binom{n}{j}} (q-1)^{j}.$$

We use this notation throughout the paper. Further details on q-ary linear codes can be found in [6–10].

3. Main result

LEMMA 3.1. Let \mathscr{C} be an [n, k, d] q-ary linear code and d' the minimum distance of \mathscr{C}^{\perp} . For $s = 0, 1, \ldots, d' - 1$,

$$\mu_s(\boldsymbol{a}) = \mu_s(\boldsymbol{b}).$$

PROOF. Applying the MacWilliams identity for q-ary codes [3, Equation (M3), page 257],

$$W_{\mathscr{C}^{\perp}}(x,y) = \frac{1}{|\mathscr{C}|} W_{\mathscr{C}}(y-x,y+(q-1)x).$$
(3.1)

Let A'_i (i = 0, 1, ..., n) be the number of codewords with weight *i* in \mathscr{C}^{\perp} and let a'_i be the MacWilliams transform of a_i with parameter $\lambda = q - 1$ [1]. Substituting y = 1 into (3.1) gives the expansion

$$\sum_{i=0}^{n} A'_{i} x^{i} = \frac{1}{q^{k}} \sum_{i=0}^{n} A_{i} (1-x)^{i} [1+(q-1)x]^{n-i} = \sum_{i=0}^{n} a_{i} (1-x)^{i} [1+(q-1)x]^{n-i}.$$

We then find that, for $i = 0, 1, \ldots, n$,

$$a_i' = \frac{A_i'}{q^{n-k}}.$$

Since $a'_1 = \cdots = a'_{d'-1} = 0$, by Delsarte [1, Lemma 4], we have $\mu_s(\mathbf{a}) = \mu_s(\mathbf{b})$ for $s = 0, 1, \ldots, d' - 1$.

Because the definitions of μ , σ^2 and μ_s in the *q*-ary case are the same as in the binary case, the formulas from Sidel'nikov's derivation [4, Equations (35)–(40), pages 285–286] remain correct. Setting r = 2[(d' - 1)/2] gives the following lemma.

LEMMA 3.2. *For all* T > 0,

$$|A(z) - \Phi(z)| \le \frac{1}{\pi} \int_{-T}^{T} \frac{1}{|t|} \left| \sum_{s=0}^{\infty} \mu_s(b) \frac{(it)^s}{s!} - e^{-t^2/2} \right| dt + \frac{2}{\pi} \int_{-T}^{T} \mu_r(b) \frac{|t|^{r-1}}{r!} dt + \frac{24}{T\pi\sqrt{2\pi}}.$$
(3.2)

LEMMA 3.3. Let \mathscr{C} be an [n, k, d] q-ary linear code and d' the minimum distance of \mathscr{C}^{\perp} . If $n \ge 6$ and $d' \ge \frac{1}{2}n + 3$, then r = 2[(d' - 1)/2] satisfies

$$r \ge \frac{n}{2} \ge \frac{d'}{2}.$$

PROOF. From the definition of *r*, together with $n \ge 6$ and $d' \ge \frac{1}{2}n + 3$,

$$r = 2\left[\frac{d'-1}{2}\right] \ge 2\left(\frac{d'-3}{2}\right) \ge \frac{n}{2} \ge \frac{d'}{2}.$$

This completes the proof.

We now focus on the right-hand side of (3.2) and give upper bounds for each of the three terms.

LEMMA 3.4. For all T with $0 < T \le T_0 = n^{1/6}/(3\rho^{1/3})$,

$$\frac{1}{\pi} \int_{-T}^{T} \frac{1}{|t|} \left| \sum_{s=0}^{\infty} \mu_s(\boldsymbol{b}) \frac{(it)^s}{s!} - e^{-t^2/2} \right| dt \le \frac{8\rho}{\sqrt{2\pi}} \cdot \frac{1}{\sqrt{n}}$$

where $0 < \rho < \infty$ and ρ is a constant that depends only on q.

PROOF. Define the random variables $\xi \sim B(n, 1 - 1/q)$, $\eta = (E\xi - \xi)/\sqrt{D\xi}$, and let $\varphi_{\eta}(t)$ be the characteristic function of η . We find

$$\begin{split} \varphi_{\eta}(t) &= \sum_{j=0}^{n} \exp\left(it \cdot \frac{(q-1)n - jq}{\sqrt{n(q-1)}}\right) {\binom{n}{j}} (q-1)^{j} q^{-n} \\ &= \sum_{s=0}^{\infty} \sum_{j=0}^{n} \left(\frac{(q-1)n - jq}{\sqrt{n(q-1)}}\right)^{s} {\binom{n}{j}} (q-1)^{j} q^{-n} \frac{(it)^{s}}{s!} \\ &= \sum_{s=0}^{\infty} \mu_{s}(\mathbf{b}) \frac{(it)^{s}}{s!}. \end{split}$$

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[3]

Thus the first term on the right-hand side of (3.2) is equal to

$$\frac{1}{\pi}\int_{-T}^{T}\frac{1}{|t|}|\varphi_{\eta}(t)-e^{-t^{2}/2}|\,dt.$$

Define independent random variables ξ_1, \ldots, ξ_n , where each ξ_j satisfies

$$P(\xi_j = \sqrt{q-1}) = \frac{1}{q}, \quad P\left(\xi_j = -\frac{1}{\sqrt{q-1}}\right) = 1 - \frac{1}{q}.$$

It is easy to verify that $E\xi_j = 0$ and $D\xi_j = 1$. If we now set $s_n^2 = \sum_{j=1}^n D\xi_j$, then $\eta = s_n^{-1} \sum_{j=1}^n \xi_j$. Define $F_j(x)$ to be the distribution function of ξ_j and

$$\rho_j = \sup_{z>0} \left(\left| \int_{-z}^z x^3 dF_j(x) \right| + z \int_{|x| \ge z} x^2 dF_j \right).$$

Observe that $\rho_1 = \cdots = \rho_n$, so we can set $\rho = \rho_j$. From the definition of the Riemann–Stieltjes integral, $0 < \rho < \infty$ and ρ only depends on q. From Esseen [2, Lemma 5], for $|t| \le T \le T_0 = n^{1/6}/(3\rho^{1/3})$,

$$|\varphi_{\eta}(t) - e^{-t^{2}/2}| \leq \frac{4\sum_{j=1}^{n} \rho_{j}}{s_{n}^{3}} |t|^{3} e^{-t^{2}/2} = \frac{4\rho}{\sqrt{n}} |t|^{3} e^{-t^{2}/2}.$$

Hence

$$\begin{aligned} \frac{1}{\pi} \int_{-T}^{T} \frac{1}{|t|} |\varphi_{\eta}(t) - e^{-t^{2}/2}| \, dt &\leq \frac{4\rho}{\pi \sqrt{n}} \int_{-T}^{T} |t|^{2} e^{-t^{2}/2} \, dt \\ &\leq \frac{8\rho}{\pi \sqrt{n}} \int_{0}^{+\infty} t^{2} e^{-t^{2}/2} \, dt \\ &= \frac{8\rho}{\sqrt{2\pi}} \cdot \frac{1}{\sqrt{n}}. \end{aligned}$$

This completes the proof.

LEMMA 3.5 [5]. Define a random variable $X \sim B(n, p)$. Then, for all even r,

$$\mu_r(X) = \sigma(X)^{-r} E(X - \mu(X))^r \le \left(\frac{2}{p(1-p)}\right)^{r/2} \cdot \frac{r!}{(r/2)!}$$

LEMMA 3.6. For all T > 0 and all even r,

$$\frac{2}{\pi} \int_{-T}^{T} \mu_r(\boldsymbol{b}) \frac{|t|^{r-1}}{r!} dt \le \frac{4e^{1/24}T^r}{\pi\sqrt{r} \cdot r!} \cdot \left(\frac{4rq^2}{e(q-1)}\right)^{r/2}.$$

PROOF. Since $\mathbf{b} \sim B(n, 1 - 1/q)$, by Lemma 3.5,

$$\mu_r(\mathbf{b}) \le \left(\frac{2q^2}{q-1}\right)^{r/2} \cdot \frac{r!}{(r/2)!}.$$
(3.3)

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Substituting (3.3) into the second term on the right-hand side of (3.2),

$$\frac{2}{\pi} \int_{-T}^{T} \mu_{r}(\mathbf{b}) \frac{|t|^{r-1}}{r!} dt \leq \frac{4T^{r}}{\pi r \cdot r!} \cdot \left(\frac{2q^{2}}{q-1}\right)^{r/2} \cdot \frac{r!}{(r/2)!} \\
< \frac{4e^{1/24}T^{r}}{\pi\sqrt{r} \cdot r!} \cdot \left(\frac{2q^{2}}{q-1}\right)^{r/2} \cdot \left(\frac{2r}{e}\right)^{r/2} \\
= \frac{4e^{1/24}T^{r}}{\pi\sqrt{r} \cdot r!} \cdot \left(\frac{4rq^{2}}{e(q-1)}\right)^{r/2}.$$
(3.4)

The observation $\sqrt{2\pi m}(m/e)^m < m! < \sqrt{2\pi m}(m/e)^m e^{1/12m}$ for all $m \in N^*$ has been used in the second inequality in (3.4).

If we choose a suitable *T* satisfying $0 < T \le T_0 = n^{1/6}/(3\rho^{1/3})$ and collect all the results above, we reach the following bound for the right-hand side of (3.2), which gives the generalisation of Sidel'nikov's theorem for *q*-ary linear codes.

THEOREM 3.7. Let \mathscr{C} be an [n, k, d] q-ary linear code and d' the minimum distance of \mathscr{C}^{\perp} . If $n \ge 6$ and $d' \ge \frac{1}{2}n + 3$, then

$$|A(z) - \Phi(z)| \le \frac{C}{\sqrt[6]{d'}}$$

where C is a constant that depends only on q.

PROOF. Choose T with $0 < T \le T_0 = n^{1/6}/(3\rho^{1/3})$, so that $\sqrt{n} \ge 27\rho T^3$. Using this along with Lemmas 3.4 and 3.6, the right-hand side of (3.2) is

$$\leq \frac{8}{27T^3\sqrt{2\pi}} + \frac{4e^{1/24}T^r}{\pi\sqrt{r}\cdot r!} \cdot \left(\frac{4rq^2}{e(q-1)}\right)^{r/2} + \frac{24}{T\pi\sqrt{2\pi}}.$$
(3.5)

Let

$$c = \min\left(\frac{1}{16}, \frac{(q-1)e}{9\rho^{2/3}}\right)$$
 and $T = \left(\frac{cr}{(q-1)e}\right)^{1/2} \cdot \frac{1}{n^{1/3}}$.

Then $0 < T \le (n/(9\rho^{2/3}))^{1/2} \cdot n^{-1/3} = T_0$ and we can substitute *T* into (3.5). Finally, from the inequality $(r/e)^r/r! < 1/\sqrt{2\pi r}$ and Lemma 3.3,

$$\begin{split} |A(z) - \Phi(z)| &\leq \frac{8[(q-1)e]^{3/2}}{27c^{3/2}\sqrt{2\pi}} \cdot \left(\frac{n^{1/3}}{\sqrt{r}}\right)^3 + \frac{4e^{1/24}}{\pi\sqrt{r}\cdot r!} \cdot \left(\frac{r}{e}\right)^r \cdot \left(\frac{2qc^{1/2}}{(q-1)n^{1/3}}\right)^r \\ &+ \frac{24[(q-1)e]^{1/2}}{c^{1/2}\pi\sqrt{2\pi}} \cdot \frac{n^{1/3}}{\sqrt{r}} \\ &\leq \frac{16[(q-1)e]^{3/2}}{27c^{3/2}\sqrt{2\pi}} \cdot \frac{1}{\sqrt{r}} + \frac{4e^{1/24}}{\pi\sqrt{2\pi}} \cdot \frac{1}{r} + \frac{24[(q-1)e]^{1/2} \cdot 2^{1/3}}{c^{1/2}\pi\sqrt{2\pi}} \cdot \frac{1}{\sqrt[6]{r}} \\ &\leq \frac{16[(q-1)e]^{3/2}}{27c^{3/2}\sqrt{2\pi}} \cdot \frac{1}{\sqrt{r}} + \frac{8e^{1/24}}{\pi\sqrt{2\pi}} \cdot \frac{1}{r} + \frac{24[(q-1)e]^{1/2}}{c^{1/2}\pi\sqrt{2\pi}} \cdot \frac{1}{\sqrt[6]{r}} \\ &\leq \frac{16[(q-1)e]^{3/2}}{27c^{3/2}\sqrt{\pi}} \cdot \frac{1}{\sqrt{d'}} + \frac{8e^{1/24}}{\pi\sqrt{2\pi}} \cdot \frac{1}{d'} + \frac{24[(q-1)e]^{1/2}}{c^{1/2}\pi\sqrt{\pi}} \cdot \frac{1}{\sqrt[6]{d'}} \\ &= \frac{c_1}{\sqrt{d'}} + \frac{c_2}{d'} + \frac{c_3}{\sqrt[6]{d'}} \leq \frac{C}{\sqrt[6]{d'}}, \end{split}$$

where *C* is a constant that depends only on *q*.

4. Conclusion and open problem

The estimate $|A(z) - \Phi(z)| = O(1/\sqrt[6]{d'})$ when $n \to \infty$ that we have obtained for *q*-ary codes is coarser than Sidel'nikov's upper-bound $O(1/\sqrt{d'})$ for the binary case. A challenging open problem is to improve the above estimates.

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